Ensemble Learning Targeted Maximum Likelihood Estimation for Stata Users

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Table of Contents

- Background and notation
- ATE estimators
 - Estimators: Drawbacks
- 3 Targeted Maximum Likelihood Estimation
- Stata Implementation
 - Simulations
 - Links: online tutorial and GitHub open source eltmle
 - eltmle one sample simulation
- Next steps

References

- Additional material
 - Why care about TMLE
 - TMLE road map
 - Non-parametric theory and empirical efficiency: Influence Cy
 - Machine learning: ensemble learning



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Causal Inference Tutorial

Link to the tutorial

https://onlinelibrary.wiley.com/doi/10.1002/sim.9234?af=R





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https://github.com/migariane/TutorialComputationalCausalInferenceEstimators



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Stata ELTMLE

Syntax eltmle Stata command

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Stata Implementation: overall structure

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ELTMLE

Stata Implementation: R code for calling the SL

```
program tmle
// Write R Code dependencies: foreign Surperlearner
set more off
qui: file close all
qui: file open rcode using SLS.R, write replace
qui: file write rcode ///
                  "set.seed(123)"' newline ///
                  "list.of.packages <- c("foreign", "SuperLearner")"' newline ///
                  "new.packages <- list.of.packages[!(list.of.packages %in% installed.packages()[,"Package"])]"' newline ///
                  "if (length (new.packages)) install.packages (new.packages, repos='http://cran.us.r-project.org')" newline ///
                  "library(SuperLearner)"' newline ///
                  "library(foreign)"' newline ///
                  "data <- read.csv("data.csv", sep=",")"' newline ///
                  "attach(data)"' newline ///
                 "SL.library <- c("SL.glm", "SL.step", "SL.glm.interaction")"' newline ///
                  "n <- nrow(data)"' newline ///
                 "nvar <- dim(data)[[2]]"' newline ///
                  "Y <- data[,1]"' newline 7//
                  "A <- data[,2]"' _newline ///
                 "X <- data[,2:nvar]"' newline ///
                  "W <- data[,3:nvar]"' newline ///
                  "X1 <- X0 <- X"' newline ///
                 "X1[,1] <- 1"' newline ///
                  "X0[,1] <- 0"' newline ///
                  "newdata <- rbind(X,X1,X0)"' newline ///
                  "Q <- try(SuperLearner(Y = data[,1], X = X, SL.library=SL.library, family=binomial(), newX=newdata, method="met
                  "Q <- as.data.frame(Q[[4]])"' newline ///
                 "QAW <- Q[1:n,]"' newline ///
                  "Q1W <- Q[((n+1):(2*n)),]"' newline ///
                  "QOW <- Q[((2*n+1):(3*n)),]" newline ///
                  "g <- suppressWarnings(SuperLearner(Y = data[,2], X = W, SL.library = SL.library, family = binomial(), method =
                  "ps <- g[[4]]"' newline ///
                 "ps[ps<0.025] <- 0.025"' newline ///
"ps[ps>0.975] <- 0.975"' newline ///
                  "data <- cbind(data,QAW,Q1W,Q0W,ps,Y,A)"' newline ///
                  "write.dta(data, "data2.dta")"'
mui: file close rcode
                                                                                                                                                                                     b d link b d "sustained and also
also de la sustaine de distriction de la sustaine de la sustaine
Sustaine de la sustaine d
```

Stata Implementation: Batch file executing R

qui: file close rcode 114 // Write bacth file to find R.exe path and R version set more off 116 qui: file close all qui: file open bat using setup.bat, write replace 118 gui: file write bat /// 119 "@echo off"' newline /// "SET PATHROOT=C:\Program Files\R\"' newline /// "echo Locating path of R..."' newline /// "echo."' newline /// "if not exist "%PATHROOT%" goto:NO R"' newline /// 124 `"for /f "delims=" %%r in (' dir /b "%PATHROOT%R*" ') do ("' newline /// "echo Found %%r"' newline /// "echo shell "%PATHROOT%%%r\bin\x64\R.exe" CMD BATCH SLS.R > runr.do"' newline /// "echo All set!"' newline /// "goto:DONE"' newline /// 129 `")"' newline /// ":NO R"' newline /// "echo R is not installed in your system."' newline /// 132 "echo."' newline /// 133 "echo Download it from https://cran.r-project.org/bin/windows/base/"' newline /// 134 "echo Install it and re-run this script" newline /// ":DONE"' newline /// "echo."' newline /// 136 "pause" 138 qui: file close bat 139 140 //Run batch 141 shell setup.bat 142 //Run R 143 do runr.do 144 145 // Read Revised Data Back to Stata 146 clear 147 guietly: use "data2.dta", clear 148 149 // O to logit scale 150 gen logOAW = log(OAW / (1 - OAW))gen logOlW = log(OlW / (1 - OlW))gen logOOW = log(OOW / (1 - OOW))154 // Clever covariate HAW



