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# A command for Laplace regression

Nicola Orsini

Joint work with Matteo Bottai



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# Outline

- Laplace regression model with censored data
- Example 1 – Randomized Clinical Trial
- Laplace regression model with no censored data
- Example 2 – Observational study

# Laplace regression with censored data

- Let  $T_i$  be a sample of a continuous variable (time to an event) with  $i = 1, \dots, n$
- $T_i$  may be censored (study end, lost to follow-up)
- $y_i = \min(T_i, C_i)$
- $\delta_i = I(T_i < C_i)$

Let  $x_i$  be a  $k$ -dimensional vectors of covariates

- $y_i = x'_i \beta_p + u_i$
- $P(u_i | x_i \leq 0) = p \quad \text{with} \quad p \in (0, 1)$
- $P(y_i \leq x'_i \beta_p | x_i) = p$
- $x'_i \beta_p$  is the  $p$ -quantile of the conditional distribution of  $y_i$  given  $x_i$

# Asymmetric Laplace Distribution

We assume that conditionally on covariates the response variable follows an asymmetric Laplace distribution with probability density function

- $f(y_i | x_i) = \exp[(I(B) - p)z] p(1-p)/\sigma_p$

and cumulative distribution function

- $F(y_i | x_i) = \exp[(I(B) - p)z](p - I(A)) + I(A)$

where  $I(B) = (y_i \leq x'_i \beta_p)$ ,  $I(A) = (y_i > x'_i \beta_p)$ , and  $z = (y_i - x'_i \beta_p)/\sigma_p$

# Log-Likelihood

The contribution of the  $i$ -th observation to the log-likelihood is

- $l(\beta_p, \sigma_p | y_i, x_i, \delta_i) = \sum_{i=1}^n [\delta_i \log f(y_i | x_i) + (1 - \delta_i) \log(1 - F(y_i | x_i))]$

The likelihood estimating equations (first derivatives in  $\beta_p$  and  $\sigma_p$  equal to zero) do not have a closed form solution.

# Estimation algorithm

- Maximize the log-likelihood function directly
- Continuity and concavity of the log-likelihood
- Maximum exists and is global
- Gradient search maximization algorithm
- Standard errors based on bootstrap samples

