

GMM with first-step residuals: A recipe for control-function S.E.s

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The plan

- Use gmm to get standard errors for control function type estimators
 - ▶ linear cross-sectional model
 - ▶ fractional (binary) outcome cross-sectional model
 - ▶ exponential mean (Poisson) panel-data model
- Control function estimates imply:
 - ▶ A test for endogeneity
 - ▶ A structural function interpretation of effects
- It is common to use the bootstrap
- Excuse to show you some gmm Jujutsu
- Discuss some estimation and postestimation considerations

Model I: Linear model

$$y = X_1\beta_1 + X_2\beta_2 + \varepsilon$$

$$E(X'_1\varepsilon) = \mathbf{0}$$

$$E(X'_2\varepsilon) \neq \mathbf{0}$$

- Different estimators arise from the following:

$$E(Z'\varepsilon) = 0 \quad \text{Instrumental variables}$$

$$X_2 = Z\Pi + \nu \quad \text{Two stage least squares (TSLS)}$$

$$E(Z'\nu) = 0$$

- Control function approaches additionally assume

$$\varepsilon = \rho\nu + \epsilon$$

$$y = X_1\beta_1 + X_2\beta_2 + \rho\nu + \epsilon$$

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$$\varepsilon = \rho\nu + \epsilon$$

$$y = X_1\beta_1 + X_2\beta_2 + \rho\nu + \epsilon$$

TSLS and Control function

TSLS:

- ① Regression of X_2 on Z , get part of X_2 without endogeneity, \hat{X}_2
 - ▶ We get $\hat{X}_2 = Z(Z'Z)^{-1}Z'X_2 = P_z X_2$
- ② Regression of y on X_1 and \hat{X}_2
 - ▶ We get $\hat{\beta} = (X_2'P_zX_2)^{-1}X_2P_zy$

CONTROL FUNCTION (CF):

- ① Get residuals from Regression of X_2 on Z
 - ▶ $\hat{\nu} = X_2 - \hat{X}_2 = I - P_z X_2$
- ② Regress y on X_1 , X_2 , and $\hat{\nu}$
 - We need to address uncertainty in estimation of ν

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GMM for CF

- GMM needs specification $E\{W'e\} = \mathbf{0}$ or $E\{W'e(\theta)\} = \mathbf{0}$
- In gmm W are exogenous and specified as options
- $e(\theta)$ change at each iteration as θ converges to it's minimum
- The control function approach does not quite fit into gmm's framework. Let $W(\Pi) = [X, X_2 - Z\Pi]$

$$\begin{aligned} E\{Z'(X_2 - Z\Pi)\} &= \mathbf{0} \\ E\{W(\Pi)'[y - X_1\beta_1 + X_2\beta + \rho(X_2 - Z\Pi)]\} &= \mathbf{0} \end{aligned}$$

- Estimating $\hat{\nu}$ and feeding it to gmm will give incorrect standard errors

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- Estimating $\hat{\nu}$ and feeding it to gmm will give incorrect standard errors

Give gmm what it needs not what it wants

- gmm wants a set of instruments for each equation
- Rewrite the system as

$$E \{ Z' (X_2 - Z\Pi) \} = \mathbf{0}$$

$$E \{ X' [y - X_1\beta_1 + X_2\beta_2 + \rho(X_2 - Z\Pi)] \} = \mathbf{0}$$

$$E \{ (X_2 - Z\Pi)' [y - X_1\beta_1 + X_2\beta_2 + \rho(X_2 - Z\Pi)] \} = \mathbf{0}$$

- The last equation satisfies the framework

$$\begin{aligned} E \{ (X_2 - Z\Pi)' [y - X_1\beta_1 + X_2\beta_2 + \rho(X_2 - Z\Pi)] \} &= E \{ \eta(\Pi, \rho, \beta) \} \\ E \{ \eta(\Pi, \rho, \beta) \} &= \mathbf{0} \end{aligned}$$