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Output for continuous outcome

.use http://www.stata-press.com/data/r14/cattaneo2.dta .eltmle bweight mbsmoke mage medu prenatal mmarried, tmle										
Variable	Obs	Mean	Std. dev.	Min	Max					
POM1 POM0 ps	4,642 4,642 4,642	2832.69 3062.695 .1861267	74.9141 91.22898 .1106222	2550.819 2844.977 .0377472	2968.504 3177.975 .8479414					
TMLE: Average Treatment Effect										
ATE: -230 SE: 24 P-value: 0.00 95%CI: -277	.0 .5 00 .9, -182.1									
TMLE: Causal Risk Ratio (CRR)										
CRR: 0.93; 95%CI	:(0.91, 0.94									
TMLE: Marginal C	dds Ratio (M	10R)								
MOR: 0.83; 95%CI	:(0.80, 0.87	7)								

Output for continuous outcome and balance option

.eltmle bweight mbsmoke mage medu prenatal mmarried, tmle bal



Simulations comparing Stata ELTMLE vs R-TMLE





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ELTMLE

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Link to the tutorial

https://migariane.github.io/TMLE.nb.html



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Stata Implementation: source code



17/44

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https://github.com/migariane/eltmle

Stata installation and step by step commented syntax



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github install migariane/eltmle which eltmle viewsource eltmle.ado



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eltmle

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Statistics in Medicine tutorial: TMLE

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Next steps

• Stata Journal manuscript introducing eltmle.



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- Stata Journal manuscript introducing eltmle.
- Improved eltmle display and user interface.



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Thank YOU

THANK YOU FOR YOUR TIME



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Background: Potential Outcomes framework

Rubin and Heckman

- This framework was developed first by statisticians (Rubin, 1983) and econometricians (Heckman, 1978) as a new approach for the estimation of **causal effects** from observational data.
- We will keep separate the **causal framework** (a conceptual issue briefly introduce here) and the "**how to estimate causal effects**" (an statistical issue also introduced here)



Causal effects with OBSERVATIONAL data

ASSUMPTIONS for Identification

 Rosebaum & Rubin, 1983: The Ignorable Treatment Assignment (A.K.A Ignorability, Unconfoundeness or Conditional Mean Independence).

• POSITIVITY.

• SUTVA.



Causal effect with OBSERVATIONAL data

IGNORABILITY

$(\boldsymbol{Y}_i(1),\boldsymbol{Y}_i(0)) \bot \boldsymbol{A}_i \mid \boldsymbol{W}_i$

POSITIVITY

POSITIVITY: P(A = a | W) > 0 for all a, W

SUTVA

- We have assumed that there is only on version of the treatment (consistency) Y(1) if A = 1 and Y(0) if A = 0.
- The assignment to the treatment to one unit doesn't affect the outcome of another unit (no interference) or IID random variables.
- The model used to estimate the assignment probability has to **be correctly specified**.

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Causal effect

Potential Outcomes

We only observe:

$$Y_i(1) = Y_i(A = 1)$$
 and $Y_i(0) = Y_i(A = 0)$

However we would like to know what would have happened if:

Treated $Y_i(1)$ would have been non-treated $Y_i(A = 0) = Y_i(0)$.

Controls $Y_i(0)$ would have been treated $Y_i(A = 1) = Y_i(1)$.