# Example 1

*Lancet* 1999; **353**: 14–17

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ARTICLES

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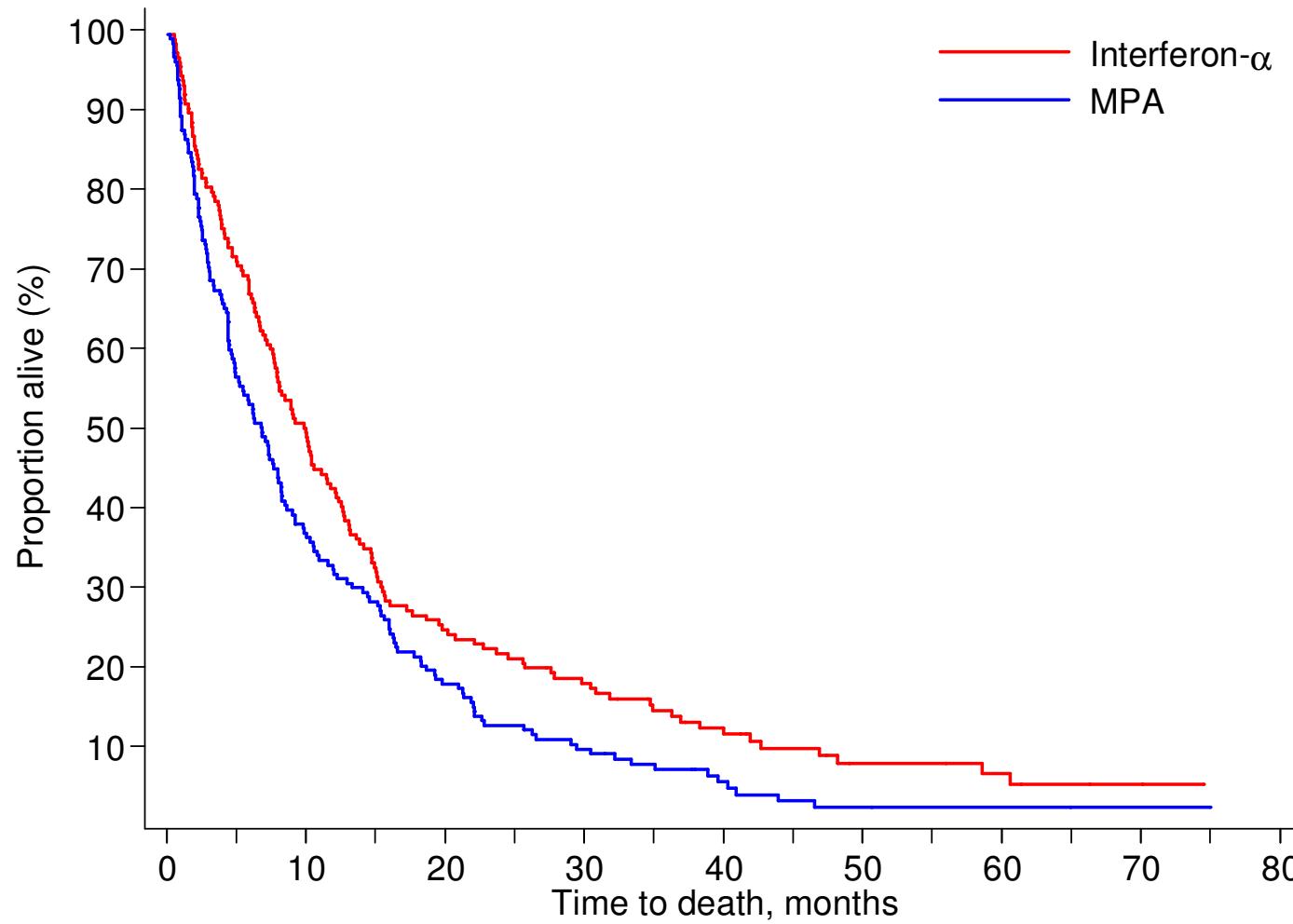
## **Interferon- $\alpha$ and survival in metastatic renal carcinoma: early results of a randomised controlled trial**

