

# xtdpdqml: Quasi-maximum likelihood estimation of linear dynamic short-T panel data models

Sebastian Kripfganz

University of Exeter Business School, Department of Economics, Exeter, UK

UK Stata Users Group Meeting  
London, September 9, 2016



```
net install xtdpdqml, from(http://www.kripfganz.de/stata/)
```

# Estimation of short-T linear dynamic panel models in Stata

- Least-squares estimation of dynamic models (i.e. models with a lagged dependent variable) with random or fixed effects (`xtreg` in Stata) yields biased coefficient estimates when the time horizon is short (Nickell, 1981).
- Predominant estimation technique in empirical research is the generalized method of moments (GMM):
  - Arellano and Bond (1991) “difference GMM”: `xtabond`,
  - Arellano and Bover (1995) and Blundell and Bond (1998) “system GMM”: `xtdpdsys`.
  - ⇒ Both Stata commands are wrappers for the more flexible command `xtdpd`.
  - ⇒ Alternative user-written command with full flexibility and many additional options by Roodman (2009): `xtabond2`.

# Estimation of short-T linear dynamic panel models in Stata

- Other promising approaches that can be more efficient alternatives to GMM with potentially better finite-sample performance remain underrepresented in empirical work:
  - Bias-correction procedures by Kiviet (1995), Bun and Kiviet (2003), and Everaert and Pozzi (2007): user-written implementations `xt1sdvc` (Bruno, 2005) and `xtbcfe` (De Vos, Everaert, and Ruysen, 2015).
  - Full-information maximum likelihood / structural equation modeling: `xtdpdm1` command by Williams, Allison, and Moral-Benito (2015) as a wrapper for `sem`.
  - Limited-information quasi-maximum likelihood (QML) estimation for dynamic random-effects models (Bhargava and Sargan, 1983) and dynamic fixed-effects models (Hsiao, Pesaran, and Tahmiscioglu, 2002): new `xtdpdqml` package.

# Linear dynamic panel data model

- Linear panel model with first-order autoregressive dynamics:

$$y_{it} = \lambda y_{i,t-1} + \mathbf{x}'_{it}\boldsymbol{\beta} + \mathbf{f}'_i\boldsymbol{\gamma} + \epsilon_{it}, \quad \epsilon_{it} = u_i + e_{it},$$

$t = 1, 2, \dots, T_i$  (potentially unbalanced but without gaps),  
and where  $e_{it} \stackrel{iid}{\sim} (0, \sigma_e^2)$ . The regressors  $\mathbf{x}_{it}$  and  $\mathbf{f}_i$  are required  
to be strictly exogenous with respect to  $e_{it}$ .

- The lagged dependent variable  $y_{i,t-1}$  is correlated by construction with the unit-specific error component  $u_i$ .
- Dynamic random-effects model:
  - The time-varying regressors  $\mathbf{x}_{it}$  and the time-invariant regressors  $\mathbf{f}_i$  are uncorrelated with  $u_i$ .
- Dynamic fixed-effects model:
  - All regressors are allowed to be correlated with  $u_i$ .

# Dynamic random-effects model

$$y_{it} = \lambda y_{i,t-1} + \mathbf{x}'_{it}\boldsymbol{\beta} + \mathbf{f}'_i\boldsymbol{\gamma} + \epsilon_{it}, \quad \epsilon_{it} = u_i + e_{it},$$

- Random-effects assumption:
  - $u_i \stackrel{iid}{\sim} (0, \sigma_u^2)$ , uncorrelated with  $\mathbf{x}_{it}$  and  $\mathbf{f}_i$ .
- The classical random-effects estimator is a least-squares estimator treating the initial observations  $y_{i0}$  as exogenous. Consequently, it is biased when  $T$  is small due to the correlation of  $y_{i,t-1}$  (and therefore also  $y_{i0}$ ) with  $u_i$ .
- To account for this correlation with a likelihood approach, the joint distribution of  $(y_{i0}, y_{i1}, \dots, y_{iT_i})$  needs to be specified.