- It divided the set of instruments for one of the equations

gmm: Substitutable expressions I

$$\begin{aligned} \text{mpg} &= \beta_0 + \beta_1 \text{turn} + \beta_2 \text{foreign} + \varepsilon \\ \text{turn} &= \pi_0 + \pi_1 \text{foreign} + \pi_2 \text{weight} + \nu \end{aligned}$$

```
. sysuse auto, clear  
(1978 Automobile Data)  
. // Writing down substitutable expression  
. local zp {p1}*l.foreign + {p2}*weight + {p0}  
. local u turn -(`zp')  
. local xb {b1}*turn + {b2}*l.foreign + {b3}*(`u') + {b0}  
. local e mpg - (`xb')  
. // Computing  
. gmm (eq3: `u') ///  
> (eq2: (`e')*(`u')) ///  
> (eq1: `e'), ///  
> instruments(eq3: weight i.foreign) ///  
> instruments(eq1: turn i.foreign) ///  
> winitial(unadjusted, independent) ///  
> from(C) onestep quickderivatives
```

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. local e mpg - (`xb')  
. // Computing  
. gmm (eq3: `u') ///  
> (eq2: (`e')*(`u')) ///  
> (eq1: `e'), ///  
> instruments(eq3: weight i.foreign) ///  
> instruments(eq1: turn i.foreign) ///  
> winitial(unadjusted, independent) ///  
> from(C) onestep quickderivatives
```

Considerations

- Do not use `gmm` as a computation engine
- You know the fitted values of θ from `regress`
- Use `gmm` to compute standard errors
- Start optimization at `regress` values
- Use `quickderivatives` uses numerical recipes does not go through `deriv()`
- GMM is exactly identified, use `onestep`
- Sanity check: `ivregress gmm` should give you the same standard errors and point estimates

gmm: Substitutable expressions II

Step 1

Iteration 0: GMM criterion Q(b) = 1.364e-28

Iteration 1: GMM criterion Q(b) = 3.309e-29

note: model is exactly identified

GMM estimation

Number of parameters = 7

Number of moments = 7

Initial weight matrix: Unadjusted

Number of obs = 74

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
/p1	-1.809802	.6316408	-2.87	0.004	-3.047795	-.5718085
/p2	.0042183	.0003777	11.17	0.000	.003478	.0049587
/p0	27.44963	1.347933	20.36	0.000	24.80773	30.09153
/b1	-1.56173	.2071108	-7.54	0.000	-1.96766	-1.1558
/b2	-4.476451	1.790227	-2.50	0.012	-7.985232	-.9676693
/b3	1.326564	.2608633	5.09	0.000	.8152814	1.837847
/b0	84.54861	8.641913	9.78	0.000	67.61077	101.4865

Instruments for equation eq3: weight 0b.foreign 1.foreign _cons

Instruments for equation eq2: _cons

Instruments for equation eq1: turn 0b.foreign 1.foreign _cons

gmm: Starting values

```
. // Getting starting values
. quietly regress turn i.foreign weight
. predict double uhat, residuals
. matrix B = e(b)
. matrix B = B[1,2..colsof(B)]
. quietly regress mpg turn i.foreign uhat
. matrix A = e(b)
. matrix A = A[1,1], A[1,3..colsof(A)]
. matrix C = B,A
```

Evaluator

```
gmm eval [if][in][weight], equations(eqnames)  
parameters(parameter_names)  
[youropts stataopts]
```

- I would write an evaluator instead of using substitutable expressions
- Evaluators allow me to add options
- Evaluators are .ado files so they can be used more widely

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gmm eval [if][in][weight], equations(eqnames)  
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Example

```
program _cf_linear
    version 16
    syntax varlist if [fweight iweight pweight],
        at(name) ///
        [///
        uhat(varlist) ///
        y1(varname) ///
        y2(varlist) ///
        * ///
        ]
        ]///

tempvar zp xb xbu

tokenize `varlist'
local main `1'
local reduced `2'
local aux `3'

matrix score double `xb' = `at' `if', eq(#1)
matrix score double `zp' = `at' `if', eq(#2)
replace `reduced' = `y2' - `zp' `if'
replace `uhat' = `y2' - `zp' `if' // i.v random
matrix score double `xbu' = `at' `if', eq(#3)
replace `main' = `y1' - `xb' - `xbu' `if'
replace `aux' = (`y1' - `xb' - `xbu') * `uhat' `if'