*Medical Research Council Renal Cancer Collaborators\**

# Kaplan-Meier



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```
. laplace month trt , failure(cens) q(10 50 90) reps(500)
```

Simultaneous Laplace regression				Number of obs = 347		
Optimization: Gradient Search						
bootstrap(500) SEs						
				q10 Log likelihood = -1.5581		
				q50 Log likelihood = -1.9044		
				q90 Log likelihood = -2.6071		
month	Observed Coef.	Bootstrap Std. Err.	t	P> t	Normal-based [95% Conf. Interval]	
q10						
trt	0.6	0.3	1.84	0.066	-0.0	1.2
_cons	1.0	0.2	5.67	0.000	0.6	1.3
q50						
trt	3.1	1.2	2.50	0.013	0.7	5.5
_cons	6.8	0.8	8.54	0.000	5.3	8.4
q90						
trt	11.7	5.9	1.99	0.047	0.1	23.2
_cons	29.5	4.2	6.98	0.000	21.2	37.8

- Ninety percent of patients assigned to interferon- $\alpha$  therapy have at least two more months to live

***10th percentile of survival is 2 month***

- Half of the patients assigned to interferon- $\alpha$  therapy live longer than ten months

***Median survival is 10 months***

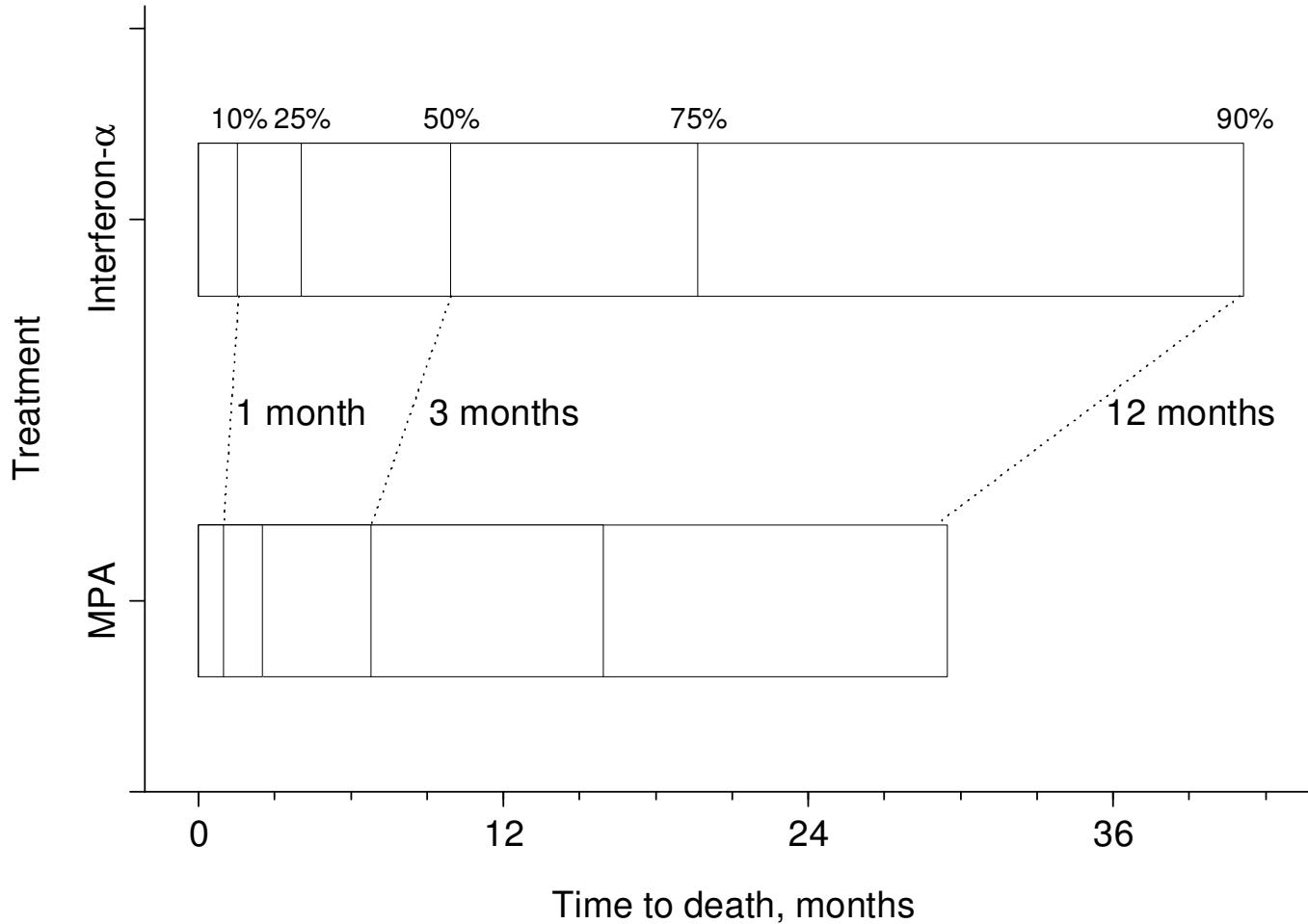
- Ten percent of the patients assigned to interferon- $\alpha$  therapy live longer than 41 months

***90th percentile of survival is 41 months***

# Tabulate percentiles

<b>Time to death, months</b>	<b>Interferon-<math>\alpha</math></b>	<b>MPA</b>	<b>Percentile Difference (95% CI)</b>	<b>P-value</b>
10th Percentile	1.6	1.0	0.6 (-0.0, 1.2)	0.066
Median	9.9	6.8	3.1 (0.7, 5.5)	0.013
90th Percentile	41.2	29.5	11.7 (0.1, 23.2)	0.047

# Plot percentiles



# Testing across percentiles