# Dynamic random-effects model

$$y_{it} = \lambda y_{i,t-1} + \mathbf{x}'_{it}\boldsymbol{\beta} + \mathbf{f}'_i\boldsymbol{\gamma} + \epsilon_{it}, \quad \epsilon_{it} = u_i + e_{it},$$

- Bhargava and Sargan (1983) propose to model the initial observations as a function of the observed exogenous variables:

$$y_{i0} = \sum_{s=0}^{T^*} \mathbf{x}'_{is}\boldsymbol{\pi}_{x,s} + \mathbf{f}'_i\boldsymbol{\pi}_f + \nu_{i0},$$

with  $T^* = \min(T_i)$ ,  $\text{Var}(\nu_{i0}) = \sigma_0^2$ , and  $\text{Cov}(\nu_{i0}, \epsilon_{it}) = \phi\sigma_0^2$ .

- Implied restrictions under stationarity of all variables:
  - $\phi = \frac{\sigma_u^2}{(1-\lambda)\sigma_0^2}$  in the presence of time-varying regressors  $\mathbf{x}_{it}$ ,
  - $\boldsymbol{\pi}_f = \frac{\boldsymbol{\gamma}}{1-\lambda}$ ,  $\sigma_0^2 = \frac{\sigma_u^2}{(1-\lambda)^2} + \frac{\sigma_e^2}{1-\lambda^2}$ , and  $\phi = \frac{\sigma_u^2}{(1-\lambda)\sigma_0^2}$  in the absence of time-varying regressors  $\mathbf{x}_{it}$ .

# Dynamic fixed-effects model

$$y_{it} = \lambda y_{i,t-1} + \mathbf{x}'_{it}\beta + \mathbf{f}'_i\gamma + \epsilon_{it}, \quad \epsilon_{it} = u_i + e_{it},$$

- Fixed-effects assumption:
  - $u_i$  allowed to be arbitrarily correlated with  $\mathbf{x}_{it}$  and  $\mathbf{f}_i$ .
- First-difference transformation to remove the fixed effects:

$$\Delta y_{it} = \lambda \Delta y_{i,t-1} + \Delta \mathbf{x}'_{it}\beta + \Delta e_{it},$$

- The lagged dependent variable  $\Delta y_{i,t-1}$  (and therefore also  $\Delta y_{i1}$ ) is correlated by construction with the transformed error term  $\Delta e_{it}$ . Consequently, an estimator that treats  $\Delta y_{i1}$  as exogenous is biased.
- To account for this correlation with a likelihood approach, the joint distribution of  $(\Delta y_{i1}, \Delta y_{i2}, \dots, \Delta y_{iT_i})$  needs to be specified.

# Dynamic fixed-effects model

$$\Delta y_{it} = \lambda \Delta y_{i,t-1} + \Delta \mathbf{x}'_{it} \boldsymbol{\beta} + \Delta e_{it},$$

- Hsiao, Pesaran, and Tahmiscioglu (2002) justify the following representation for the initial observations:

$$\Delta y_{i1} = b + \sum_{s=1}^{T^*} \Delta \mathbf{x}'_{is} \boldsymbol{\pi}_s + \nu_{i1},$$

with  $T^* = \min(T_i)$ ,  $\text{Var}(\nu_{i1}) = \omega \sigma_e^2$ ,  $\text{Cov}(\nu_{i0}, \Delta e_{i2}) = -\sigma_e^2$ , and  $\text{Cov}(\nu_{i0}, \Delta e_{it}) = 0$  for  $t > 2$ .

- Implied restrictions under stationarity of all (first-differenced) variables:
  - $b = 0$  in the presence of regressors  $\Delta \mathbf{x}_{it}$ ,
  - $b = 0$  and  $\omega = \frac{2}{1+\lambda}$  in the absence of regressors  $\Delta \mathbf{x}_{it}$ .



# Quasi-maximum likelihood estimation

- Given the assumptions on the error components and treating all of them as if they were normally distributed, the log-likelihood function for the system of equations can be maximized with a gradient-based optimization technique.
- This iterative procedure needs appropriate starting values:<sup>1</sup>
  - By default, `xtdpdqml` obtains initial estimates for the model coefficients from a consistent GMM estimator (`xtdpd`).
  - Initial estimates for the initial-observations coefficients are obtained from a separate least-squares estimation.
  - The initial variance parameter estimates are computed from the respective residuals.
  - Alternative initial estimates for the model coefficients and variance parameters can be specified by the user.
- Analytical first-order and second-order derivatives largely speed up the computations.

