

Regression Control Method with Stata

回归控制法及Stata应用

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1. Introduction

- Regression control method (RCM)
 - Aka a panel data approach for program evaluation (Hsiao et al. 2012)
 - Exploits cross-sectional correlation to construct counterfactual outcomes for a single treated unit
 - Hsiao and Zhou (2019) propose to add covariates in model for helping prediction

1. Introduction

- Stata command `rcm` :
 - Implement regression control method
 - Facilitate model selection and estimation
 - Support to add covariates
 - Support placebo tests for statistical inferences

2. Model

2.1 Basic Model

- N cross-sectional units indexed by $i = 1, \dots, N$ over $t = 1, \dots, T_0 + 1, \dots, T$ periods
- $i = 1$ indexes the treated unit, whereas $\{2, \dots, N\}$ is the index set of $N - 1$ control units (donor pool)
- Policy intervention occurs at time $T_0 + 1$, which partitions the time series into two sections:
 - Pre-treatment periods: $1, \dots, T_0$
 - Post-treatment periods: $T_0 + 1, \dots, T$

2.1 Basic Model

- y_{it}^1 and y_{it}^0 be the outcomes of the unit i in period t with and without intervention respectively
- y_{it} is observed that in the form:

$$y_{it} = d_{it}y_{it}^1 + (1 - d_{it})y_{it}^0$$

- $d_{it} = 1$ if $i = 1$ and unit i is under intervention in period t , or $d_{it} = 0$ if not
- The treatment effect for can be expressed as

$$\Delta_{it} = y_{it}^1 - y_{it}^0 \quad t = T_0 + 1, \dots, T$$

2.1 Basic Model

- Assume that y_{it}^0 is generated by a **pure** factor model of the form

$$y_{it}^0 = \mathbf{b}'_i \mathbf{f}_t + \varepsilon_{it}$$

- f_t : $K \times 1$ (unobserved) common factors
- \mathbf{b}'_i : $1 \times K$ (unobserved) factor loadings
- ε_{it} : Random idiosyncratic component with $E(\varepsilon_{it}) = 0$.

2.1 Basic Model

Stack y_{it}^0 for $i \in \{2, \dots, N\}$, we get

$$\tilde{\mathbf{y}}_t = \tilde{\mathbf{B}}\mathbf{f}_t + \tilde{\boldsymbol{\varepsilon}}_t$$

- $\tilde{\mathbf{y}}_t = (y_{2t}^0, \dots, y_{Nt}^0)'$
- $\tilde{\boldsymbol{\varepsilon}}_t = (\varepsilon_{2t}, \dots, \varepsilon_{Nt})'$
- $\tilde{\mathbf{B}}$: $(N - 1) \times K$ factor loading matrix $(\mathbf{b}_2, \dots, \mathbf{b}_N)'$

$$\mathbf{f}_t = \left(\tilde{\mathbf{B}}' \tilde{\mathbf{B}} \right)^{-1} \tilde{\mathbf{B}}' (\tilde{\mathbf{y}}_t - \tilde{\boldsymbol{\varepsilon}}_t)$$

2.1 Basic Model

Then y_{1t}^0 can be represented as

$$\begin{aligned}y_{1t}^0 &= \mathbf{b}'_1 \mathbf{f}_t + \varepsilon_{1t} \\ &= \mathbf{b}'_1 \left(\tilde{\mathbf{B}}' \tilde{\mathbf{B}} \right)^{-1} \tilde{\mathbf{B}}' (\tilde{\mathbf{y}}_t - \tilde{\boldsymbol{\varepsilon}}_t) + \varepsilon_{1t} \\ &= \boldsymbol{\gamma}' \tilde{\mathbf{y}}_t + \varepsilon_{1t} - \boldsymbol{\gamma}' \tilde{\boldsymbol{\varepsilon}}_t\end{aligned}$$

- $\boldsymbol{\gamma}' = \mathbf{b}'_1 \left(\tilde{\mathbf{B}}' \tilde{\mathbf{B}} \right)^{-1} \tilde{\mathbf{B}}'$
- $\tilde{\mathbf{y}}_t$ is correlated to $\varepsilon_{1t} - \boldsymbol{\gamma}' \tilde{\boldsymbol{\varepsilon}}_t$

2.1 Basic Model

$$\begin{aligned}y_{1t}^0 &= \gamma' \tilde{\mathbf{y}}_t + \varepsilon_{1t} - \gamma' \tilde{\boldsymbol{\varepsilon}}_t \\ &= \gamma' \tilde{\mathbf{y}}_t + L(\varepsilon_{1t} - \gamma' \tilde{\boldsymbol{\varepsilon}}_t | \tilde{\mathbf{y}}_t) + v_{1t} \\ &= \gamma' \tilde{\mathbf{y}}_t + c_1 + \mathbf{c}' \tilde{\mathbf{y}}_t + v_{1t}\end{aligned}$$

- $L(\varepsilon_{1t} - \gamma' \tilde{\boldsymbol{\varepsilon}}_t | \tilde{\mathbf{y}}_t)$ is the linear projection of $\varepsilon_{1t} - \gamma' \tilde{\boldsymbol{\varepsilon}}_t$ onto $(1, \tilde{\mathbf{y}}_t')$, and c_1 and \mathbf{c} are the minimizers of

$$\min_{c_1, \mathbf{c}} E \left[(\varepsilon_{1t} - \gamma' \tilde{\boldsymbol{\varepsilon}}_t - c_1 - \mathbf{c}' \tilde{\mathbf{y}}_t)^2 \right]$$

2.1 Basic Model

$$\begin{aligned}y_{1t}^0 &= \boldsymbol{\gamma}' \tilde{\mathbf{y}}_t + c_1 + \mathbf{c}' \tilde{\mathbf{y}}_t + v_{1t} \\ &= \delta_1 + \boldsymbol{\delta}' \tilde{\mathbf{y}}_t + v_{1t}\end{aligned}$$

- $\delta_1 = c_1, \boldsymbol{\delta}' = \boldsymbol{\gamma}' + \mathbf{c}'$
- Hsiao et al. (2012) advocate estimating $\hat{\delta}_1$ and $\hat{\boldsymbol{\delta}}'$ by OLS with the pre-treatment subsample, and predicting the counterfactual outcomes as

$$\hat{y}_{1t}^0 = \hat{\delta}_1 + \hat{\boldsymbol{\delta}}' \tilde{\mathbf{y}}_t$$

2.1 Basic Model

- So the treatment effect Δ_{1t} can be predicted using

$$\hat{\Delta}_{1t} = y_{1t}^1 - \hat{y}_{1t}^0 \quad t = T_0 + 1, \dots, T$$

- Average treatment effect (ATE) is estimated by averaging $\hat{\Delta}_{1t}$ s over the post-treatment periods:

$$\hat{\Delta}_1 = \frac{1}{T - T_0} \sum_{t=T_0+1}^T \hat{\Delta}_{1t}$$

2.1 Basic Model

- Use all of control units for estimation may not be the best choice
- Large estimation variance in turn leads to poor out-of-sample predictions
- Hsiao et al. (2012) suggest using information criterion approach to select control units in exhaustive search (best subset selection)
- Use $\tilde{\mathbf{y}}_t^*$ instead of $\tilde{\mathbf{y}}_t$, where $\tilde{\mathbf{y}}_t^*$ is the best subset of $(y_{2t}, \dots, y_{Nt})'$

2.2 Model with Covariates (Hsiao and Zhou, 2019)

- Assume y_{it}^0 is a function of p observable variables \mathbf{x}_{it} :

$$y_{it}^0 = \mathbf{x}_{it}'\boldsymbol{\beta} + \mathbf{b}_i'\mathbf{f}_t + \varepsilon_{it}$$

- y_{1t}^0 can be predicted as

$$\hat{y}_{1t}^0 = \hat{\delta}_1 + \hat{\boldsymbol{\delta}}'\mathbf{z}_t^*$$

- \mathbf{z}_t^* includes any subset of $\mathbf{z}_t = (y_{2t}, \dots, y_{Nt}, \mathbf{x}_{1t}, \dots, \mathbf{x}_{Nt})'$ that helps to predict y_{1t}^0
(forward stepwise method or Lasso method suggested)

3 Extension

- Step 1 - Select the Suboptimal Model
Select the suboptimal model $M_{OLS}^*(j)$ using best subset selection, forward stepwise selection or backward stepwise selection, or $M_{lasso}^*(\lambda)$ using Lasso selection ($j \in \{1, \dots, T_0 - 1\}$)
- Step 2 - Select the Optimal Model
Choose the optimal model M^* from all of suboptimal models in terms of information criterion or cross-validation
- After model selection, OLS or Lasso regression is used to fit the optimal model for counterfactual prediction

3.1 Select the Suboptimal Model

3.1.1 Best Subset Selection

- Hsiao et. al. (2012) suggest using best subset selection

- *exhaustion*:

For $k = 1, 2, \dots, p$:

1. Fit all $\binom{p}{k}$ models that contain k predictors.

2. Choose the smallest *RSS* model as $M_{OLS}^*(k)$

- 2^p possible combination of the p predictors

3.1.1 Best Subset Selection

- *leaps and bounds*: quickly calculate best subsets without examining all possible subsets
- Fundamental inequality:

$$RSS(A) \leq RSS(B) \quad B \subset A$$

- If $RSS(\{3, 4\}) \geq RSS(1)$, $RSS(\{3\})$ and $RSS(\{4\})$ do not need to be calculated

3.1.1 Best Subset Selection

- **Initial** : Reorder the variables by their impact on RSS
- Regression and bound tree (pair tree)
- Traverse all the subsets of the root node in level 1
- Traverse the tree from right to left