```
. test [q10]trt = [q50]trt
```

```
( 1) [q10]trt - [q50]trt = 0
```

```
F( 1, 345) = 4.51
```

```
Prob > F = 0.0343
```

The treatment effect on the 10th (1 month) and 50th (3 months) survival percentiles are significantly different.

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# Flexible modeling of predictors

We can investigate the change in survival percentiles according to a quantitative covariate using flexible tools (fractional polynomials, splines).

Let's consider white cell count ( $wcc$ ) as predictor of the median survival (or any other percentile).

# Laplace with restricted cubic splines

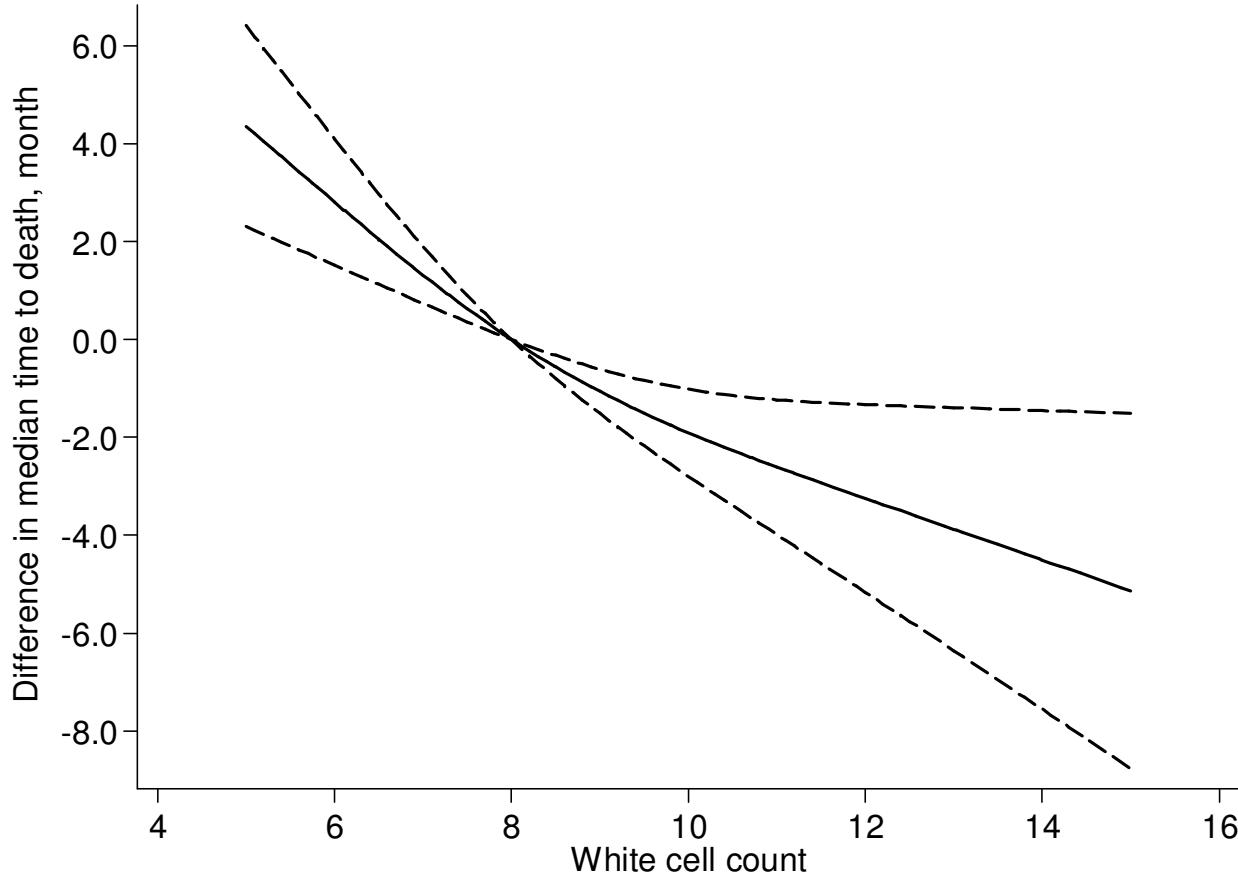
```
. mkspline wccs = wcc, nk(3) cubic
```

```
. laplace month trt wccs1 wccs2, failure(cens)
```

Laplace regression				Number of obs =			343
Optimization: Gradient Search				q50 Log likelihood =			-1.8687
month	Observed	Bootstrap		Normal-based			
	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]		
q50							
trt	2.3	1.1	2.20	0.029	0.2	4.4	
wccs1	-1.6	0.4	-3.94	0.000	-2.3	-0.8	
wccs2	0.8	0.4	1.79	0.075	-0.1	1.7	
_cons	19.1	3.3	5.78	0.000	12.6	25.6	

Some indication of departure from linearity for white cell count ( $p=0.075$ ).