---

<sup>1</sup>See the paper and the online appendix at [www.kripfganz.de](http://www.kripfganz.de) for details.

# Stata syntax of the `xtdpdqml` command

```
xtdpdqml depvar [indepvars] [if] [in] [, options]
```

- Selected options:

- `fe`: uses the fixed-effects estimator, the default,
- `re`: uses the random-effects estimator,
- `projection(varlist [, leads(#) nodifference omit])`: specifies the initial-observations projection,
- `stationary`: imposes restrictions valid under stationarity,
- `vce(robust)`: uses the sandwich VC estimator for valid inference under cross-sectional heteroskedasticity (Hayakawa and Pesaran, 2015),
- `mlparams`: reports all ML parameter estimates,
- `from(init_specs)` and `initval(numlist)`: specify alternative starting values,
- additional `display_options`, `maximize_options`, ...

- Selected postestimation commands:

- `predict`: similar to `xtreg` plus equation-level scores,
- `estat`, `hausman`, `lrtest`, `nlcom`, `suest`, `test`, ...

# Example

- Estimation of an employment equation for 140 UK companies, 1976–1984, based on the Arellano and Bond (1991) data set:

```
. webuse abdata
```

- Dependent variable:
  - Logarithm of the number of employees ( $n$ ).
- Strictly exogenous explanatory variables:
  - Real wage ( $w$ ),
  - Gross capital stock ( $k$ ),
  - Time dummies ( $yr1978$ – $yr1984$ ).

# Example

## ● QML estimation of the dynamic fixed-effects model:

```
. xtddpqml n w k yr1978-yr1984, nolog
```

Quasi-maximum likelihood estimation

```
Group variable: id                Number of obs      =       891
Time variable: year              Number of groups   =       140
```

```
Fixed effects                    Obs per group:    min =         6
                                avg =    6.364286
                                max =         8
(Estimation in first differences)
```

n	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
n						
L1.	.7181159	.0349792	20.53	0.000	.6495579	.7866738
w	-.4210157	.0512701	-8.21	0.000	-.5215034	-.3205281
k	.2487324	.0255407	9.74	0.000	.1986736	.2987911
yr1978	-.0214489	.0149487	-1.43	0.151	-.0507478	.00785
yr1979	-.0319754	.0149372	-2.14	0.032	-.0612518	-.0026991
yr1980	-.0637126	.0148821	-4.28	0.000	-.092881	-.0345441
yr1981	-.1130657	.0150739	-7.50	0.000	-.14261	-.0835213
yr1982	-.0844508	.0160798	-5.25	0.000	-.1159666	-.052935
yr1983	-.0461928	.0197008	-2.34	0.019	-.0848057	-.0075798
yr1984	-.0115354	.0241271	-0.48	0.633	-.0588236	.0357528
_cons	1.74826	.1705756	10.25	0.000	1.413938	2.082582

# Example

## ● Reporting of all parameter estimates:

```
. xtldpdqml n w k yr1978-yr1984, mlparams nolog
```

Quasi-maximum likelihood estimation

```
Group variable: id                Number of obs      =       891
Time variable: year              Number of groups   =       140
```

```
Fixed effects                    Obs per group:    min =         6
                                      avg =    6.364286
                                      max =         8
```

	D.n	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----							
_model							
	n						
	LD.	.7181159	.0349792	20.53	0.000	.6495579	.7866738
	w						
	D1.	-.4210157	.0512701	-8.21	0.000	-.5215034	-.3205281
	k						
	D1.	.2487324	.0255407	9.74	0.000	.1986736	.2987911
	yr1978						
	D1.	-.0214489	.0149487	-1.43	0.151	-.0507478	.00785

(Continued on next page)