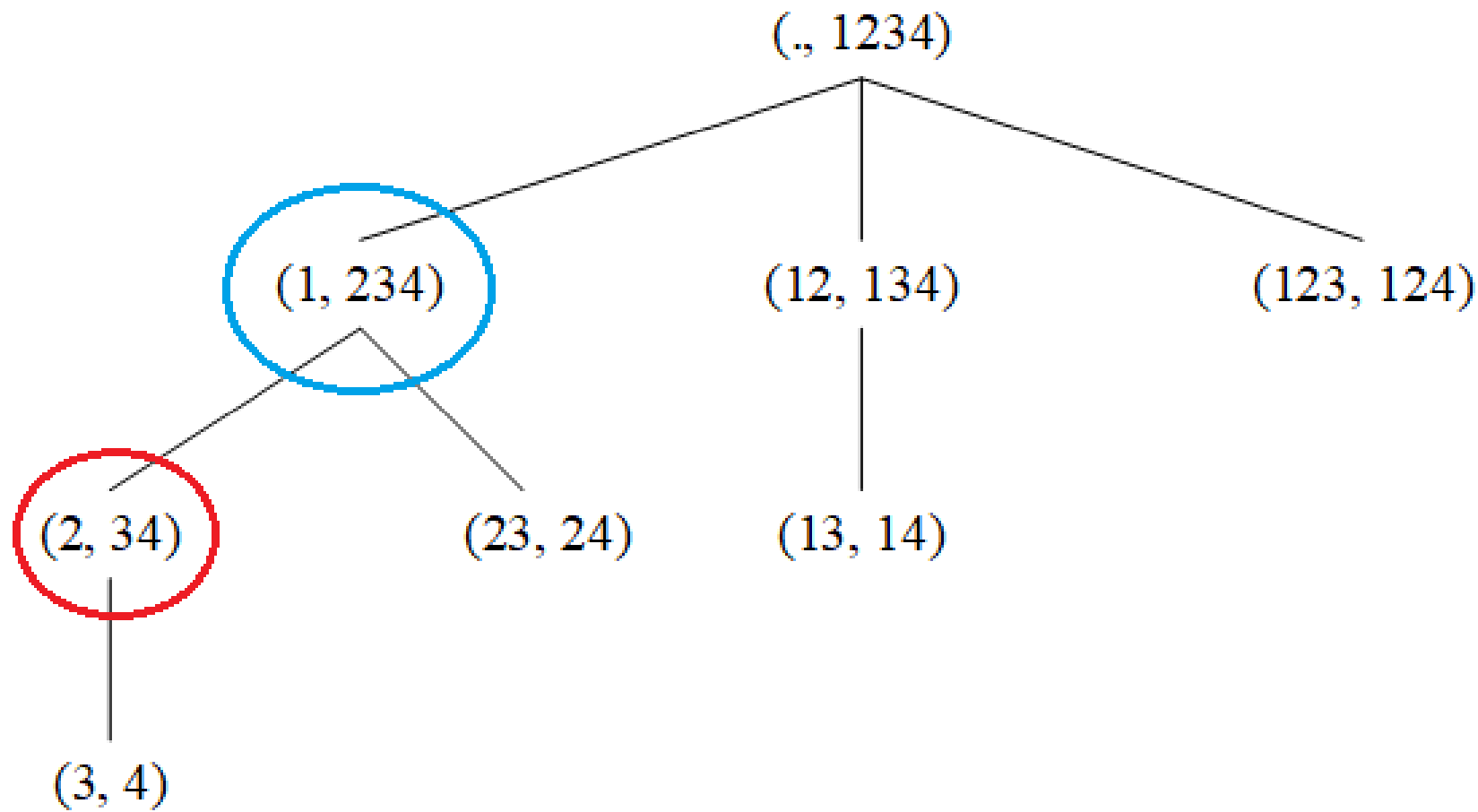


Figure : The pair tree

3.1.1 Best Subset Selection

- Suppose subset $\Omega \subset \{1, 2, \dots, p + 1\}$, we have

$$\text{RSS}(\Omega) = y'y - (\Phi'_{\Omega} \cdot y)' (\Phi'_{\Omega} \Phi_{\Omega})^{-1} (\Phi'_{\Omega} \cdot y)$$

- Φ is $T_0 \times (p + 1)$ matrix with a constant column and predictors
- Precomputed matrix :

$$(y, \Phi)'(y, \Phi) = \begin{pmatrix} y'y & y'\Phi \\ \Phi'y & \Phi'\Phi \end{pmatrix}$$

3.1.2 Forward Stepwise Selection

- Hsiao and Zhou (2019) suggest using forward stepwise selection
- $M_{OLS}^*(0)$ denote the smallest model, which contains no predictors
- For $k = 0, \dots, p - 1$:
 1. Consider models that augment the predictors in $M_{OLS}^*(k)$ with one additional predictor, fit OLS regression
 2. Fit these models with OLS regression
 3. Choose the smallest *RSS* model as $M_{OLS}^*(k + 1)$
- Shi and Huang (2021) : Forward-Selected PDA (fsPDa)

3.1.3 Backward Stepwise Selection

- $M_{OLS}^*(p)$ denote the largest possible model, which contains all p predictors
- For $k = p, \dots, 1$:
 1. Consider models that contain all but one of the predictors in $M_{OLS}^*(k)$
 2. Fit these models with OLS regression
 3. Choose the smallest *RSS* model as $M_{OLS}^*(k - 1)$
- Require that $T_0 > p$ (the largest possible model can be fit)

3.1.4 Lasso Selection

- Lasso selects β to minimize:

$$\hat{\beta}_{lasso}(\lambda) = \arg \min_{\beta} \sum_{i=1}^n (y_i - \mathbf{x}'_i \beta)^2 + \lambda \sum_{j=1}^p |\beta_j|$$

- $\lambda \geq 0$ is a tuning parameter
- Two extreme cases : $\hat{\beta}_{lasso}(\lambda = 0) = \hat{\beta}_{ols}$, $\hat{\beta}_{lasso}(\lambda = \infty) = \mathbf{0}$

3.1.4 Lasso Selection

- For $\lambda_l = \lambda_1, \dots, \lambda_L$:

1. Calculate lasso estimator $\hat{\beta}_{lasso}(\lambda_l)$

2. Consider the model with a set of predictors $\Omega(\lambda_l)$, where

$$\Omega(\lambda_l) = \{predictor_j | j \in \{1, \dots, p\}, \beta_{lasso,j}(\lambda_l) \neq 0\}$$

3. Choose the model as $M_{lasso}^*(\lambda_l)$ if $\Omega(\lambda_l)$ is different from $\Omega(\lambda_1), \dots, \Omega(\lambda_{l-1})$

3.2 Select Optimal Model

3.2.1 Information criterion

$$\text{AIC}(p) = T_0 \ln \left(\frac{\mathbf{e}'_0 \mathbf{e}_0}{T_0} \right) + 2(p + 2)$$

$$\text{BIC}(p) = T_0 \ln \left(\frac{\mathbf{e}'_0 \mathbf{e}_0}{T_0} \right) + (p + 2) \ln(T_0)$$

$$\text{AICc}(p) = \text{AIC}(p) + \frac{2(p + 2)(p + 3)}{T_1 - (p + 1) - 2}$$

- \mathbf{e}'_0 : OLS or Lasso residuals fitted in pre-treatment periods.

3.2.2 Cross-Validation

- Common practice after Lasso selection
- Optimize the out-of-sample prediction performance
- The mean squared error for each fold is computed as

$$\text{MSE}(\lambda, k) = \frac{1}{n_k} \sum_{i \in \mathcal{F}_k} \left(y_i - \mathbf{x}'_i \hat{\boldsymbol{\beta}}_{\text{lasso},k}(\lambda) \right)^2$$

- K : Number of groups in which pre-treatment data is splitted
- n_k : Size of pre-treatment data partition k for $k = 1, \dots, K$
- \mathcal{F}_k : Set of observations in k -fold

3.2.2 Cross-Validation

- K -fold Cross-Validation estimate of **CV MSE**, which serves as a measure of prediction performance, is

$$\begin{aligned} \text{CV MSE}(\lambda, K) &= \frac{1}{K} \sum_{k=1}^K \text{MSE}(\lambda, k) \\ &= \frac{1}{K} \sum_{k=1}^K \left(\frac{1}{n_k} \sum_{i \in \mathcal{F}_k} \left(y_i - \mathbf{x}'_i \hat{\boldsymbol{\beta}}_{\text{lasso},k}(\lambda) \right)^2 \right) \end{aligned}$$

3.3 Post-Estimation of the Optimal Model

- Post-Estimation OLS :
 - Common practice
 - Fit OLS regression to the optimal model and obtain β_{ols}
 - Use β_{ols} for counterfactual prediction
 - **Post-Lasso OLS** : Fit OLS regression after Lasso selected
- Post-Estimation Lasso:
 - Obtain $\beta_{lasso}(\lambda)$ in **Lasso selection process**
 - Use $\beta_{lasso}(\lambda)$ for counterfactual prediction
 - **Can only be used after Lasso selection**

3.4 Placebo Test

3.4.1 Placebo Test Using Fake Treatment Unit

- Reassign the treatment to control units (donor pool) where no intervention actually occurred
- Determine statistical significance of treatment effect
- $p\text{-val}(t)$: p-value of estimated effect for a particular period is

$$p\text{-val}(t) = \frac{1}{N-1} \sum_{i=2}^N \mathbf{1} \left(\left| \hat{\Delta}_{it} \right| \geq \left| \hat{\Delta}_{1t} \right| \right) \quad t = T_0 + 1, \dots, T$$

3.4.1 Placebo Test Using Fake Treatment Unit

- *p-val*: The probability of obtaining a post/pre-MSPE ratio as large as that of treated unit, is

$$p\text{-val} = \frac{1}{N} \sum_{i=1}^N \mathbf{1} \left(\frac{\text{MSPE}_{i,post}}{\text{MSPE}_{i,pre}} \geq \frac{\text{MSPE}_{1,post}}{\text{MSPE}_{1,pre}} \right)$$

- Cutoff : Discard the fake units $i \in \{2, \dots, N\}$ with extreme values of $\text{MSPE}_{i,pre}$ (Abadie et al. , 2010)

3.4.2 Placebo Test Using Fake Treatment Time

- Reassign the treatment to periods previous to the intervention when no treatment actually occurred
- Whether there is a perceivable effect during $1, \dots, T_0 + 1$ periods?

4 The **rcm** command

- Implement "Regression Control Method (RCM)" in Stata
- Install: `ssc install rcm, all replace` with Stata version ≥ 16
- Installed files :
 - `rcm.ado` and `rcm.sthlp`: Stata ado and help file
- Ancillary files :
 - `growth.dta` : The dataset obtained from Hsiao et al. (2012) which has been reshaped to long form
 - `repgermany.dta` : The dataset obtained from Abadie et al. (2015) of which panel variable has been reencoded

4 The **rcm** command

- Syntax:

```
rcm depvar [indepvars] , trunit(#) trperiod(#) [options]
```

- `xtset panelvar timevar` must be used to declare a panel dataset in the usual long form
 - `depvar` and `indepvars` must be numeric variables, and abbreviations are not allowed.
- Without `indepvars` : Basic model (Section 2.1)
 - With `indepvars` : Model with covariates (Section 2.2)

5 Examples

5.1 Replicate Hsiao et al.(2012)

- Consider the impact on Hong Kong real GDP growth rate with the reversion of sovereignty on 1 July 1997 from the UK to China
 - Treatment period : 1997Q3
 - Pre-treatment periods : 1993Q1-1997Q2
 - Post-treatment periods : 1997Q3-2004Q1
 - Treated unit : Hong Kong
 - Control units : 10 countries/regions that are either in the region or economically closely associated with Hong Kong

5.1 Replicate Hsiao et al.(2012)

```
use growth, clear  
xtset region time  
des
```

```
Observations:      1,525  
Variables:         3      16 Jun 2019 00:03
```

Variable name	Storage type	Display format	Value label	Variable label
time	float	%tq		
gdp	float	%8.0g		
region	long	%13.0g	region	

```
Sorted by: region time
```

5.1 Replicate Hsiao et al.(2012)

label list

region:

- | | |
|--------------|------------------|
| 1 Australia | 14 Malaysia |
| 2 Austria | 15 Mexico |
| 3 Canada | 16 Netherlands |
| 4 China | 17 NewZealand |
| 5 Denmark | 18 Norway |
| 6 Finland | 19 Philippines |
| 7 France | 20 Singapore |
| 8 Germany | 21 Switzerland |
| 9 HongKong | 22 Taiwan |
| 10 Indonesia | 23 Thailand |
| 11 Italy | 24 UnitedKingdom |
| 12 Japan | 25 UnitedStates |
| 13 Korea | |

5.1 Replicate Hsiao et al.(2012)

```
rcm gdp, trunit(9) trperiod(150) ctrlunit(4 10 12 13 14 19 20 22 23 25) postperiod(150/175)
```

- `gdp` : Specifies "gdp" as dependent variable (outcome variable)
- `trunit(9)` : Specifies "HongKong" as the treated unit
- `trperiod(150)` : Specifies "1997q3" as the treatment period (`150` is obtained from `di tq(1997q3)`)
- `ctrlunit(4 10 12 13 14 19 20 22 23 25)` : Specifies 10 countries/regions as the control units
- `postperiod(150/175)` : Specifies "1997q3-2003q4" as the post-treatment periods (`175` is obtained from `di tq(2003q4)`)

5.1 Replicate Hsiao et al.(2012)

Step 1: Select the suboptimal models
(method **best** specified)

Note: If this takes too long, you may wish to try `method(lasso)`(recommended), `method(forward)` or `method(backward)`. Alternatively, you may restrict `indepvars`, and/or the donor pool by the option `ctrlunit()`.