```
. levelsof wcc if inrange(wcc, 5, 15)  
  
. xblc wccs1 wccs2 , cov(wcc) at(`r(levels)') line ///  
    ytitle("Difference in median time to death, month") ///  
    xtitle("White cell count") eq(q50) ref(8)
```



# Interactions

**British Journal of Cancer (2004) 90, 794–799**

Is treatment with interferon- $\alpha$  effective in all patients with metastatic renal carcinoma? A new approach to the investigation of interactions

**P Royston<sup>\*1</sup>, W Sauerbrei<sup>2</sup> and A Ritchie<sup>3</sup>**

# Laplace with restricted cubic splines and interactions

```
. gen inter1 = wccs1*trt
. gen inter2 = wccs2*trt
```

```
laplace month trt wccs1 wccs2 inter1 inter2 , failure(cens)
```

		Observed		Bootstrap		Normal-based	
		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
month	q50						
q50	trt	21.4	1.2	18.07	0.000	19.1	23.8
	wccs1	-0.0	0.3	-0.13	0.900	-0.7	0.6
	wccs2	-0.6	0.3	-1.78	0.075	-1.3	0.1
	inter1	-2.4	0.1	-27.28	0.000	-2.5	-2.2
	inter2	1.6	0.3	4.77	0.000	1.0	2.3
	_cons	7.8	2.5	3.10	0.002	2.9	12.8

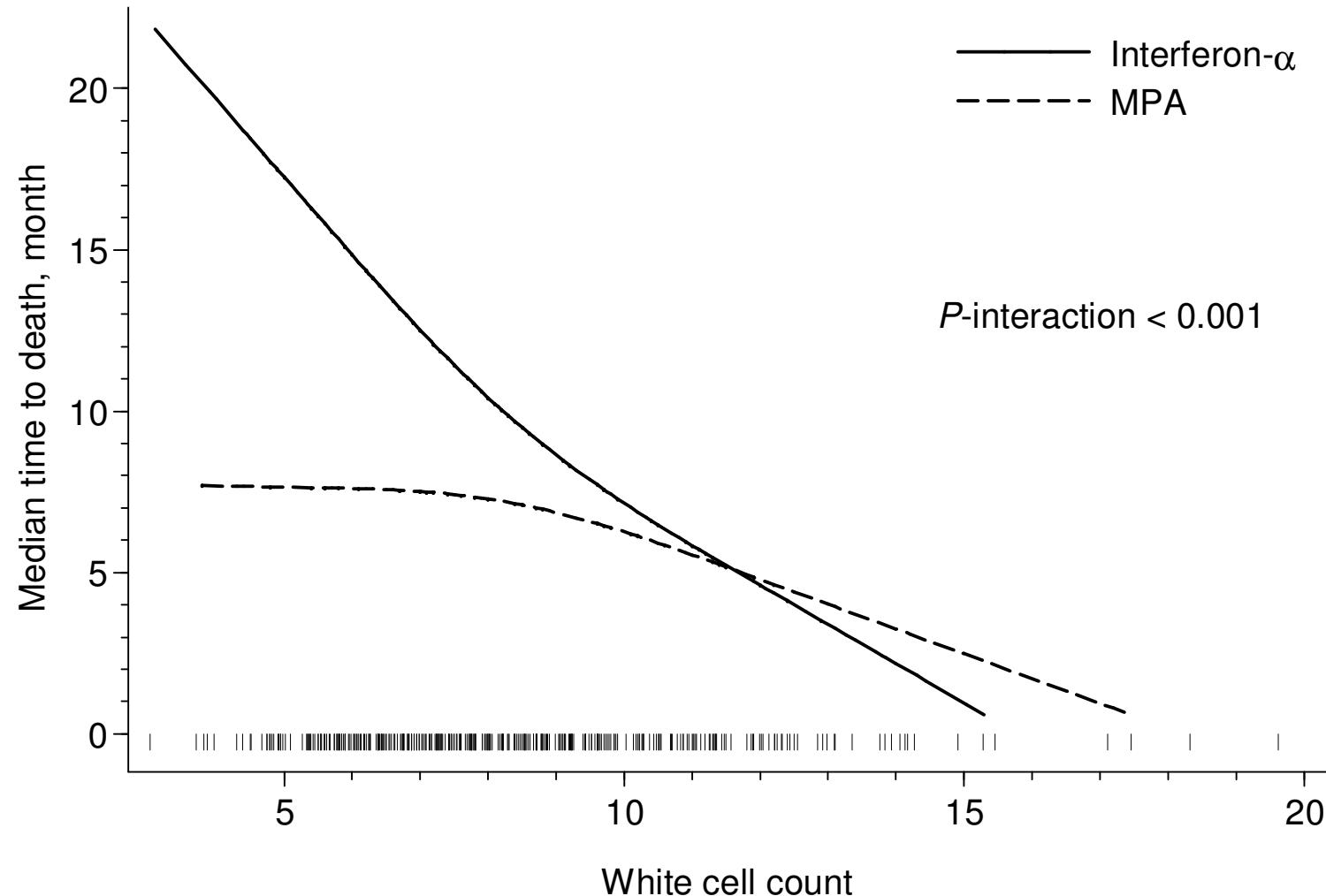
## Is the treatment effect dependent of white cell counts?