# Example

```

yr1979 |
  D1. | -.0319754 .0149372 -2.14 0.032 -.0612518 -.0026991
      |
yr1980 |
  D1. | -.0637126 .0148821 -4.28 0.000 -.092881 -.0345441
      |
yr1981 |
  D1. | -.1130657 .0150739 -7.50 0.000 -.14261 -.0835213
      |
yr1982 |
  D1. | -.0844508 .0160798 -5.25 0.000 -.1159666 -.052935
      |
yr1983 |
  D1. | -.0461928 .0197008 -2.34 0.019 -.0848057 -.0075798
      |
yr1984 |
  D1. | -.0115354 .0241271 -0.48 0.633 -.0588236 .0357528
-----|-----
_initobs |
      w |
  D1. | .1745629 .0835193 2.09 0.037 .010868 .3382578
  FD. | .4866594 .1160984 4.19 0.000 .2591107 .714208
  F2D. | .234992 .0921914 2.55 0.011 .0543001 .4156838
  F3D. | .180422 .0831649 2.17 0.030 .0174218 .3434222
  F4D. | .1587507 .0822884 1.93 0.054 -.0025316 .3200329
  F5D. | .1828358 .0801948 2.28 0.023 .025657 .3400147

```

(Continued on next page)

# Example

```

      k |
      D1. | .2516903 .0514379 4.89 0.000 .1508739 .3525068
      FD. | -.0759983 .0442764 -1.72 0.086 -.1627784 .0107819
      F2D. | .0345647 .0402481 0.86 0.390 -.0443201 .1134496
      F3D. | .0426643 .0416536 1.02 0.306 -.0389754 .1243039
      F4D. | .0180357 .0354471 0.51 0.611 -.0514394 .0875108
      F5D. | .1373772 .0420249 3.27 0.001 .0550099 .2197445
      |
      yr1978 |
      D1. | .0472505 .0347851 1.36 0.174 -.0209269 .115428
      FD. | .0336196 .0205327 1.64 0.102 -.0066237 .073863
      |
      _cons | .0034106 .0211468 0.16 0.872 -.0380363 .0448575
-----+-----
      /_sigma2e | .0107403 .0005952 .0095737 .011907
      /_omega | 1.219196 .0690326 1.083894 1.354497
-----+-----

```

```
. estimates store fe
```

# Example

## ● Restricted model versions:

```
. xtddpqml n w k yr1978-yr1984, stationary mlparams nolog  
(Output omitted)
```

```
. lrtest fe
```

```
Likelihood-ratio test                                LR chi2(1) =      0.03  
(Assumption: . nested in fe)                       Prob > chi2 =      0.8720
```

```
. xtddpqml n w k yr1978-yr1984, stationary projection(yr*, omit) mlparams nolog  
(Output omitted)
```

```
. lrtest fe
```

```
Likelihood-ratio test                                LR chi2(3) =      6.29  
(Assumption: . nested in fe)                       Prob > chi2 =      0.0983
```

```
. estimates restore fe  
(results fe are active now)
```

```
. test [_initobs]: D.yr1978 FD.yr1978 _cons
```

```
( 1)  [_initobs]D.yr1978 = 0  
( 2)  [_initobs]FD.yr1978 = 0  
( 3)  [_initobs]_cons = 0
```

```
        chi2( 3) =      6.36  
        Prob > chi2 =      0.0955
```



# Example

```
. xtdpdqml n w k yr1978-yr1984, stationary projection(w k, leads(0)) mlparams nolog
(Output omitted)
```

```
. lrtest fe
```

```
Likelihood-ratio test                    LR chi2(11) =      42.49
(Assumption: . nested in fe)             Prob > chi2 =      0.0000
```

- Alternative starting values from “system GMM” estimator (default starting values are from “difference GMM” estimator):

```
. quietly xtdpsys n w k yr1978-yr1984, twostep
```

```
. matrix b = e(b)
```

```
. xtdpdqml n w k yr1978-yr1984, stationary from(b, skip)
(Output omitted)
```

```
. estimates store fe
```

# Example

- QML estimation of the dynamic random-effects model:

```
. xtddpqml n w k yr1978-yr1984, re nolog
```

```
Quasi-maximum likelihood estimation  
initial values not feasible
```