Selecting the suboptimal model with number of predictors **1-10...**

Step 2: Select the optimal model from the suboptimal models
(criterion **aicc** specified)

Comparing the suboptimal models containing different set of predictors:

K	AICc	AIC	BIC	R-squared
1	-144.7514	-146.4657	-143.7946	0.4034
2	-160.5063	-163.5832	-160.0217	0.7937
3	-170.6492	-175.6492	-171.1973	0.9056
4	-171.7725	-179.4088	-174.0666	0.9314
5	-169.7878	-180.9878	-174.7552	0.9438
6	-164.2937	-180.2937	-173.1707	0.9477
7	-156.6834	-179.1834	-171.1701	0.9503
8	-146.2921	-177.7207	-168.8169	0.9517
9	-131.7464	-175.7464	-165.9523	0.9518
10	-111.3603	-173.7603	-163.0758	0.9518

Among models with **1-10** predictors, the optimal model contains **4** predictors with **AICc = -171.7725**.

5.1 Replicate Hsiao et al.(2012)

Fitting results in the pre-treatment periods using OLS:

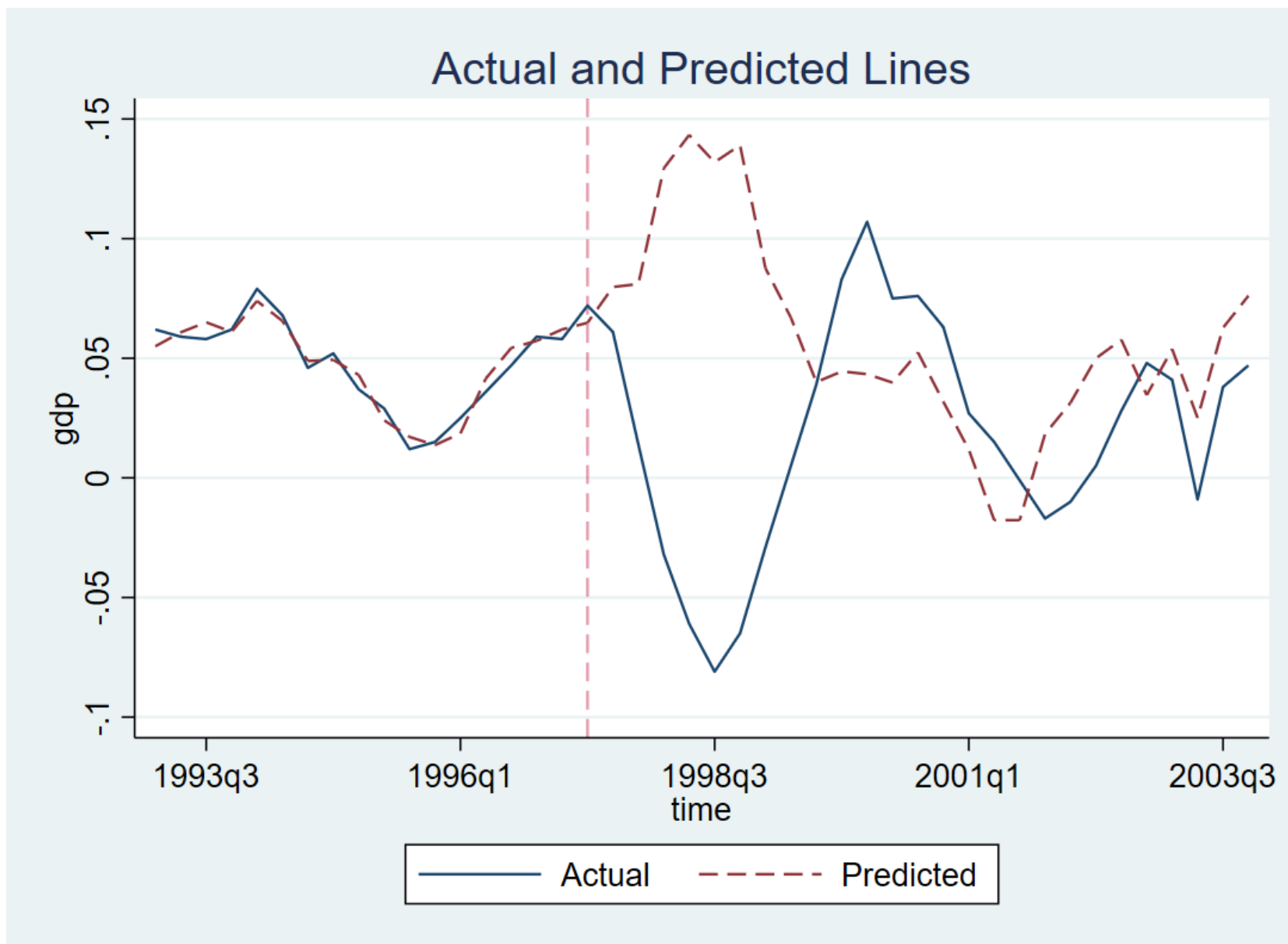
Mean Absolute Error	=	0.00611	Number of Observations	=	18
Mean Squared Error	=	0.00003	Number of Predictors	=	4
Root Mean Squared Error	=	0.00578	R-squared	=	0.93144

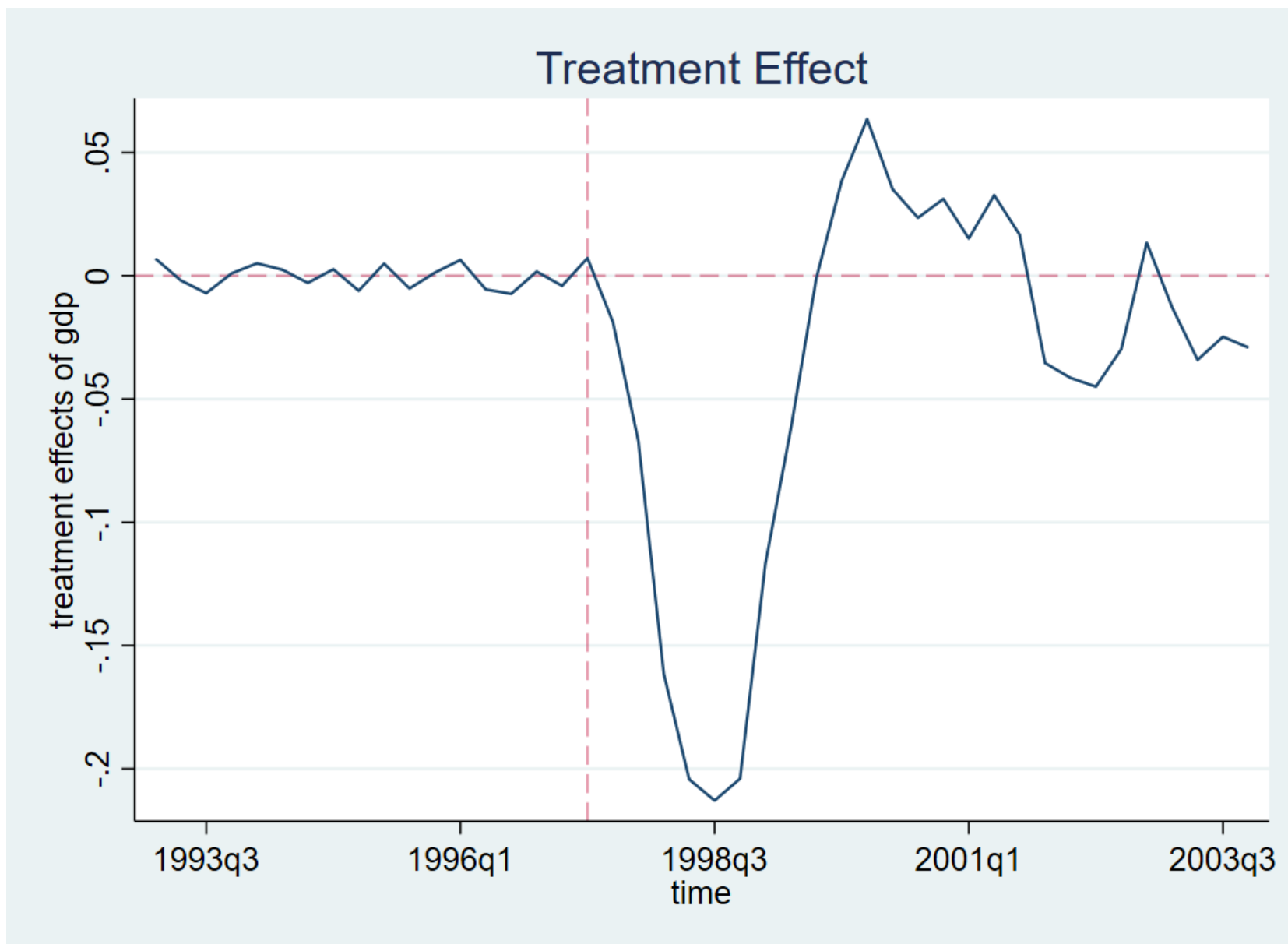
gdp·HongKong	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
gdp·Korea	-0.4323	0.0634	-6.82	0.000	-0.5692	-0.2954
gdp·Japan	-0.6760	0.1117	-6.05	0.000	-0.9172	-0.4347
gdp·Taiwan	0.7926	0.3099	2.56	0.024	0.1231	1.4621
gdp·UnitedStates	0.4860	0.2195	2.21	0.045	0.0118	0.9603
_cons	0.0263	0.0170	1.54	0.147	-0.0105	0.0631

5.1 Replicate Hsiao et al.(2012)

Time	Treated	Predicted	Tr. Effect
1997q3	0.0610	0.0798	-0.0188
1997q4	0.0140	0.0810	-0.0670
1998q1	-0.0320	0.1294	-0.1614
1998q2	-0.0610	0.1433	-0.2043
1998q3	-0.0810	0.1319	-0.2129
1998q4	-0.0650	0.1390	-0.2040
1999q1	-0.0290	0.0876	-0.1166
1999q2	0.0050	0.0670	-0.0620
1999q3	0.0390	0.0400	-0.0010
1999q4	0.0830	0.0445	0.0385
2000q1	0.1070	0.0434	0.0636
2000q2	0.0750	0.0398	0.0352
2000q3	0.0760	0.0524	0.0236
2000q4	0.0630	0.0318	0.0312
2001q1	0.0270	0.0118	0.0152
2001q2	0.0150	-0.0177	0.0327
2001q3	-0.0010	-0.0177	0.0167
2001q4	-0.0170	0.0184	-0.0354
2002q1	-0.0100	0.0314	-0.0414
2002q2	0.0050	0.0500	-0.0450
2002q3	0.0280	0.0577	-0.0297
2002q4	0.0480	0.0346	0.0134
2003q1	0.0410	0.0538	-0.0128
2003q2	-0.0090	0.0251	-0.0341
2003q3	0.0380	0.0628	-0.0248
2003q4	0.0470	0.0761	-0.0291
Mean	0.0180	0.0576	-0.0396

Note: The average treatment effect over the post-treatment periods is **-0.0396**.