Identifiability

- How we can identify the effect of the potential outcomes **Y**^a if they are not observed?
- How we can estimate the expected difference between the potential outcomes E[Y(1) - Y(0)], namely the ATE.

G-Formula for the identification of the ATE with observational data

$$E(Y^{a}) = \sum_{y} E(Y^{a} \mid W = w)P(W = w)$$

=
$$\sum_{y} E(Y^{a} \mid A = a, W = w)P(W = w)$$
 by consistency
=
$$\sum_{y} E(Y = y \mid A = a, W = w)P(W = w)$$
 by ignorability

The ATE=

$$\sum_{\mathbf{w}} \left[\sum_{\mathbf{y}} \mathbf{P}(\mathbf{Y} = \mathbf{y} \mid \mathbf{A} = \mathbf{1}, \mathbf{W} = \mathbf{w}) - \sum_{\mathbf{y}} \mathbf{P}(\mathbf{Y} = \mathbf{y} \mid \mathbf{A} = \mathbf{0}, \mathbf{W} = \mathbf{w}) \right] \mathbf{P}(\mathbf{W} = \mathbf{w})$$
$$P(W = w) = \sum_{\mathbf{y}} P(W = w, \mathbf{A} = \mathbf{a}, \mathbf{Y} = \mathbf{y})$$

y,a

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MEDICINE

G-Formula, (Robins, 1986)

G-Formula for the identification of the ATE with observational data The ATE=

$$\sum_{\mathbf{w}} \left[\sum_{\mathbf{y}} \mathbf{P}(\mathbf{Y} = \mathbf{y} \mid \mathbf{A} = \mathbf{1}, \mathbf{W} = \mathbf{w}) - \sum_{\mathbf{y}} \mathbf{P}(\mathbf{Y} = \mathbf{y} \mid \mathbf{A} = \mathbf{0}, \mathbf{W} = \mathbf{w}) \right] \mathbf{P}(\mathbf{W} = \mathbf{w})$$

$$P(W = w) = \sum_{y,a} P(W = w, A = a, Y = y)$$

G-Formula

- The sums is generic notation. In reality, likely involves sums and integrals (we are just integrating out the W's).
- The **g-formula** is a **generalization of standardization** and allow to estimate unbiased treatment effect estimates.

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Regression-adjustment

$$\widehat{ATE}_{RA} = N^{-1} \sum_{i=1}^{N} [E(Y_i \mid A = 1, W_i) - E(Y_i \mid A = 0, W_i)]$$

$$m_{A}(w_{i}) = E(Y_{i} \mid A_{i} = A, W_{i})$$

$$\widehat{ATE}_{RA} = N^{-1} \sum_{i=1}^{N} \left[\hat{m}_1(w_i) - \hat{m}_0(w_i) \right]$$

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IPTW

IPTW (Inverse probability treatment weighting)

Survey theory (Horvitz-Thompson)

$$\hat{\mathsf{P}}_i = E(A_i \mid W_i)$$
; So, $\frac{1}{\hat{\mathsf{p}}_i}$, if A = 1 and, $\frac{1}{(1 - \hat{\mathsf{p}}_i)}$, if A = 0

over the total number of individuals

$$\widehat{ATE}_{IPTW} = N^{-1} \sum_{i=1}^{N} \frac{A_i Y_i}{\hat{p}_i} - N^{-1} \sum_{i=1}^{N} \frac{(1 - A_i) Y_i}{(1 - \hat{p}_i)}$$

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AIPTW

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AIPTW (Augmented Inverse probability treatment weighting)

Solving Estimating Equations

$$\widehat{TE}_{AIPTW} = N^{-1} \sum_{i=1}^{N} \left[(Y(1) \mid A_i = 1, W_i) - (Y(0) \mid A_i = 0, W_i) \right] + N^{-1} \sum_{i=1}^{N} \left(\frac{(A_i = 1)}{P(A_i = 1 \mid W_i)} - \frac{(A_i = 0)}{P(A_i = 0 \mid W_i)} \right) \left[Y_i - E(Y \mid A_i, W_i) \right]$$



Targeted learning

Springer Series in Statistics

Targeted Learning

Causal Inference for Observational and Experimental Data

🙆 Springer

Source: Mark van der Laan and Sherri Rose. Targeted learning: causal inference for observational and experimental data. Springer Series in Statistics, 2011.