- Feasible starting values for the variance parameters  $(\sigma_u^2, \sigma_e^2, \sigma_0^2, \phi)$  need to satisfy the restriction

$$(\sigma_u^2 - \phi^2 \sigma_0^2) \max(T_i) > -\sigma_e^2.$$

```
. xtddpqml n w k yr1978-yr1984, re initval(.1 .2 .2 .3) nolog  
(Output omitted)
```

```
. estimates store re
```

# Example

## • Traditional Hausman test:

```
. hausman fe re, df(3)
```

---- Coefficients ----				
	(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
	fe_eq1	re_eq1	Difference	S.E.
n				
L1.	.7175701	.6827449	.0348253	.0226022
w	-.4219682	-.304499	-.1174692	.0284715
k	.2493912	.2630639	-.0136728	.0131214
yr1978	-.0212959	-.0215183	.0002224	.0016011
yr1979	-.0317929	-.0326742	.0008813	.0015725
yr1980	-.0633101	-.0639498	.0006397	.
yr1981	-.1125881	-.1171753	.0045871	.
yr1982	-.0839164	-.0953542	.0114378	.0042314
yr1983	-.0455604	-.0651054	.019545	.006765
yr1984	-.0107753	-.035986	.0252107	.0069979

b = consistent under Ho and Ha; obtained from xtddpqml

B = inconsistent under Ha, efficient under Ho; obtained from xtddpqml

Test: Ho: difference in coefficients not systematic

```
chi2(3) = (b-B)'[(V_b-V_B)^(-1)](b-B)
          = 240.26
Prob>chi2 = 0.0000
(V_b-V_B is not positive definite)
```

# Example

## ● Generalized (robust) Hausman test:

```
. quietly xtldpdqml n w k yr1978-yr1984, mlparams
. estimates store fe
. quietly xtldpdqml n w k yr1978-yr1984, re initval(.1 .2 .2 .3) mlparams
. estimates store re
. suest fe re, vce(cluster id)
```

Simultaneous results for fe, re

```

                                     Number of obs      =      1,031
                                     (Std. Err. adjusted for 140 clusters in id)
-----+-----
                                     |
                                     |           Robust
                                     |           Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
fe_model |
   n |
   LD. |   .7181159   .0806002    8.91  0.000    .5601424   .8760893
       |
   w |
   D1. |  -.4210157   .1316838   -3.20  0.001   -0.6791113 -0.1629202
       |

```

(Continued on next page)

# Example

k							
D1.		.2487324	.047906	5.19	0.000	.1548384	.3426263
yr1978							
D1.		-.0214489	.0142647	-1.50	0.133	-.0494072	.0065095
yr1979							
D1.		-.0319754	.016767	-1.91	0.057	-.0648382	.0008873
yr1980							
D1.		-.0637126	.0180549	-3.53	0.000	-.0990996	-.0283255
yr1981							
D1.		-.1130657	.0209338	-5.40	0.000	-.1540952	-.0720362
yr1982							
D1.		-.0844508	.0190163	-4.44	0.000	-.121722	-.0471796
yr1983							
D1.		-.0461928	.0209038	-2.21	0.027	-.0871635	-.005222
yr1984							
D1.		-.0115354	.02833	-0.41	0.684	-.0670612	.0439905

(Continued on next page)