5.1 Replicate Hsiao et al.(2012)

- Consider the impact on Hong Kong real GDP growth rate with the implementation of CEPA starting in 2004Q1 between mainland China and Hong Kong.
 - Treatment period : 2004Q1
 - Pre-treatment periods : 1993Q1-2003Q4
 - Post-treatment periods : 2004Q1-2008Q1
 - Treated unit : Hong Kong
 - Control units : All countries/regions in dataset except Hong Kong

5.1 Replicate Hsiao et al.(2012)

```
rcm gdp, trunit(9) trperiod(176)
```

- `trperiod(176)` : Specifies "2004q1" as the treatment period
(`176` is obtained from `di tq(2004q1)`)

5.1 Replicate Hsiao et al.(2012)

Step 1: Select the suboptimal models
 (method **best** specified)

Note: If this takes too long, you may wish to try **method(lasso)**(recommended), **method(forward)** or **method(backward)**. Alternatively, you may restrict *indepvars*, and/or the donor pool by the option **ctrlunit()**.

Selecting the suboptimal model with number of predictors **1-24...**

Step 2: Select the optimal model from the suboptimal models
 (criterion **aicc** specified)

Comparing the suboptimal models containing different set of predictors:

K	AICc	AIC	BIC	R-squared
1	-313.8269	-314.4269	-309.0743	0.5877
2	-335.2386	-336.2642	-329.1275	0.7602
3	-348.2800	-349.8590	-340.9380	0.8318
4	-365.6420	-367.9122	-357.2071	0.8933
5	-377.4412	-380.5523	-368.0630	0.9235
6	-378.9426	-383.0569	-368.7833	0.9310
7	-378.9074	-384.2016	-368.1439	0.9357
8	-378.5854	-385.2521	-367.4102	0.9400
9	-377.5003	-385.7503	-366.1242	0.9433
10	-375.0098	-385.0744	-363.6641	0.9450
11	-372.4606	-384.5939	-361.3994	0.9469
12	-369.2578	-383.7405	-358.7619	0.9483
13	-365.9158	-383.0586	-356.2958	0.9498
14	-362.5660	-382.7142	-354.1671	0.9516
15	-358.3157	-381.8542	-351.5230	0.9529
16	-353.3736	-380.7336	-348.6182	0.9538
17	-348.1579	-379.8246	-345.9250	0.9549
18	-342.4931	-379.0149	-343.3311	0.9561
19	-335.8492	-377.8492	-340.3812	0.9570
20	-328.0881	-376.2785	-337.0264	0.9574
21	-319.2286	-374.4286	-333.3922	0.9575
22	-309.3373	-372.4952	-329.6747	0.9576
23	-298.3113	-370.5335	-325.9288	0.9576
24	-285.9617	-368.5499	-322.1610	0.9576

Among models with **1-24** predictors, the optimal model contains **6** predictors with **AICc = -378.9426**.

5.1 Replicate Hsiao et al.(2012)

Fitting results in the pre-treatment periods using OLS:

Mean Absolute Error	=	0.01070	Number of Observations	=	44
Mean Squared Error	=	0.00014	Number of Predictors	=	6
Root Mean Squared Error	=	0.01170	R-squared	=	0.93097

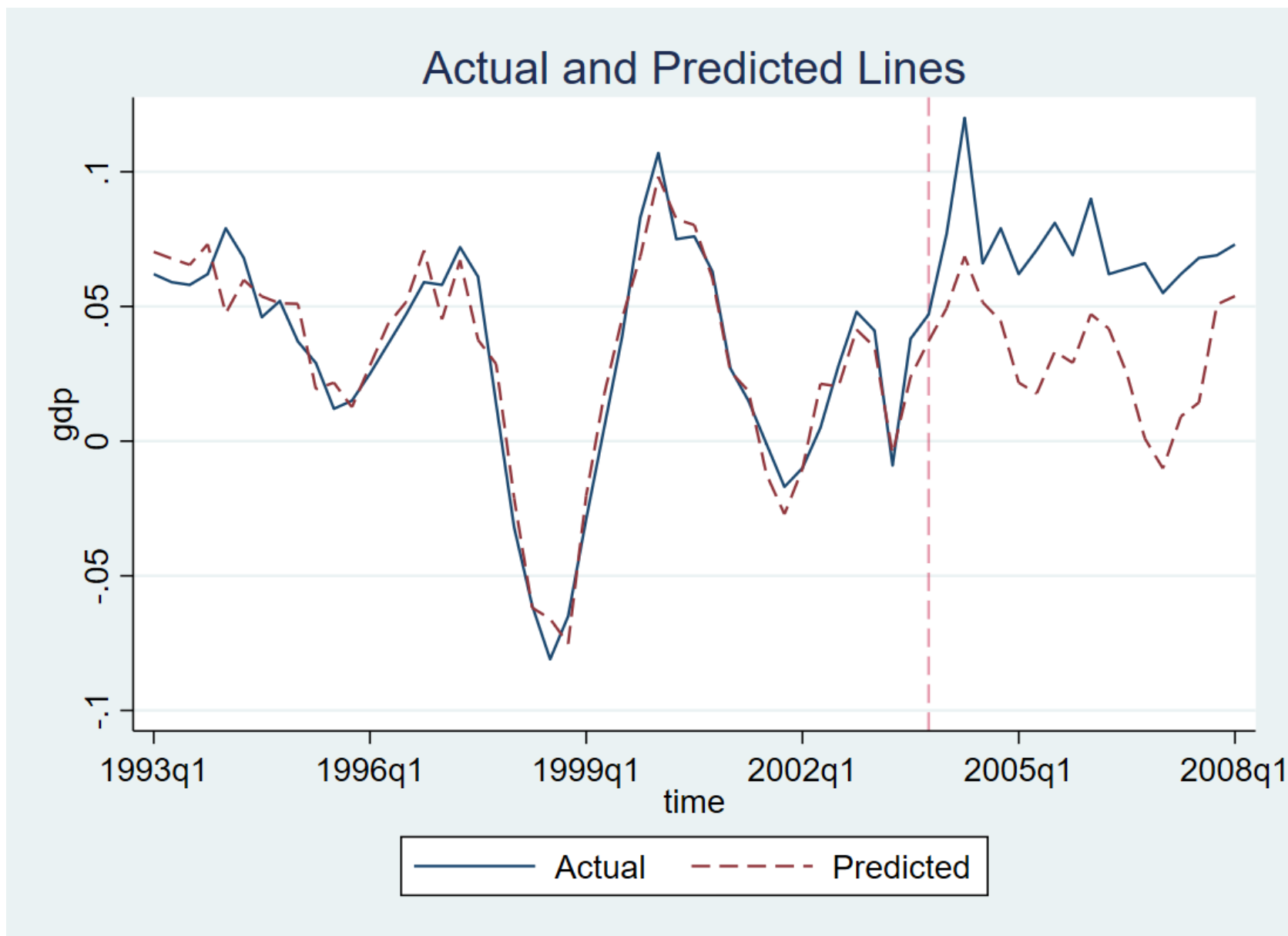
gdp·HongKong	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
gdp·Norway	0.3222	0.0538	5.99	0.000	0.2132	0.4311
gdp·Austria	-1.0115	0.1682	-6.01	0.000	-1.3524	-0.6707
gdp·Korea	0.3447	0.0469	7.35	0.000	0.2497	0.4398
gdp·Mexico	0.3129	0.0510	6.13	0.000	0.2095	0.4162
gdp·Italy	-0.3177	0.1591	-2.00	0.053	-0.6400	0.0046
gdp·Singapore	0.1845	0.0546	3.38	0.002	0.0739	0.2951
_cons	-0.0019	0.0037	-0.52	0.603	-0.0094	0.0056

5.1 Replicate Hsiao et al.(2012)

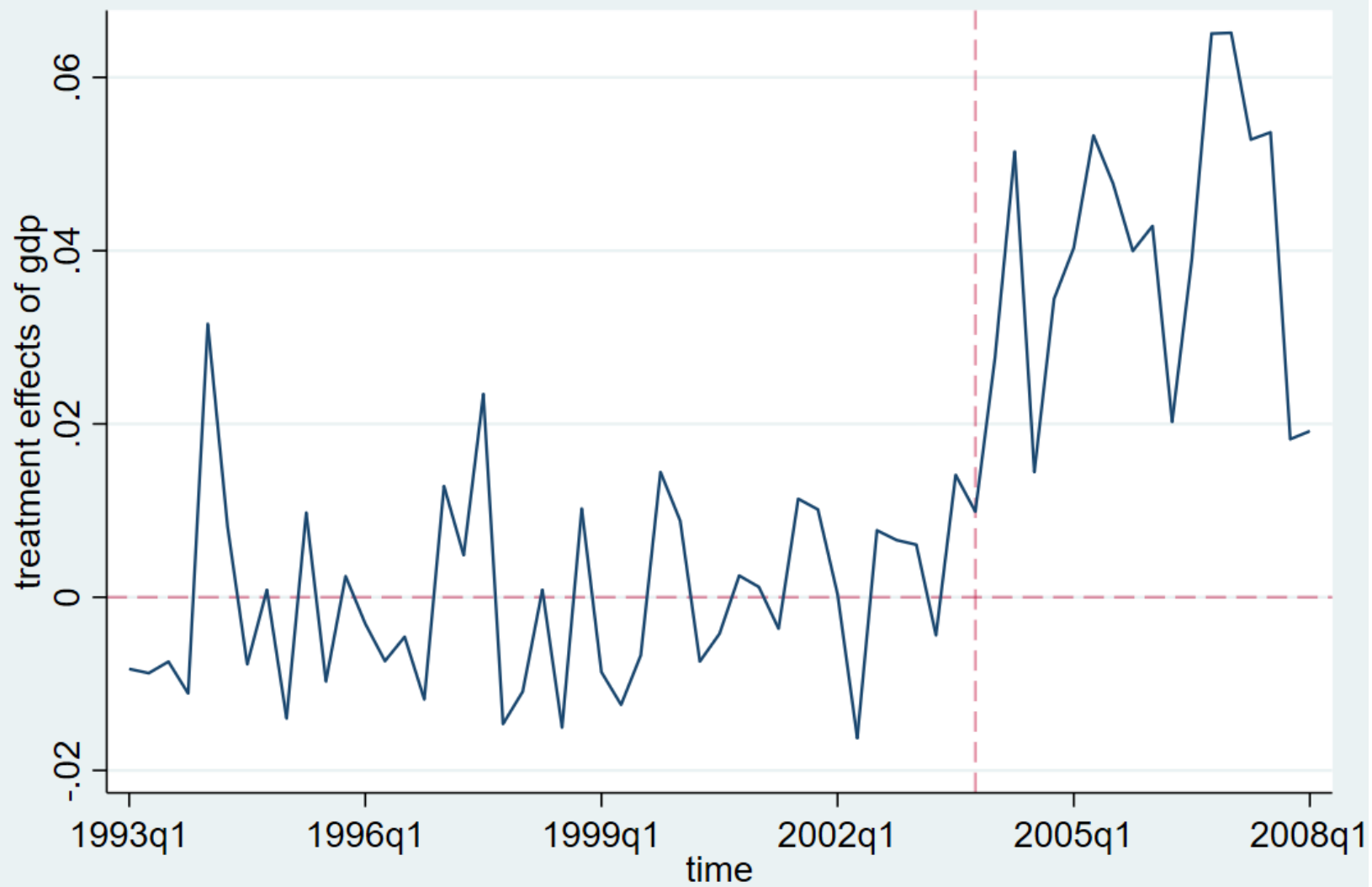
Prediction results in the post-treatment periods using OLS:

Time	Treated	Predicted	Tr. Effect
2004q1	0.0770	0.0493	0.0277
2004q2	0.1200	0.0686	0.0514
2004q3	0.0660	0.0515	0.0145
2004q4	0.0790	0.0446	0.0344
2005q1	0.0620	0.0217	0.0403
2005q2	0.0710	0.0177	0.0533
2005q3	0.0810	0.0333	0.0477
2005q4	0.0690	0.0290	0.0400
2006q1	0.0900	0.0471	0.0429
2006q2	0.0620	0.0417	0.0203
2006q3	0.0640	0.0250	0.0390
2006q4	0.0660	0.0009	0.0651
2007q1	0.0550	-0.0101	0.0651
2007q2	0.0620	0.0092	0.0528
2007q3	0.0680	0.0143	0.0537
2007q4	0.0690	0.0508	0.0182
2008q1	0.0730	0.0538	0.0192
Mean	0.0726	0.0323	0.0403

Note: The average treatment effect over the post-treatment periods is **0.0403**.



Treatment Effect



5.2 Illustrate RCM with Covariates and Placebo Test

- Use the same dataset as Abadie et al. (2015)
- Estimate the economic impact of the 1990 German reunification
 - Treatment period : 1990
 - Pre-treatment periods : 1960-1989
 - Post-treatment periods : 1990-2003
 - Treated unit : West Germany
 - Control units : 16 OECD member countries

5.2 Illustrate RCM with Covariates and Placebo Test

```
use reppgermany.dta, clear
xtset country year
des
```

Contains data from **reppgermany.dta**

Observations: **748**
 Variables: **10** 12 Aug 2021 08:25

Variable name	Storage type	Display format	Value label	Variable label
year	float	%8.0g		Year
gdp	long	%8.0g		GDP per-capita (annual)
infrate	float	%9.0g		Inflation Rate (annual)
trade	float	%9.0g		Trade openness (annual)
schooling	float	%9.0g		Schooling (every 5 years)
invest60	float	%9.0g		Investment rate (average 60-65)
invest70	float	%9.0g		Investment rate (average 70-75)
invest80	float	%9.0g		Investment rate (average 80-85)
industry	float	%9.0g		Industry Share (annual)
country	long	%12.0g	country	Country Name

Sorted by: **country year**

5.2 Illustrate RCM with Covariates and Placebo Test

```
label list
```

```
country:
```

- 1 Australia
- 2 Austria
- 3 Belgium
- 4 Denmark
- 5 France
- 6 Greece
- 7 Italy
- 8 Japan
- 9 Netherlands
- 10 New Zealand
- 11 Norway
- 12 Portugal
- 13 Spain
- 14 Switzerland
- 15 UK
- 16 USA
- 17 West Germany

5.2 Illustrate RCM with Covariates and Placebo Test

```
rcm gdp infrate trade industry, tru(17) trp(1990) me(lasso) cr(cv) fold(10)
```

- `infrate trade industry` : Specifies three covariates
- `tru(17)` : Abbr. for `trunit(17)`
- `trp(1990)` : Abbr. for `trperiod(1990)`
- `me(lasso)` : Abbr. for `method(lasso)`, which specifies Lasso as selection method
- `cr(cv)` : Abbr. for `criterion(cv)`, which specifies cross-validation as the selection criterion
- `fold(10)` : Specifies that cross-validation with 10 folds

5.2 Illustrate RCM with Covariates and Placebo Test

Step 1: Select the suboptimal models
 (method **lasso** specified)

Selecting the suboptimal model...

Step 2: Select the optimal model from the suboptimal models
 (criterion **cv** specified for **10-fold** cross-validation)

Comparing the suboptimal models containing different set of predictors:

K	AICc	AIC	BIC	CV MSE	R-squared	lambda	Operation
1	526.5586	525.6355	529.8391	2.187e+07	0.0888	4287.1880	add gdp·Italy
2	506.8945	505.2945	510.8993	1.038e+07	0.5668	2954.9924	add gdp·Netherlands
3	504.0514	501.5514	508.5574	8.616e+06	0.6403	2692.4790	add gdp·Austria
4	484.0258	480.3737	488.7808	4.061e+06	0.8289	1855.8213	add gdp·Denmark
5	478.7213	473.6304	483.4388	3.058e+06	0.8706	1614.0987	add gdp·USA
7	407.2748	398.2748	410.8855	2.420e+05	0.9894	459.7011	add gdp·Greece gdp·Norway
6	357.2235	350.3663	361.5759	58902.4925	0.9976	218.3953	drop gdp·Netherlands
7	334.6504	325.6504	338.2611	29433.5466	0.9988	150.5314	add gdp·Switzerland
8	328.6835	317.1046	331.1166	25005.4310	0.9990	137.1586	add industry·Spain
9	285.8035	271.1368	286.5500	6852.9127	0.9998	59.3727	add gdp·Netherlands
9	270.7571	256.0904	271.5036	4898.0417	0.9999	44.9133	.

Among models with 1-67 predictors, the optimal model contains 9 predictors with CV MSE = 4898.0417.

5.2 Illustrate RCM with Covariates and Placebo Test

Fitting results in the pre-treatment periods using **post-lasso OLS**:

Mean Absolute Error	=	12.87119	Number of Observations	=	30
Mean Squared Error	=	3.0e+02	Number of Predictors	=	9
Root Mean Squared Error	=	17.30368	R-squared	=	0.99999

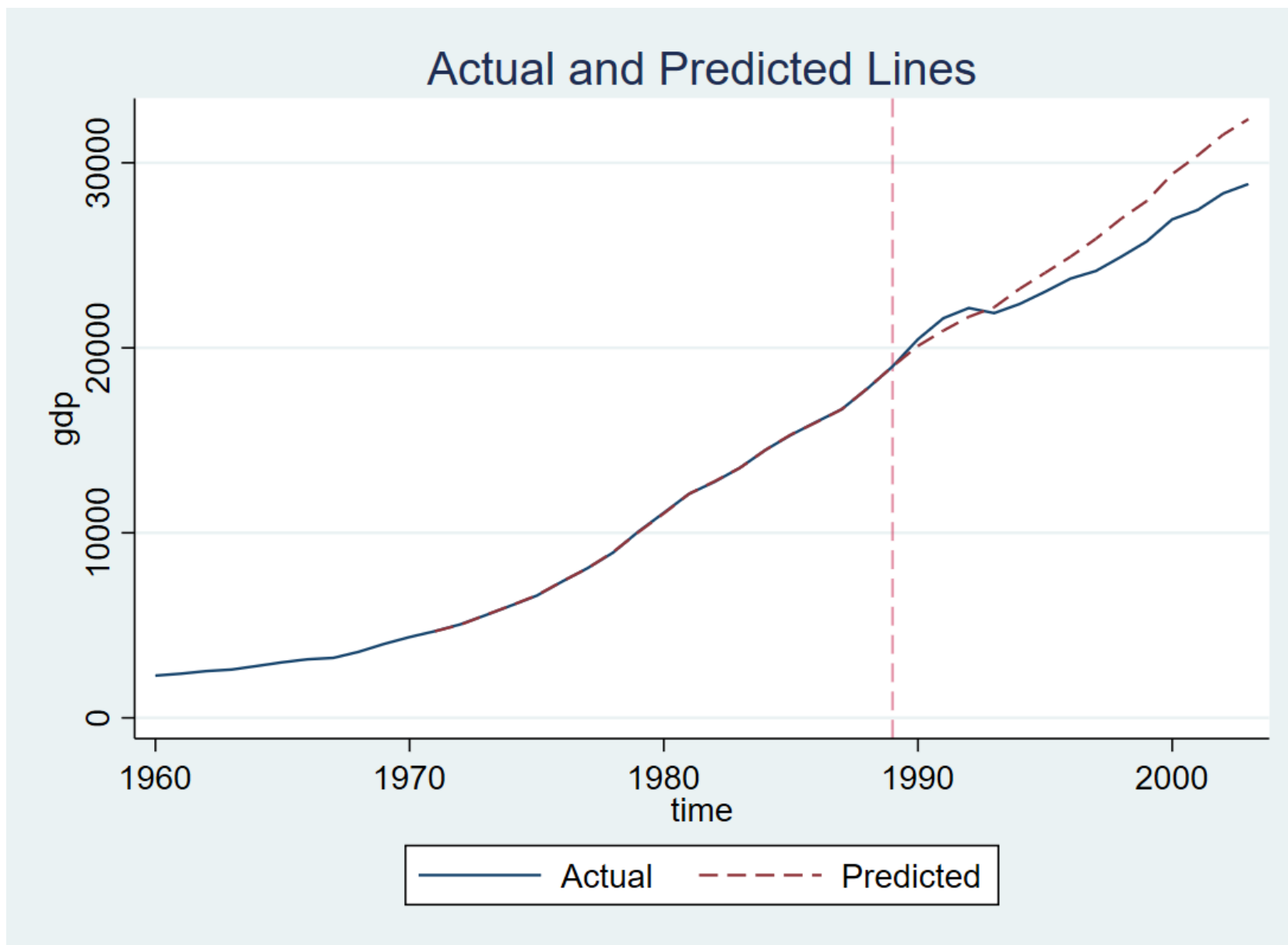
gdp·WestGermany	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
gdp·Austria	0.1331	0.0708	1.88	0.093	-0.0271	0.2934
gdp·Denmark	0.1046	0.0788	1.33	0.217	-0.0735	0.2828
gdp·Greece	0.1513	0.0422	3.58	0.006	0.0557	0.2468
gdp·Italy	0.2914	0.0838	3.48	0.007	0.1018	0.4809
gdp·Netherlands	0.1334	0.1097	1.22	0.255	-0.1148	0.3816
gdp·Norway	-0.0313	0.0598	-0.52	0.613	-0.1665	0.1039
industry·Spain	-44.6926	11.6744	-3.83	0.004	-71.1019	-18.2833
gdp·Switzerland	0.0548	0.0343	1.60	0.144	-0.0228	0.1324
gdp·USA	0.2282	0.0439	5.20	0.001	0.1289	0.3276
_cons	1755.2090	419.7465	4.18	0.002	805.6764	2704.7417