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Why Targeted learning?



Source: Mark van der Laan and Sherri Rose. Targeted learning: causal inference for observational and experimental data. Springer Series in Statistics, 2011.



TMLE ROAD MAP

MC simulations: Luque-Fernandez et al, 2017 (in press, American Journal of Epidemiology)

	ATE		BIAS (%)		RMSE		95%CI coverage (%)	
	N=1,000	N=10,000	N=1,000	N=10,000	N=1,000	N=10,000	N=1,000	N=10,000
First scenario* (correctly specified models)								
True ATE	-0.1813							
Naïve	-0.2234	-0.2218	23.2	22.3	0.0575	0.0423	77	89
AIPTW	-0.1843	-0.1848	1.6	1.9	0.0534	0.0180	93	94
IPTW-RA	-0.1831	-0.1838	1.0	1.4	0.0500	0.0174	91	95
TMLE	-0.1832	-0.1821	1.0	0.4	0.0482	0.0158	95	95
Second scenario ** (misspecified models)								
True ATE	-0.1172							
Naïve	-0.0127	-0.0121	89.2	89.7	0.1470	0.1100	0	0
BFit AIPTW	-0.1155	-0.0920	1.5	11.7	0.0928	0.0773	65	65
BFit IPTW-RA	-0.1268	-0.1192	8.2	1.7	0.0442	0.0305	52	73
TMLE	-0.1181	-0.1177	0.8	0.4	0.0281	0.0107	93	95

*First scenario : correctly specified models and near-positivity violation

**Second scenario: misspecification, near-positivity violation and adaptive model selection



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TMLE ROAD MAP



TMLE steps

ELTMLE

October 20, 2022 36/44

Substitution estimation: $\hat{E}(Y | A, W)$



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October 20, 2022

Substitution estimation: $\hat{E}(Y \mid A, W)$

First compute the outcome regression E(Y | A, W) using the Super-Learner to then derive the Potential Outcomes and compute Ψ⁽⁰⁾ = E(Y(1) | A = 1, W) - E(Y(0) | A = 0, W).



Substitution estimation: $\hat{E}(Y \mid A, W)$

- First compute the outcome regression E(Y | A, W) using the Super-Learner to then derive the Potential Outcomes and compute Ψ⁽⁰⁾ = E(Y(1) | A = 1, W) E(Y(0) | A = 0, W).
- Estimate the exposure mechanism P(A=1|,W) using the Super-Learner to predict the values of the propensity score.



Substitution estimation: $\hat{E}(Y \mid A, W)$

- First compute the outcome regression $\mathbf{E}(\mathbf{Y} | \mathbf{A}, \mathbf{W})$ using the **Super-Learner** to then derive the Potential Outcomes and compute $\Psi^{(0)} = \mathbf{E}(Y(1) | \mathbf{A} = 1, \mathbf{W}) \mathbf{E}(Y(0) | \mathbf{A} = 0, \mathbf{W}).$
- Estimate the exposure mechanism P(A=1|,W) using the Super-Learner to predict the values of the propensity score.
- Compute **HAW** = $\left(\frac{\mathbb{I}(A_i=1)}{P(A_i=1|W_i)} \frac{\mathbb{I}(A_i=0)}{P(A_i=0|W_i)}\right)$ for each individual, named the **clever covariate H**.



Substitution estimation: $\hat{E}(Y \mid A, W)$

- First compute the outcome regression $\mathbf{E}(\mathbf{Y} | \mathbf{A}, \mathbf{W})$ using the **Super-Learner** to then derive the Potential Outcomes and compute $\Psi^{(0)} = \mathbf{E}(Y(1) | \mathbf{A} = 1, \mathbf{W}) \mathbf{E}(Y(0) | \mathbf{A} = 0, \mathbf{W}).$
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Fluctuation step $(\hat{\epsilon}_0, \hat{\epsilon}_1)$

 Update Ψ⁽⁰⁾ through a fluctuation step incorporating the information from the exposure mechanism:

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Fluctuation step $(\hat{\epsilon}_0, \hat{\epsilon}_1)$