# Example

```

fe__initobs |
      w
      D1. | .1745629 .0898936 1.94 0.052 -.0016253 .3507512
      FD. | .4866594 .1895771 2.57 0.010 .1150951 .8582237
      F2D. | .234992 .1322934 1.78 0.076 -.0242983 .4942823
      F3D. | .180422 .1104639 1.63 0.102 -.0360833 .3969272
      F4D. | .1587507 .0902785 1.76 0.079 -.018192 .3356933
      F5D. | .1828358 .0940585 1.94 0.052 -.0015154 .3671871
      |
      k
      D1. | .2516903 .078033 3.23 0.001 .0987485 .4046322
      FD. | -.0759983 .0668488 -1.14 0.256 -.2070196 .055023
      F2D. | .0345647 .0385317 0.90 0.370 -.0409561 .1100856
      F3D. | .0426643 .0470128 0.91 0.364 -.0494791 .1348077
      F4D. | .0180357 .0278761 0.65 0.518 -.0366004 .0726719
      F5D. | .1373772 .0447742 3.07 0.002 .0496213 .2251331
      |
      yr1978
      D1. | .0472505 .0210911 2.24 0.025 .0059127 .0885884
      FD. | .0336196 .0155646 2.16 0.031 .0031136 .0641256
      |
      _cons | .0034106 .0205965 0.17 0.868 -.0369577 .0437789
-----+-----
fe__sigma2e |
      _cons | .0107403 .0014299 7.51 0.000 .0079379 .0135428
-----+-----
fe__omega
      _cons | 1.219196 .0819172 14.88 0.000 1.058641 1.379751
-----+-----

```

(Continued on next page)

# Example

```

re__model |
      n |
      L1. | .6827449 .0631622 10.81 0.000 .5589492 .8065406
      |
      w | -.3044499 .1153329 -2.64 0.008 -.5305473 -.0784507
      k | .2630639 .0511424 5.14 0.000 .1628267 .3633012
yr1978 | -.0215183 .0141391 -1.52 0.128 -.0492304 .0061938
yr1979 | -.0326742 .0160256 -2.04 0.041 -.0640839 -.0012645
yr1980 | -.0639498 .0177469 -3.60 0.000 -.0987331 -.0291664
yr1981 | -.1171753 .0216733 -5.41 0.000 -.1596542 -.0746964
yr1982 | -.0953542 .0222249 -4.29 0.000 -.1389142 -.0517943
yr1983 | -.0651054 .0240963 -2.70 0.007 -.1123333 -.0178774
yr1984 | -.035986 .0317191 -1.13 0.257 -.0981542 .0261823
  _cons | 1.43717 .4311311 3.33 0.001 .5921688 2.282172
-----
re__initobs |
      w |
      --. | .4486646 .2996806 1.50 0.134 -.1386987 1.036028
      F1. | -.0795423 .5469361 -0.15 0.884 -1.151517 .9924327
      F2. | -.8357704 .5370137 -1.56 0.120 -1.888298 .2167572
      F3. | -.1347361 .3832975 -0.35 0.725 -.8859854 .6165132
      F4. | .1016144 .3492035 0.29 0.771 -.5828119 .7860408
      F5. | .1846765 .1168485 1.58 0.114 -.0443424 .4136954
      F6. | -.5300617 .2599228 -2.04 0.041 -1.039501 -.0206224

```

(Continued on next page)

# Example

k						
--.		.8302629	.1898999	4.37	0.000	.4580658 1.20246
F1.		-.2463192	.2770439	-0.89	0.374	-.7893152 .2966768
F2.		.3583677	.2750527	1.30	0.193	-.1807258 .8974611
F3.		.0512604	.1811207	0.28	0.777	-.3037297 .4062505
F4.		-.1772404	.2013611	-0.88	0.379	-.5719008 .21742
F5.		.470898	.1744223	2.70	0.007	.1290366 .8127594
F6.		-.4599582	.1731522	-2.66	0.008	-.7993302 -.1205861
yr1978						
F2.		-.1260256	.1120337	-1.12	0.261	-.3456076 .0935563
yr1979						
F2.		-.1369898	.0939022	-1.46	0.145	-.3210347 .047055
_cons		4.181794	.921414	4.54	0.000	2.375856 5.987733
-----						
re__sigma2u						
_cons		.0248997	.0110377	2.26	0.024	.0032663 .0465331
-----						
re__sigma2e						
_cons		.0106025	.0014872	7.13	0.000	.0076877 .0135174
-----						
re__sigma2e0						
_cons		.3161824	.048807	6.48	0.000	.2205225 .4118423
-----						
re__phi						
_cons		.2688014	.0576002	4.67	0.000	.1559072 .3816957
-----						