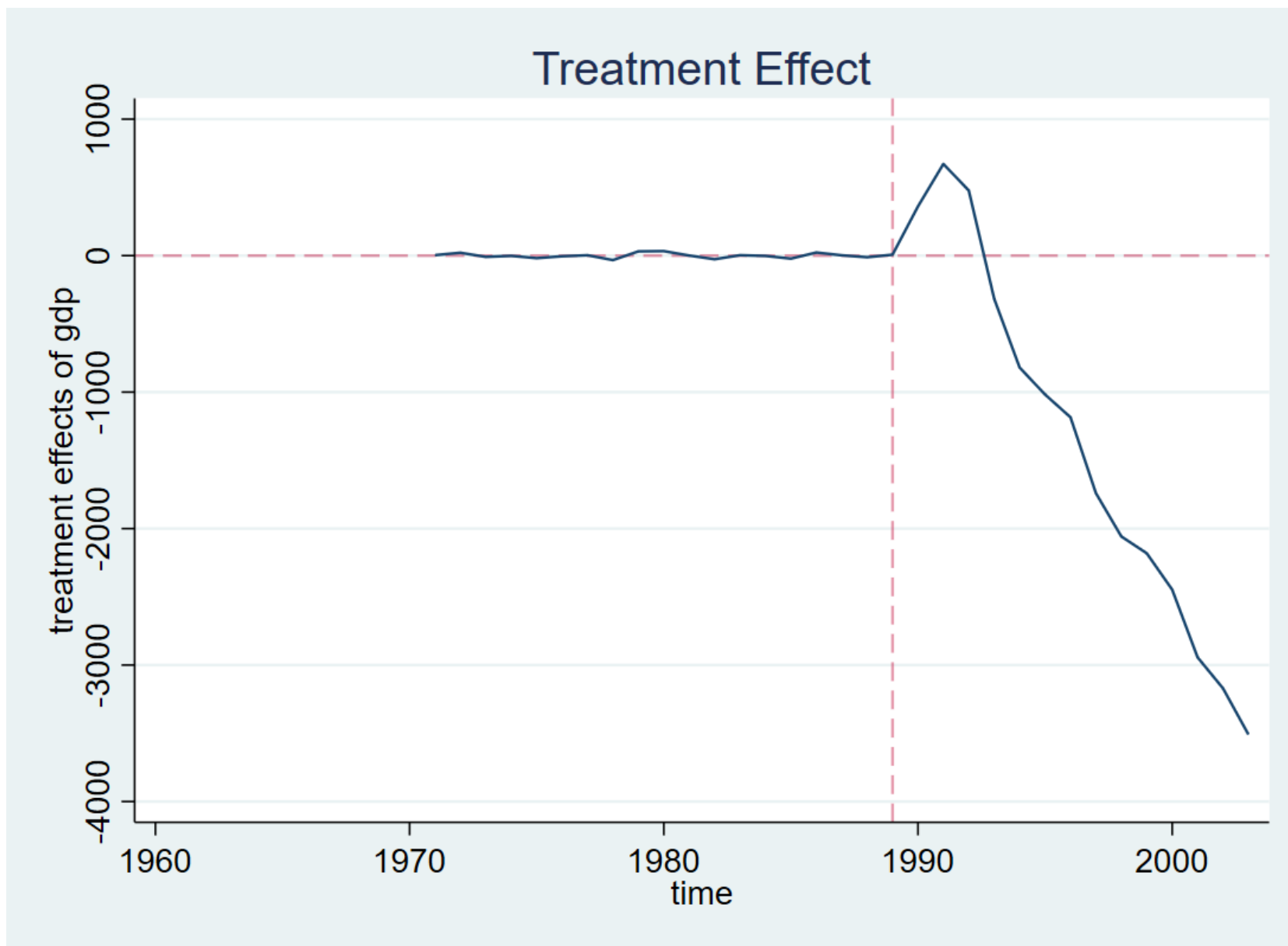
5.2 Illustrate RCM with Covariates and Placebo Test

Prediction results in the post-treatment periods using **post-lasso OLS**:

Time	Treated	Predicted	Tr. Effect
1990	20465.0000	20104.6133	360.3867
1991	21602.0000	20930.3809	671.6191
1992	22154.0000	21677.2500	476.7500
1993	21878.0000	22194.6797	-316.6797
1994	22371.0000	23190.9297	-819.9297
1995	23035.0000	24052.6563	-1017.6563
1996	23742.0000	24926.1309	-1184.1309
1997	24156.0000	25896.0313	-1740.0313
1998	24931.0000	26988.5430	-2057.5430
1999	25755.0000	27935.7734	-2180.7734
2000	26943.0000	29389.8184	-2446.8184
2001	27449.0000	30392.2207	-2943.2207
2002	28348.0000	31518.2871	-3170.2871
2003	28855.0000	32363.5508	-3508.5508
Mean	24406.0000	25825.7761	-1419.7761

Note: The average treatment effect over the post-treatment periods is **-1419.7761**.





5.2 Illustrate RCM with Covariates and Placebo Test

```
rcm gdp infrate trade industry, tru(17) trp(1990) me(lasso) cr(cv) fold(10) fill(mean) placebo(unit cut(10))
```

- `fill(mean)` : Fill in missing values by sample means for each units
- `placebo(unit cut(10))` :
 - Implement placebo tests using the fake treatment units
 - Discard the fake treatment units of which pre-treatment MSPE 10 times smaller than or equal to that of the treated unit

5.2 Illustrate RCM with Covariates and Placebo Test

Implementing placebo effects using fake treatment unit **Australia...Austria...Belgium...Denmark...France...Greece...Italy...Japan...Netherlands...NewZealand...Norway...Portugal...Spain...Switzerland...UK...USA...**

Placebo test results using fake treatment units:

Unit	Pre MSPE	Post MSPE	Post/Pre MSPE	Pre MSPE of Unit/Treated Unit
WestGermany	2009.4254	4912795.2300	2444.8757	1.0000
Austria	6616.8915	191860.4677	28.9956	3.2929
Belgium	2640.2164	81819.2402	30.9896	1.3139
France	3509.5265	241028.1689	68.6783	1.7465
Greece	16088.6537	1218658.0663	75.7464	8.0066
Italy	2180.6256	181301.3038	83.1419	1.0852
Japan	10414.5510	2085867.0714	200.2839	5.1829
Netherlands	5532.6214	1278504.6038	231.0848	2.7533
Portugal	7468.8217	858321.5550	114.9206	3.7169
Spain	8020.4273	1327015.9387	165.4545	3.9914
UK	6950.5984	443053.1737	63.7432	3.4590
USA	14405.5364	5967054.2037	414.2195	7.1690

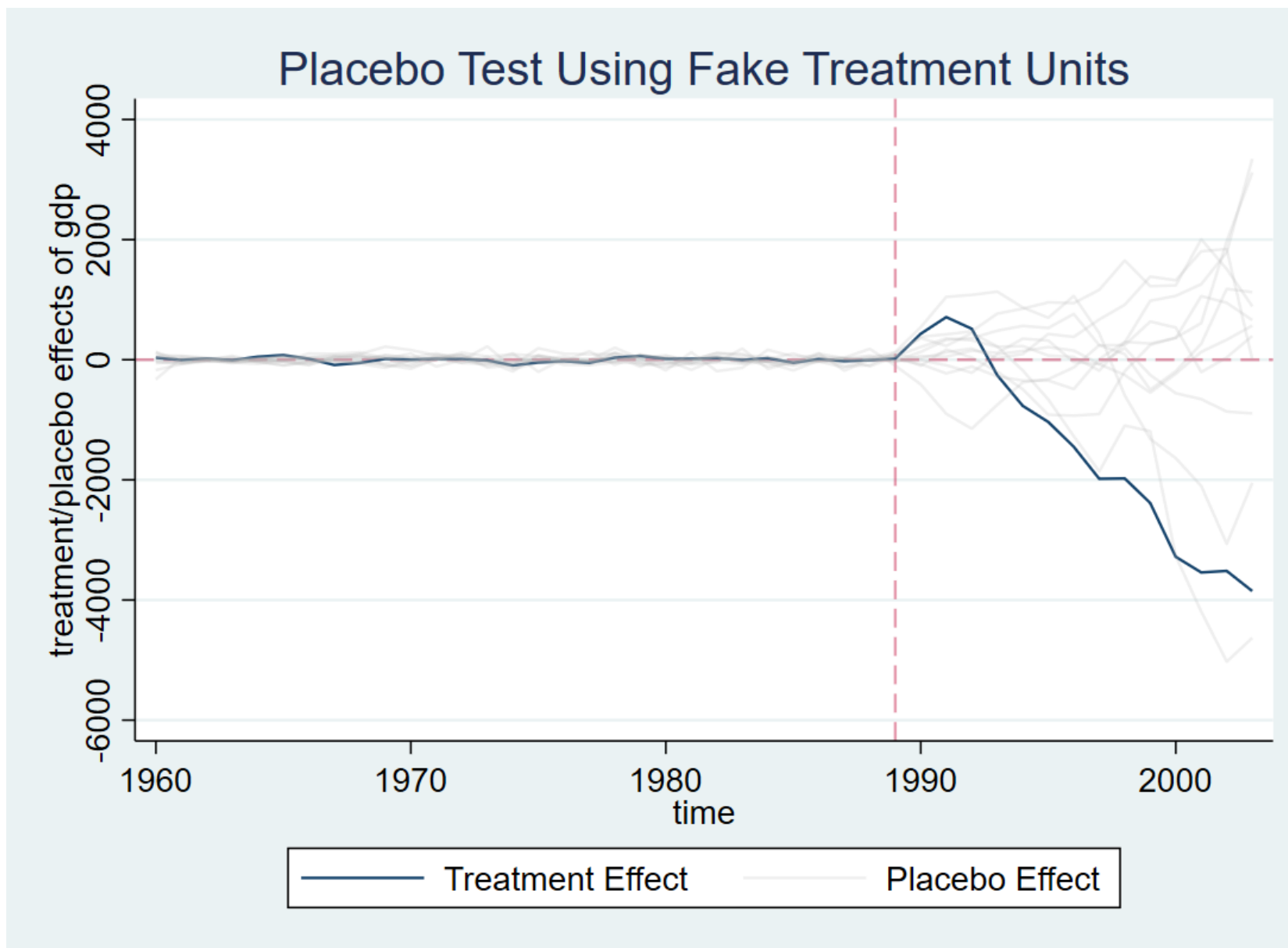
Note: The units **Australia Denmark NewZealand Norway Switzerland** with Pre-Treatment MSPE **10** times higher than **WestGermany's** are excluded. The probability of obtaining a post/pre-treatment MSPE ratio as large as **WestGermany's** is **0.0833**.