```
. testparm inter1 inter2
```

```
( 1) [q50]inter1 = 0
( 2) [q50]inter2 = 0
```

```
F(  2,    337) = 373.12
Prob > F = 0.0000
```

Yes ( $p<0.001$ ), based on a joint test that the two coefficients for interaction are equal to zero.



# Laplace regression with no censored data

- If there is no censoring ( $\delta_i = 1$  for all observations)
- Laplace likelihood estimator is simplified
- Laplace regression is equivalent to quantile regression

laplace = sqreg



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## Example 2

The NEW ENGLAND JOURNAL of MEDICINE

ORIGINAL ARTICLE

### Hyponatremia among Runners in the Boston Marathon

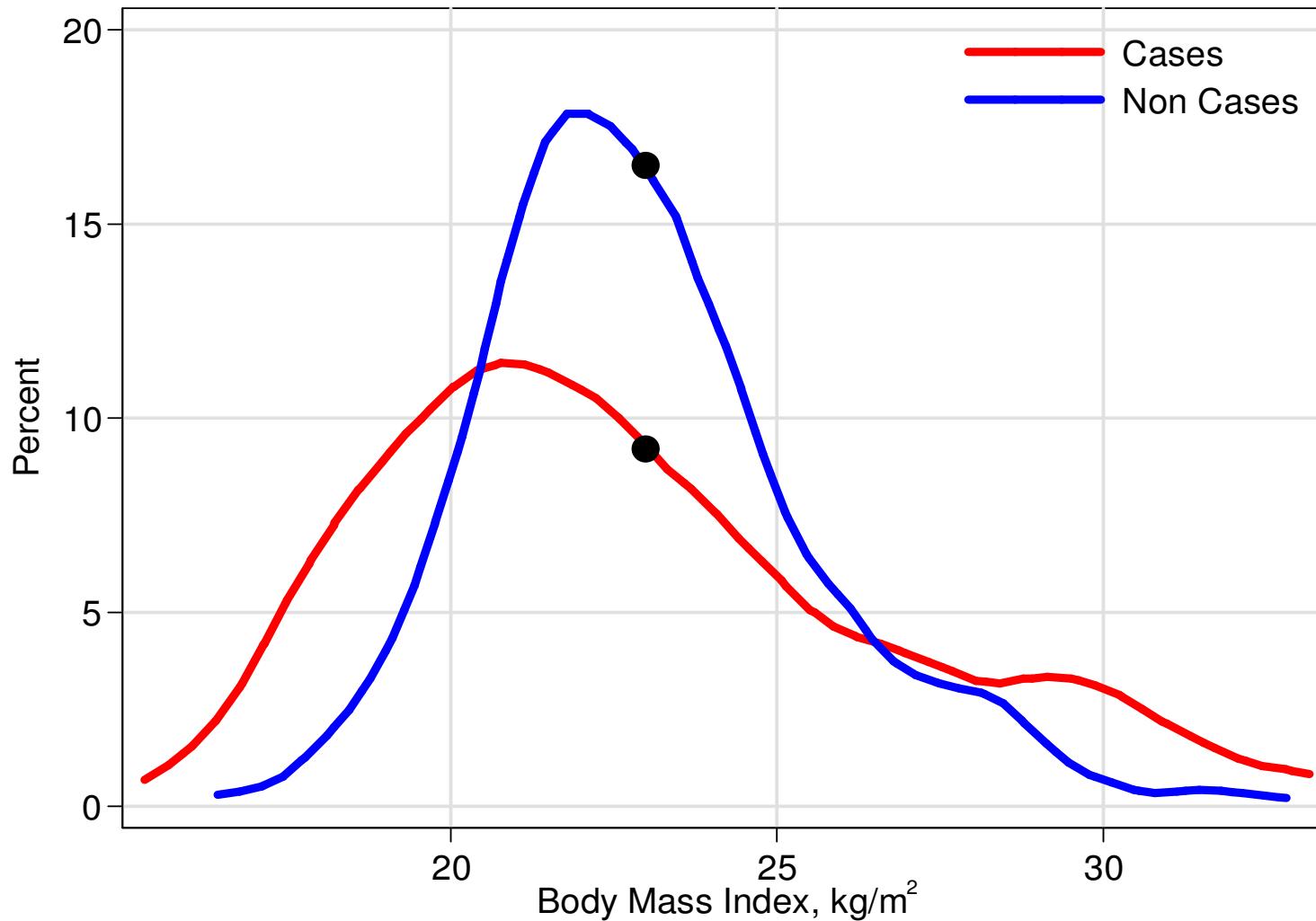
N ENGL J MED 352;15    WWW.NEJM.ORG    APRIL 14, 2005

**Table 2. Univariate and Multivariate Predictors of Hyponatremia.\***

Variable	Univariate Predictors		
	Hyponatremia (N=62)	No Hyponatremia (N=426)	P Value†
<b>Demographic characteristics</b>			
Body-mass index	22.8±3.7	23.0±2.5	0.68
Category of body-mass index			0.01
<20 (%)	25	8	—
20–25 (%)	54	73	—
>25 (%)	21	19	—

† For the univariate analysis, all continuous variables were analyzed with the use of t-tests.