# Example

```
. test ([fe__model]LD.n = [re__model]L.n) ([fe__model]D.w = [re__model]w) ([fe__model]D.k = [re__model]k)

( 1)  [fe__model]LD.n - [re__model]L.n = 0
( 2)  [fe__model]D.w - [re__model]w = 0
( 3)  [fe__model]D.k - [re__model]k = 0

      chi2( 3) =      5.97
      Prob > chi2 =      0.1132
```

## • Computation of long-run effects:

```
. xtddpqml n w k yr1978-yr1984, stationary vce(robust)
(Output omitted)

. nlcom (_b[w] / (1 - _b[L.n])) (_b[k] / (1 - _b[L.n]))

      _nl_1:  _b[w] / (1 - _b[L.n])
      _nl_2:  _b[k] / (1 - _b[L.n])
```

n	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
_nl_1	-1.494064	.4484327	-3.33	0.001	-2.372976	-.6151519
_nl_2	.8830199	.1834742	4.81	0.000	.523417	1.242623

# Summary: the new `xtdpdqml` package for Stata

- (Quasi-)maximum likelihood estimation can be an attractive alternative to widely used GMM estimators with potential efficiency gains and better finite-sample performance.
- The `xtdpdqml` implements the Bhargava and Sargan (1983) random-effects QML estimator and the Hsiao, Pesaran, and Tahmiscioglu (2002) fixed-effects QML estimator for linear dynamic panel data models.
- It provides a complement to Stata's existing estimation toolbox for dynamic panel models that can be valuable to assess the robustness of estimates obtained with different methods.

---

Kripfganz, S. (forthcoming). `xtdpdqml`: Quasi-maximum likelihood estimation of linear dynamic short-T panel data models. *Stata Journal* (accepted manuscript).

```
net install xtdpdqml, from(http://www.kripfganz.de/stata/)
help xtdpdqml
help xtdpdqml postestimation
```

# References

- Arellano, M., and S. R. Bond (1991). Some tests of specification for panel data: Monte Carlo evidence and an application to employment equations. *Review of Economic Studies* 58(2): 277–297.
- Arellano, M., and O. Bover (1995). Another look at the instrumental variable estimation of error-components models. *Journal of Econometrics* 68(1): 29–51.
- Bhargava, A., and J. D. Sargan (1983). Estimating dynamic random effects models from panel data covering short time periods. *Econometrica* 51(6): 1635–1659.
- Blundell, R., and S. R. Bond (1991). Initial conditions and moment restrictions in dynamic panel data models. *Journal of Econometrics* 87(1): 115–143.
- Bruno, G. S. F. (2005). Estimation and inference in dynamic unbalanced panel-data models with a small number of individuals. *Stata Journal* 5(4): 473–500.
- Bun, M. J. G., and J. F. Kiviet (2003). On the diminishing returns of higher-order terms in asymptotic expansions of bias. *Economics Letters* 79(2): 145–152.
- De Vos, I., G. Everaert, and I. Ruyssen (2015). Bootstrap-based bias correction and inference for dynamic panels with fixed effects. *Stata Journal* 15(4): 986–1018.
- Everaert, G., and L. Pozzi (2007). Bootstrap-based bias correction for dynamic panels. *Journal of Economic Dynamics and Control* 31(4): 1160–1184.
- Hayakawa, K., and M. H. Pesaran (2015). Robust standard errors in transformed likelihood estimation of dynamic panel data models with cross-sectional heteroskedasticity. *Journal of Econometrics* 188(1): 111–134.
- Hsiao, C., M. H. Pesaran, and A. K. Tahmiscioglu (2002). Maximum likelihood estimation of fixed effects dynamic panel data models covering short time periods. *Journal of Econometrics* 109(1): 107–150.
- Kiviet, J. F. (1995). On bias, inconsistency, and efficiency of various estimators in dynamic panel data models. *Journal of Econometrics* 68(1): 53–78.
- Roodman, D. (2009). How to do xtabond2: An introduction to difference and system GMM in Stata. *Stata Journal* 9(1): 86–136.
- Williams, R., P. D. Allison, and E. Moral-Benito (2015). Linear dynamic panel-data estimation using maximum likelihood and structural equation modeling. Presented July 30, 2015 at the Stata Conference 2015, Columbus, Ohio.