5.2 Illustrate RCM with Covariates and Placebo Test

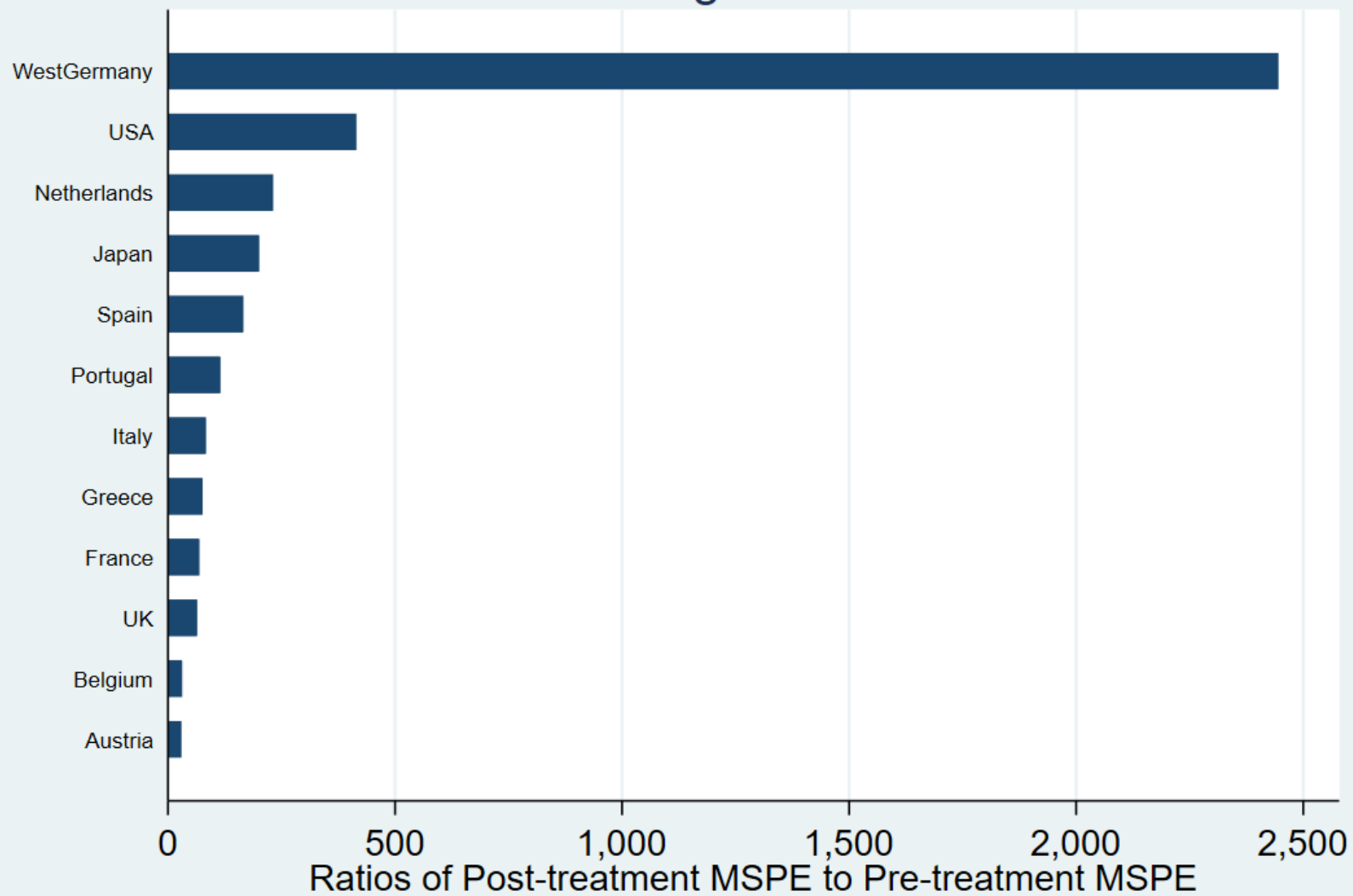
Placebo test results using fake treatment units (continued):

Time	Tr. Eff.	P-value
1990	432.1543	0.0909
1991	709.9277	0.1818
1992	515.8379	0.1818
1993	-256.7402	0.4545
1994	-766.9336	0.1818
1995	-1035.2422	0.0000
1996	-1448.2324	0.0000
1997	-1981.4160	0.0000
1998	-1975.5234	0.0000
1999	-2385.0527	0.0000
2000	-3280.4258	0.0000
2001	-3540.6426	0.0909
2002	-3515.3242	0.0909
2003	-3850.5918	0.0909

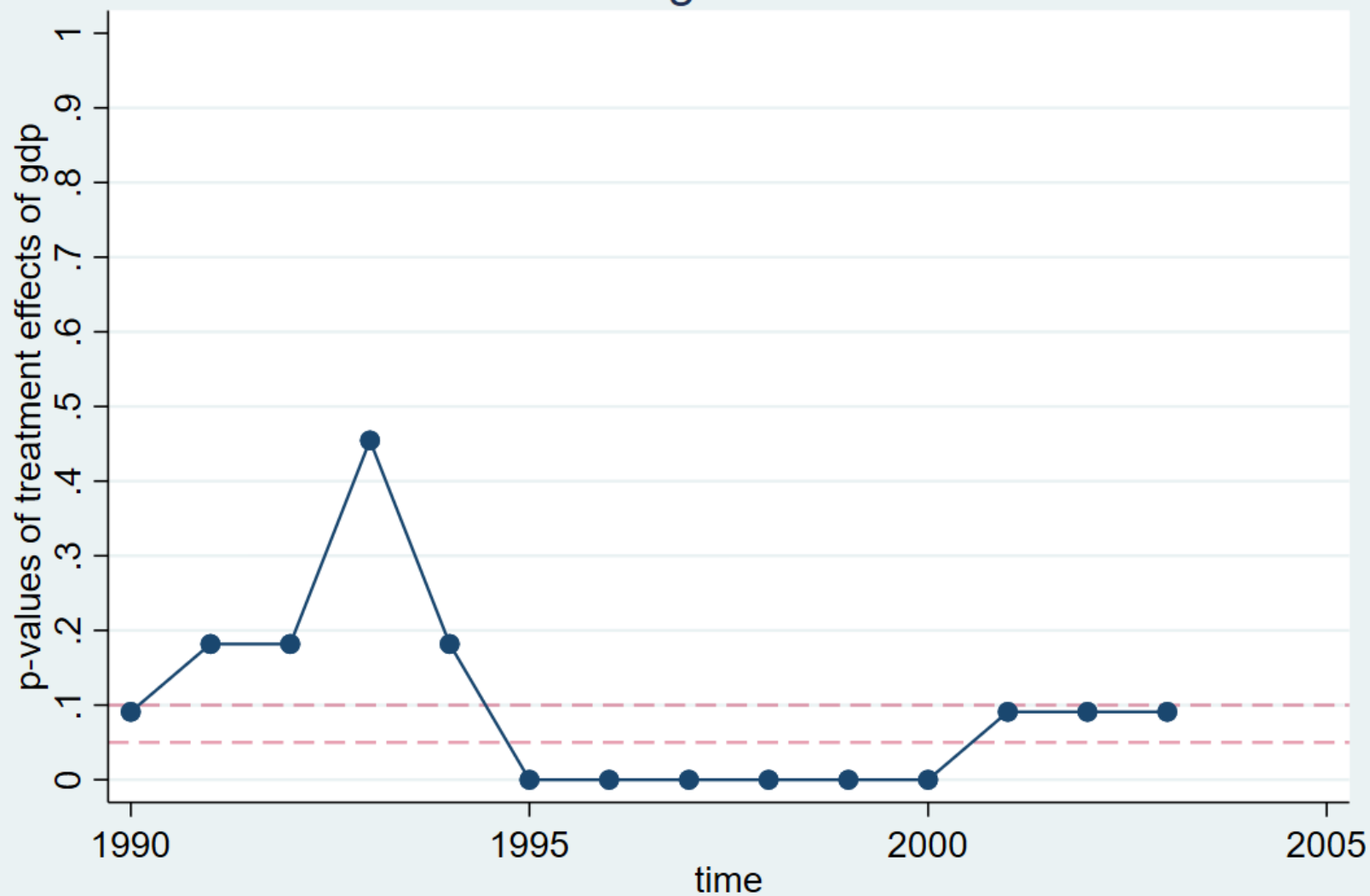
Note: The p-value of the treatment effect for a particular period is defined as the frequency that the absolute values of the placebo effects are greater or equal to the absolute value of treatment effect.



Placebo Test Using Fake Treatment Units



Placebo Test Using Fake Treatment Units



5.2 Illustrate RCM with Covariates and Placebo Test

```
rcm gdp infrate trade industry, tru(17) trp(1990) me(lasso) cr(cv) fold(10) fill(linear) placebo(period(1980 1985))
```

- `fill(linear)` : Fill in missing values by linear interpolation for each units
- `placebo(period(1980 1985))` : Implement placebo tests with fake treatment time 1980 and 1985

Implementing placebo effects using fake treatment time 1980...1985...

Placebo test results using fake treatment time 1980:

Time	Treated	Predicted	Tr. Eff.
1980	11083.0000	11062.1953	20.8047
1981	12115.0000	11937.9014	177.0986
1982	12761.0000	12584.5449	176.4551
1983	13519.0000	13487.0322	31.9678
1984	14481.0000	14485.5996	-4.5996
1985	15291.0000	15414.0498	-123.0498
1986	15998.0000	15827.1543	170.8457
1987	16679.0000	16503.3066	175.6934
1988	17786.0000	17469.7988	316.2012
1989	18994.0000	18590.4004	403.5996
1990	20465.0000	19414.7520	1050.2480
1991	21602.0000	20036.1699	1565.8301
1992	22154.0000	20710.7656	1443.2344
1993	21878.0000	21346.8164	531.1836
1994	22371.0000	22564.2402	-193.2402
1995	23035.0000	23573.0605	-538.0605
1996	23742.0000	25172.0996	-1430.0996
1997	24156.0000	26662.7520	-2506.7520
1998	24931.0000	27212.3789	-2281.3789
1999	25755.0000	28998.8145	-3243.8145
2000	26943.0000	32415.6914	-5472.6914
2001	27449.0000	33603.8672	-6154.8672
2002	28348.0000	34962.1016	-6614.1016
2003	28855.0000	36147.2969	-7292.2969
Mean	20432.9583	21674.2829	-1241.3246

Note: The average treatment effect over the post-treatment periods is **-1241.3246**.