# Histogram



. sqreg bmi nas135, reps(500) // 8 seconds

Simultaneous quantile regression  
 bootstrap(500) SEs

Number of obs =	465
.50 Pseudo R2 =	0.0014

	bmi	Bootstrap					
		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
q50							
	nas135	-0.27	0.61	-0.44	0.660	-1.48	0.94
	_cons	22.56	0.13	173.33	0.000	22.30	22.82

. laplace bmi nas135, reps(500) // 2 seconds

Laplace regression

Number of obs =	465
Optimization: Gradient Search	
bootstrap(500) SEs	
q50 Log likelihood =	0.6640

	bmi	Normal-based					
		Observed	Bootstrap	t	P> t	[95% Conf. Interval]	
q50							
	nas135	-0.27	0.60	-0.46	0.648	-1.45	0.91
	_cons	22.56	0.12	188.38	0.000	22.32	22.80

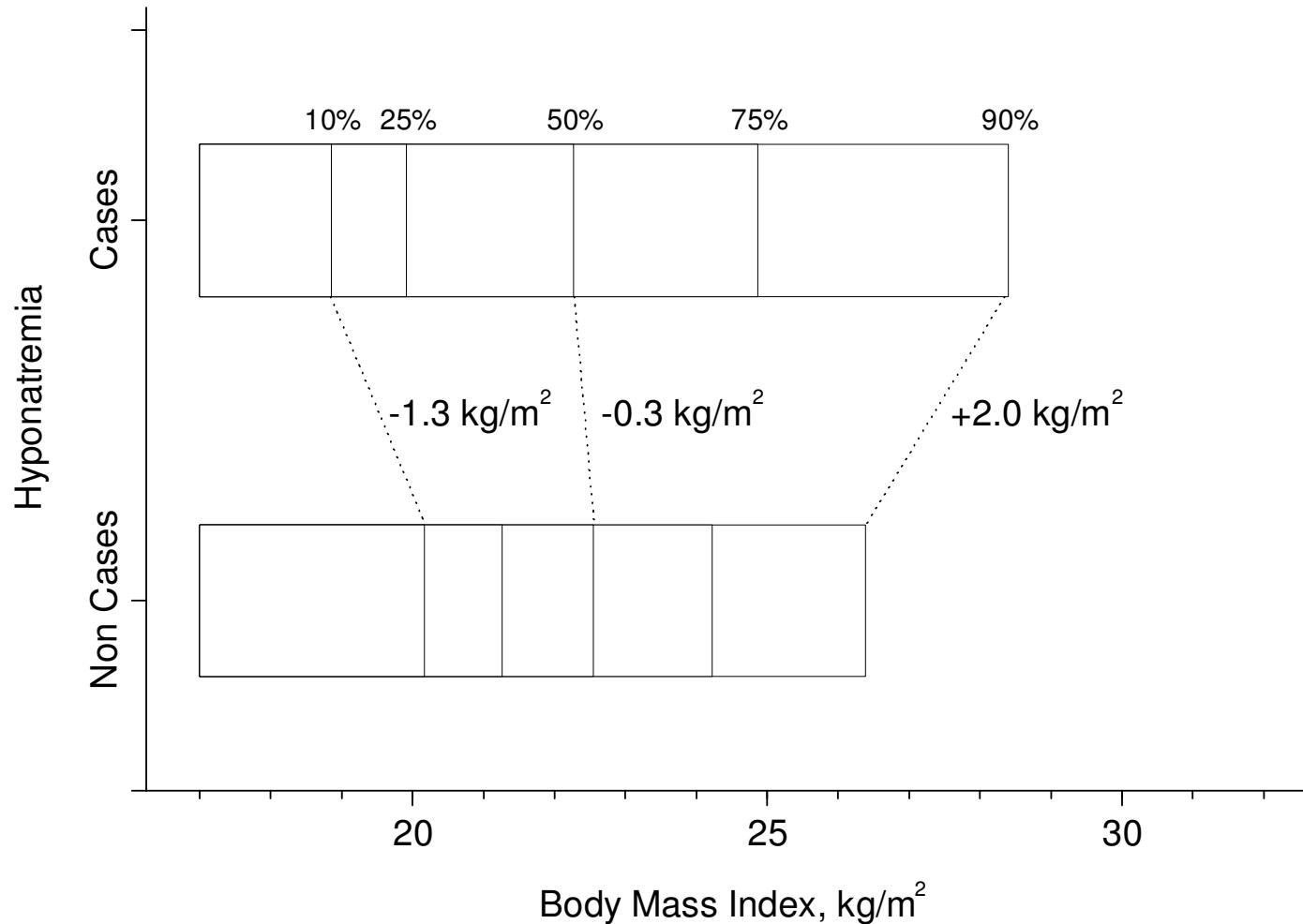
. laplace bmi nas135, q(10 50 90)

Simultaneous Laplace regression				Number of obs =			465
Optimization: Gradient Search				q10 Log likelihood =			0.2533
bootstrap(500) SEs				q50 Log likelihood =			0.6640
				q90 Log likelihood =			0.3569
bmi	Observed Coef.	Bootstrap Std. Err.	t	P> t	Normal-based [95% Conf. Interval]		
q10							
nas135	-1.33	0.42	-3.15	0.002	-2.15	-0.50	