end
```

Evaluator estimates

```
. gmm _cf_linear, equations(mpg turn uhat)           ///
>     parameters( ``y1parm' `y2parm' uhat:uhat")      ///
>     y1(mpg) y2(turn) uhat(uhat)                      ///
>     instruments(turn: i.foreign weight)              ///
>     instruments(mpg: i.foreign turn)                ///
>     winitial(unadjusted, independent)               ///
>     quickderivatives onestep from(CNEW)
```

Evaluator estimates

Iteration 0: GMM criterion Q(b) = 1.186e-28
Iteration 1: GMM criterion Q(b) = 6.355e-29
note: model is exactly identified

GMM estimation

Number of parameters = 7

Number of moments = 7

Initial weight matrix: Unadjusted

Number of obs = 74

		Robust				
		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
mpg	turn	-1.56173	.2071109	-7.54	0.000	-1.96766 -1.1558
	foreign					
	Foreign	-4.476451	1.790228	-2.50	0.012	-7.985233 -.9676679
	_cons	84.54861	8.641919	9.78	0.000	67.61076 101.4865
turn	foreign					
	Foreign	-1.809802	.6316412	-2.87	0.004	-3.047796 -.5718077
	weight	.0042183	.0003777	11.17	0.000	.003478 .0049587
	_cons	27.44963	1.347934	20.36	0.000	24.80773 30.09154
uhat	uhat	1.326564	.2608634	5.09	0.000	.8152812 1.837847

Instruments for equation mpg: 0b.foreign 1.foreign turn _cons

Instruments for equation turn: 0b.foreign 1.foreign weight _cons

Instruments for equation uhat: _cons

Sanity check

```
. estimates store gmm  
. quietly ivregress gmm mpg i.foreign (turn = weight)  
. estimates store ivreg_gmm  
. estimates table gmm ivreg_gmm, eq(1) keep(turn 1.foreign _cons) se
```

Variable	gmm	ivreg_gmm
turn	-1.5617299 .2071109	-1.5617299 .2071109
foreign Foreign	-4.4764507 1.7902282	-4.4764507 1.7902282
_cons	84.548613 8.641919	84.548613 8.641919

legend: b/se

A command

```
cfunction estimator y Xs ..., endogenous(end1 ... endk = ...) ...  
          cfvar([newvars], [...]) ...
```

```
cfunction estimator y Xs ..., endogenous(end1 = ...) ...  
          endogenous(endk = ...) ...  
          cfvar([newvars], [...]) ...
```

- *estimator* is probit, linear, ...
- *endogenous ()* might be multiple equations with different instruments
- A variable is created with the residuals of the first step
- *cfvar (newvars, [replace float])*

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- A variable is created with the residuals of the first step
- *cfvar(newvars, [replace float])*

The command

```
. cfunction linear mpg i.foreign, endogenous(turn = i.foreign weight)
Iteration 0:    EE criterion =  7.381e-29
Iteration 1:    EE criterion =  3.125e-29
Control-function linear regression                               Number of obs      =      74
Outcome model          : regress
Control-function model: regress
```

		Robust Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
mpg						
	turn	-1.56173	.2071356	-7.54	0.000	-1.967708 -1.155752
	foreign					
	Foreign	-4.476451	1.790523	-2.50	0.012	-7.985812 -.9670892
	_cons	84.54861	8.64284	9.78	0.000	67.60896 101.4883

The command

Control-function linear regression						
		Number of obs = 74				
		Outcome model : regress				
Control-function model: regress						
		Robust Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
mpg	turn	-1.56173	.2071356	-7.54	0.000	-1.967708 -1.155752
	foreign	-4.476451	1.790523	-2.50	0.012	-7.985812 -.9670892
	Foreign					
	_cons	84.54861	8.64284	9.78	0.000	67.60896 101.4883
turn	foreign					
	Foreign	-1.809802	.6316412	-2.87	0.004	-3.047796 -.5718077
	weight	.0042183	.0003777	11.17	0.000	.003478 .0049587
	_cons	27.44963	1.347934	20.36	0.000	24.80773 30.09154
__Cf_V_a_R1						
	_cons	1.326564	.2609407	5.08	0.000	.8151298 1.837998

Model II: Fractional (binary) outcomes

$$E(y|X, \nu) = \Phi(X_1\beta_1 + X_2\beta_2 + \rho\nu)$$
$$X_2 = Z\Pi + \nu$$

- You can think of y as $y = X_1\beta_1 + X_2\beta_2 + \varepsilon > 0$ and (ε, ν) being correlated and jointly normal
- y could be described by another model as long as X_2 is continuous and endogeneity is due to a relation of ε and ν
- Z are unrelated to ν

Interpretation

- Coefficients and standard errors are asymptotically equivalent to two-step estimates computed by `ivprobit ...`, `twostep`
- Coefficients should not be taken too seriously. What is important is to think about effects. (Editorial comment).
- We will be able to compute effects because the coefficient vectors and standard errors are kept for all equations

The command