 Update Ψ⁽⁰⁾ through a fluctuation step incorporating the information from the exposure mechanism:

$$\mathbf{H(1)W} = rac{\mathbb{I}(A_i=1)}{\hat{P}(A_i=1|W_i)}$$
 and,

Fluctuation step $(\hat{\epsilon}_0, \hat{\epsilon}_1)$

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$$\mathbf{H(1)W} = \frac{\mathbb{I}(A_i=1)}{\hat{P}(A_i=1|W_i)} \text{ and,} \mathbf{H(0)W} = -\frac{\mathbb{I}(A_i=0)}{\hat{P}(A_i=0|W_i)}$$

• • • • • • • • •

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 This step aims to reduce bias minimising the mean squared error (MSE) for (Ψ) and considering the bounds of the limits of Y.

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- This step aims to reduce bias minimising the mean squared error (MSE) for (Ψ) and considering the bounds of the limits of Y.
- The fluctuation parameters (\(\{\eta_0\)}, \(\{\eta_1\)}\)) are estimated using maximum likelihood procedures (in Stata):
 - . glm Y HAW, fam(binomial) nocons offset(E(Y| A, W))

38/44

Image: A matrix

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Targeted estimate of the ATE $(\widehat{\Psi})$

 $\Psi^{(0)}$ update using ϵ (epsilon)



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Targeted estimate of the ATE $(\widehat{\Psi})$

 $\Psi^{(0)}$ update using ϵ (epsilon)

 $\mathbf{E}^{*}(Y \mid A = 1, W) = \operatorname{expit}\left[\operatorname{logit}\left[E(Y \mid A = 1, W)\right] + \hat{\epsilon_{1}}H_{1}(1, W)\right]$

 $\mathbf{E}^{*}(Y \mid A = 0, W) = \operatorname{expit}\left[\operatorname{logit}\left[E(Y \mid A = 0, W)\right] + \hat{\epsilon_{0}}H_{0}(0, W)\right]$

Targeted estimate of the ATE from $\Psi^{(0)}$ to $\Psi^{(1)}$: $(\widehat{\Psi})$



Targeted estimate of the ATE $(\widehat{\Psi})$

 $\Psi^{(0)}$ update using ϵ (epsilon)

 $\mathbf{E}^{*}(Y \mid A = 1, W) = \text{expit} [\text{logit} [E(Y \mid A = 1, W)] + \hat{\epsilon}_{1}H_{1}(1, W)]$

 $\mathbf{E}^{*}(Y \mid \mathbf{A} = \mathbf{0}, \mathbf{W}) = \operatorname{expit}\left[\operatorname{logit}\left[E(Y \mid \mathbf{A} = \mathbf{0}, \mathbf{W})\right] + \hat{\epsilon_{\mathbf{0}}}H_{0}(\mathbf{0}, \mathbf{W})\right]$

Targeted estimate of the ATE from $\Psi^{(0)}$ to $\Psi^{(1)}$: $(\widehat{\Psi})$

 $\Psi^{(1)}: \hat{\Psi} = [\mathbf{E}^*(Y(1) \mid A = 1, W) - \mathbf{E}^*(Y(0) \mid A = 0, W)]$



TMLE inference



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TMLE inference



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TMLE inference

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IC: Geometric interpretation





Nonparametric Delta Method : E(x - μ)² Infinitesimal Jackknife

Estimate of the ψ Standard Error using the efficient Influence Curve. Image credit: Miguel Angel Luque-Fernandez

41/44

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Targeted learning



Source: Mark van der Laan and Sherri Rose. Targeted learning: causal inference for observational and experimental data. Springer Series in Statistics, 2011.



Super-Learner: Ensemble learning



To apply the **EIC** we need data-adaptive estimation for both, the model of the outcome, and the model of the treatment.

Asymptotically, the final weighted combination of algorithms (Super Learner) performs as well as or better than the best-fitting algorithm (van der Laan, 2007).



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An estimator is asymptotically linear with influence function φ (IC) if the estimator can be approximate by an empirical average in the sense that

$$(\hat{\theta} - \theta_0) = \frac{1}{n} \sum_{i=1}^n (IC) + Op(1/\sqrt{n})$$

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