_cons	20.18	0.15	130.82	0.000	19.88	20.48	
q50							
nas135	-0.27	0.60	-0.46	0.645	-1.44	0.89	
_cons	22.56	0.12	182.96	0.000	22.32	22.80	
q90							
nas135	2.02	0.98	2.05	0.040	0.09	3.95	
_cons	26.39	0.34	77.32	0.000	25.72	27.06	

# Tabulate percentiles

<b>BMI, kg/m<sup>2</sup></b>	<b>Hyponatremia (N=62)</b>	<b>No Hyponatremia (N=426)</b>	<b>Percentile Difference (95% CI)</b>	<b>P-value</b>
10th Percentile	18.9	20.2	-1.3 (-2.2, -0.5)	0.005
Median	22.3	22.6	-0.3 (-1.4, 0.9)	0.645
90th Percentile	28.4	26.4	2.0 (0.1, 4.0)	0.040

# Plot percentiles



## Other features

- **laplace** can model heteroschedasticity (error term dependent on covariates) using the option **sigma**(*varlist*)
- **laplace** can take on any kind of weights

# Summary

- **laplace** is a parametric model that estimates quantiles of a continuous response variable conditionally on covariates
- **laplace** is another way of analyzing censored data

## References

Bottai, M. and Zhang, J. (2010), Laplace regression with censored data. *Biometrical Journal*, 52: 487–503.

Orsini N., Bottai, M. (2011) Laplace regression. *Stata Journal*, in preparation.