```
. cfunction probit foreign mpg, endogenous(headroom = mpg weight)
Iteration 0:    EE criterion = 1.525e-24
Iteration 1:    EE criterion = 1.065e-31
Control-function fractional regression                         Number of obs      =      74
Outcome model          : fracreg probit
Control-function model: regress
```

foreign	Robust					
	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
foreign						
headroom	-5.017474	2.1125	-2.38	0.018	-9.157898	-.877051
mpg	-.1544128	.1484151	-1.04	0.298	-.445301	.1364754
_cons	17.31753	9.024816	1.92	0.055	-.370786	35.00584

Sanity check I

```
. quietly cfunction probit foreign mpg, endogenous(headroom = mpg weight)
. estimates store cfprobit
. quietly ivprobit foreign mpg (headroom = weight), twostep
. estimates store ivtwo
. estimates table cfprobit ivtwo, se eq(1) drop(`drop')
```

Variable	cfprobit	ivtwo
headroom	-5.0174745 2.1124998	-5.0174745 2.4034212
mpg	-.1544128 .14841505	-.1544128 .15947795
_cons	17.317527 9.0248155	17.317528 10.182072

legend: b/se

Sanity check II

```
. webuse cattaneo2, clear  
(Excerpt from Cattaneo (2010) Journal of Econometrics 155: 138-154)  
. quietly cfunction probit lbweight i.msmoke i.alcohol mage,           ///  
>      endogenous(medu = i.foreign fedu i.msmoke i.alcohol mage)  
. estimates store uno  
. quietly ivprobit lbweight i.msmoke i.alcohol mage (medu = i.foreign fedu)  
. estimates store dos  
. estimates table uno dos, keep(lbweight:) se
```

Variable	uno	dos
medu	-.02132736 .02605357	-.02130645 .02640689
msmoke		
1-5 daily	.01475317 .15390946	.01475978 .14990178
6-10 daily	.51695718 .09787212	.51685433 .09627577
11+ daily	.39551778 .10403561	.39544583 .10282816
alcohol		
1	.18803644 .14087207	.18800059 .14518304
mage	-.01016345 .00732448	-.01016394 .00697637
_cons	-1.1213499 .26433385	-1.121227 .25888826

legend: b/se

General postestimation considerations

- Construct predictions as a function of *cfvar*
- Return `e(covariates)`. Variables margins operates over.
 - ▶ Exclude *cfvar*
 - ▶ Exclude excluded instruments (z's)
- margins only perturbs variables in `e(covariates)`
- The coefficients on *cfvar* provide test for endogeneity
- margins quantities have a structural function interpretation

Model III: Panel exponential mean

$$E(y_{it}|X_{it1}, X_{it2}, \alpha_i, \nu_{it}) = \exp(X_{it1}\beta_1 + X_{it2}\beta_2 + \alpha_i + \rho\nu_{it})$$
$$X_{it2} = Z_{it}\Pi + \gamma_i + \nu_{it}$$

- Fit a fixed effects regression of X_{it2} on Z_{it} or a correlated random effects estimator (i.e. Mundlak) and get residuals
- Compute fixed effects Poisson regression including residuals
- Other estimators and conditions can be considered. They will imply different GMM estimators (i.e Windmeijer (2000) and Lin and Wooldridge (2019))

Moment conditions

- $\widetilde{W} \equiv W_{it} - \overline{W}_i + \overline{W}$
- $(\widetilde{X}_2 - \widetilde{Z}\Pi) \equiv \nu(\Pi)$
- $\exp(X_1\beta_1 + X_2\beta_2 + \rho\nu(\Pi)) \equiv \theta$

$$E \left\{ \widetilde{Z}' \nu(\Pi) \right\} = \mathbf{0}$$

$$E \left\{ X' \left[y - \frac{\bar{y}}{\bar{\theta}} \theta \right] \right\} = \mathbf{0}$$

$$E \left\{ \nu(\Pi)' \left[y - \frac{\bar{y}}{\bar{\theta}} \theta \right] \right\} = \mathbf{0}$$

Evaluator

```
program cfxtpoisson
    version 16
    syntax varlist if [fweight iweight pweight],           ///
        at(name)                                         ///
        [                                              ///
        at(name)                                         ///
        id(string)                                       ///
        uhat(varlist)                                     ///
        y1(varname)                                       ///
        y2(varlist)                                       ///
        *                                              ///
        ]
    tempvar zp xb xbu xbbar ybar
    tokenize `varlist'
    local main `1'
    local reduced `2'
    local aux `3'

    matrix score double `xb'      = `at' `if', eq(#1)
    matrix score double `zp'      = `at' `if', eq(#2)
    replace `reduced'            = `y2' - `zp' `if'
    replace `uhat'               = `y2' - `zp' `if' // i.v random
    matrix score double `xbu'    = `at' `if', eq(#3)

    replace `xb'                 = exp(`xb')
    replace `xbu'                = exp(`xbu')
    egen double `xbbar'          = mean(`xb'*`xbu') `if', by(`id')
    egen double `ybar'            = mean(`y1') `if', by(`id')
    replace `main'               = `y1' - `xb'*`xbu'*`ybar'/`xbbar' `if'
    replace `aux'                = (`main')*`uhat' `if'

end
```

Exponential mean simulated data

$$y2 = 1 - x1 + x2 - z1 + z2 - z3 + u1 + a$$
$$y1 = \exp(0.5(1 - y2 + x1 - x2) + u2 + a)$$

- u_1 and u_2 are correlated jointly normal time-varying unobservables
 - ▶ Correlation $\rho = .7$
- a is a time invariant unobservable and normal correlated with covariates
- All covariates are standardized chi-squares with 5 degrees of freedom

iterlogonly

```
. gmm cfxtpoisson, equations(y1 y2 uhat) id(id)           ///
>          parameters(`y1parm' `y2parm' uhat:uhat")        ///
>          y2(dmy2) y1(y1) uhat(uhat)                      ///
>          instruments(y2: dmz1 dmz2 dmz3 dmx1 dmx2)      ///
>          instruments(y1: y2 x1 x2, nocons)              ///
>          winitial(unadjusted, independent)             ///
>          quickderivatives onestep from(C) iterlogonly    ///
>          vce(cluster id)                                ///
Iteration 0:  GMM criterion Q(b) =  2.696e-20
Iteration 1:  GMM criterion Q(b) =  9.749e-31
```

Results

```
. gmm  
GMM estimation  
Number of parameters = 10  
Number of moments = 10  
Initial weight matrix: Unadjusted Number of obs = 10,000  
(Std. Err. adjusted for 2,000 clusters in id)
```

		Robust				
		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
y1	y2	-.5150497	.0183471	-28.07	0.000	-.5510094 -.47909
	x1	.4935781	.0308917	15.98	0.000	.4330315 .5541247
	x2	-.4997544	.0388162	-12.87	0.000	-.5758327 -.423676
y2	dmz1	-.9914454	.0111035	-89.29	0.000	-1.013208 -.969683
	dmz2	.9952376	.0115265	86.34	0.000	.9726461 1.017829
	dmz3	-1.000285	.0116556	-85.82	0.000	-1.023129 -.9774401
	dmx1	-1.0056	.0110415	-91.07	0.000	-1.027241 -.983959
	dmx2	.9985457	.0109718	91.01	0.000	.9770413 1.02005
	_cons	.9681836	.000386	2508.14	0.000	.967427 .9689402
uhat	uhat	.7922526	.0372532	21.27	0.000	.7192377 .8652675

Instruments for equation y1: y2 x1 x2

Instruments for equation y2: dmz1 dmz2 dmz3 dmx1 dmx2 _cons

Instruments for equation uhat: _cons

A preliminary simulation exercise

Estimator	Bias	Coverage rate
cfunction_linear	-0.000	0.950
cfunction_probit	-0.001	0.953
xtcfunction	0.000	0.945

- $\rho = .5$
- Endogeneity comes from:
 - ▶ Joint normality of time invariant unobservables (all)
 - ▶ Common component in endogenous covariate and time-invariant unobservables
- All covariates are standardized chi-square with 5 degrees of freedom
- N = 3000, T = 5 for cross section I kept one time period (T = 2)
- Results are from 1000 draws
- No time to compare to bootstrap

Conclusion

- I illustrated how to use `gmm` to compute control function estimates and their standard errors
- Along the way I illustrated some tools for those wanting to use `gmm` more efficiently
- Control function GMM standard error estimates are an attractive alternative to bootstrap standard errors
- The estimators open up the possibilities of using `margins` for control function estimates described