5.2 Illustrate RCM with Covariates and Placebo Test

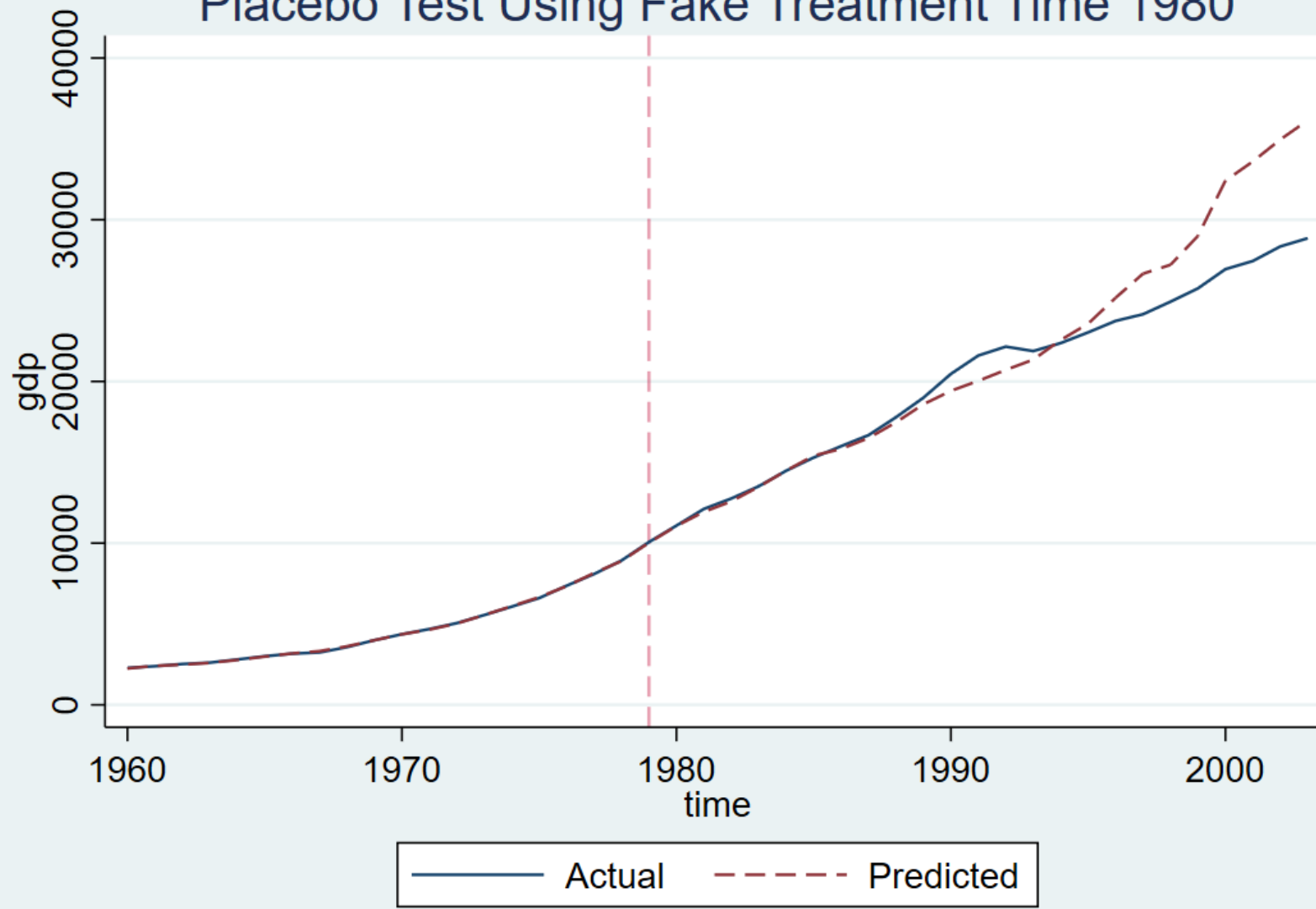
5.2 Illustrate RCM with Covariates and Placebo Test

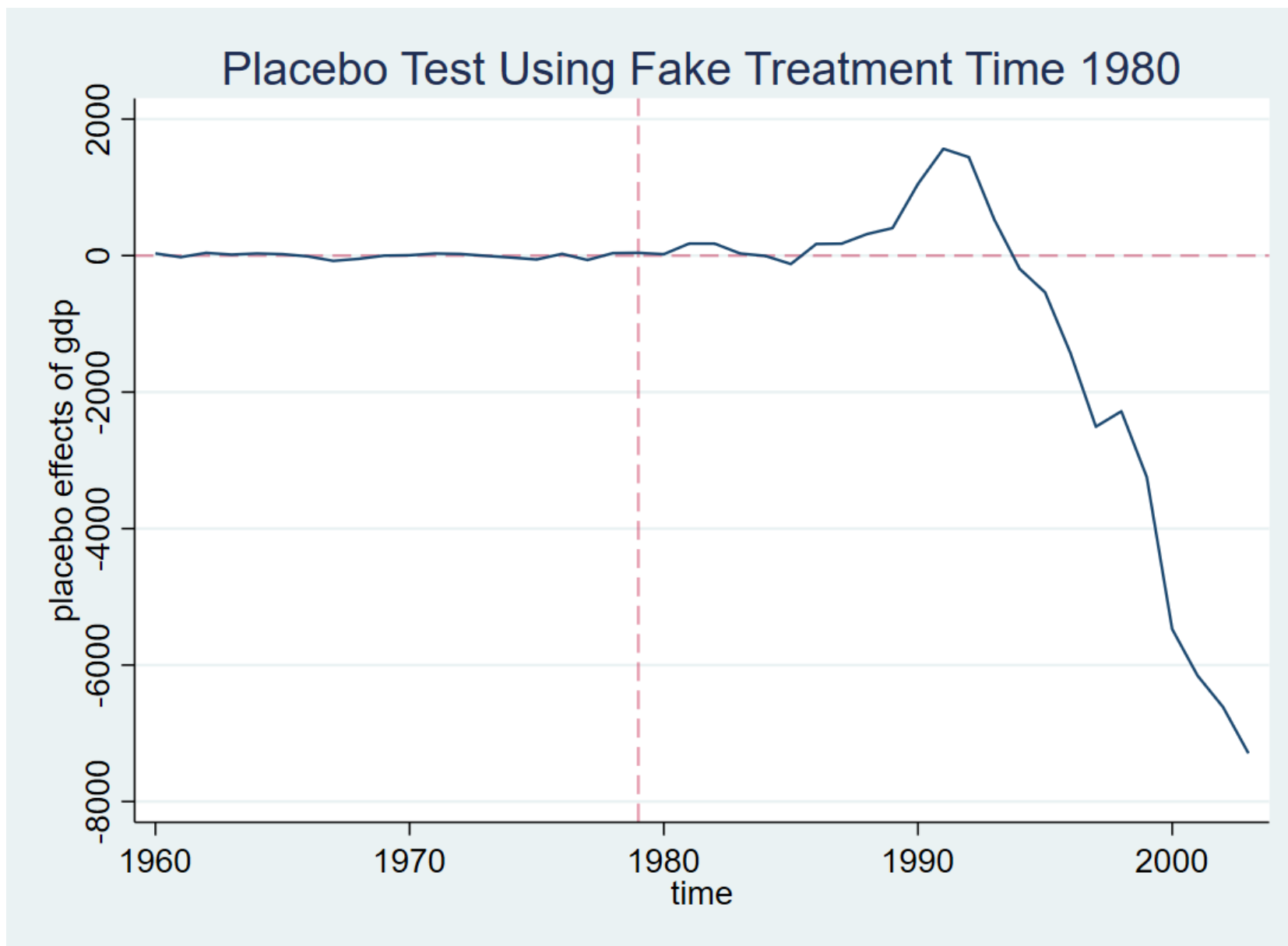
Placebo test results using fake treatment time **1985**:

Time	Treated	Predicted	Tr. Eff.
1985	15291.0000	15399.4453	-108.4453
1986	15998.0000	15950.9844	47.0156
1987	16679.0000	16604.2344	74.7656
1988	17786.0000	17587.4297	198.5703
1989	18994.0000	18738.6387	255.3613
1990	20465.0000	19745.6875	719.3125
1991	21602.0000	20602.5215	999.4785
1992	22154.0000	21363.2012	790.7988
1993	21878.0000	21871.5781	6.4219
1994	22371.0000	22915.0332	-544.0332
1995	23035.0000	23870.9434	-835.9434
1996	23742.0000	25165.6582	-1423.6582
1997	24156.0000	26213.7539	-2057.7539
1998	24931.0000	26917.2168	-1986.2168
1999	25755.0000	28444.0781	-2689.0781
2000	26943.0000	31154.6328	-4211.6328
2001	27449.0000	31974.5781	-4525.5781
2002	28348.0000	32806.0039	-4458.0039
2003	28855.0000	33786.4219	-4931.4219
Mean	22443.7895	23742.7390	-1298.9495

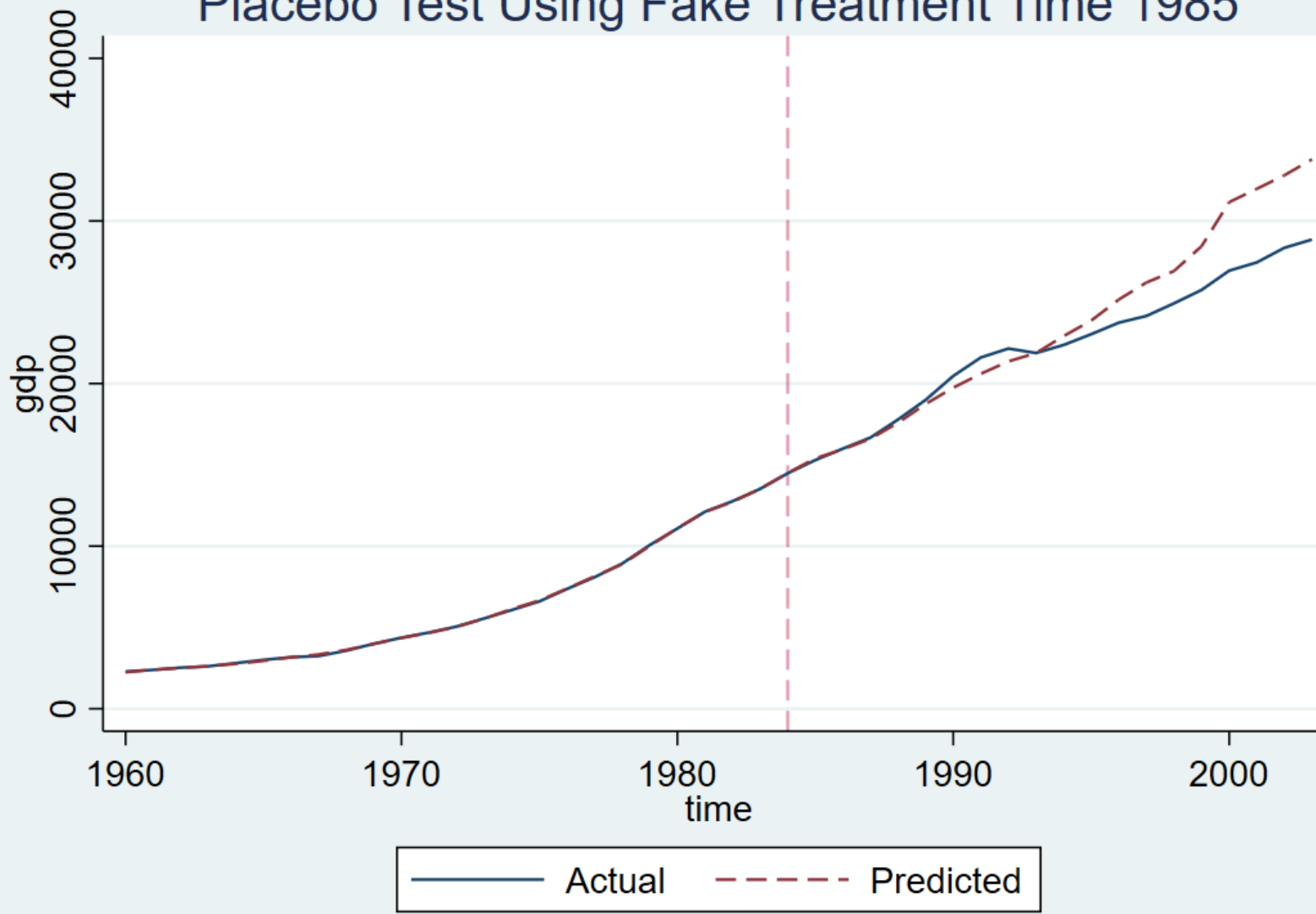
Note: The average treatment effect over the post-treatment periods is **-1298.9495**.

Placebo Test Using Fake Treatment Time 1980

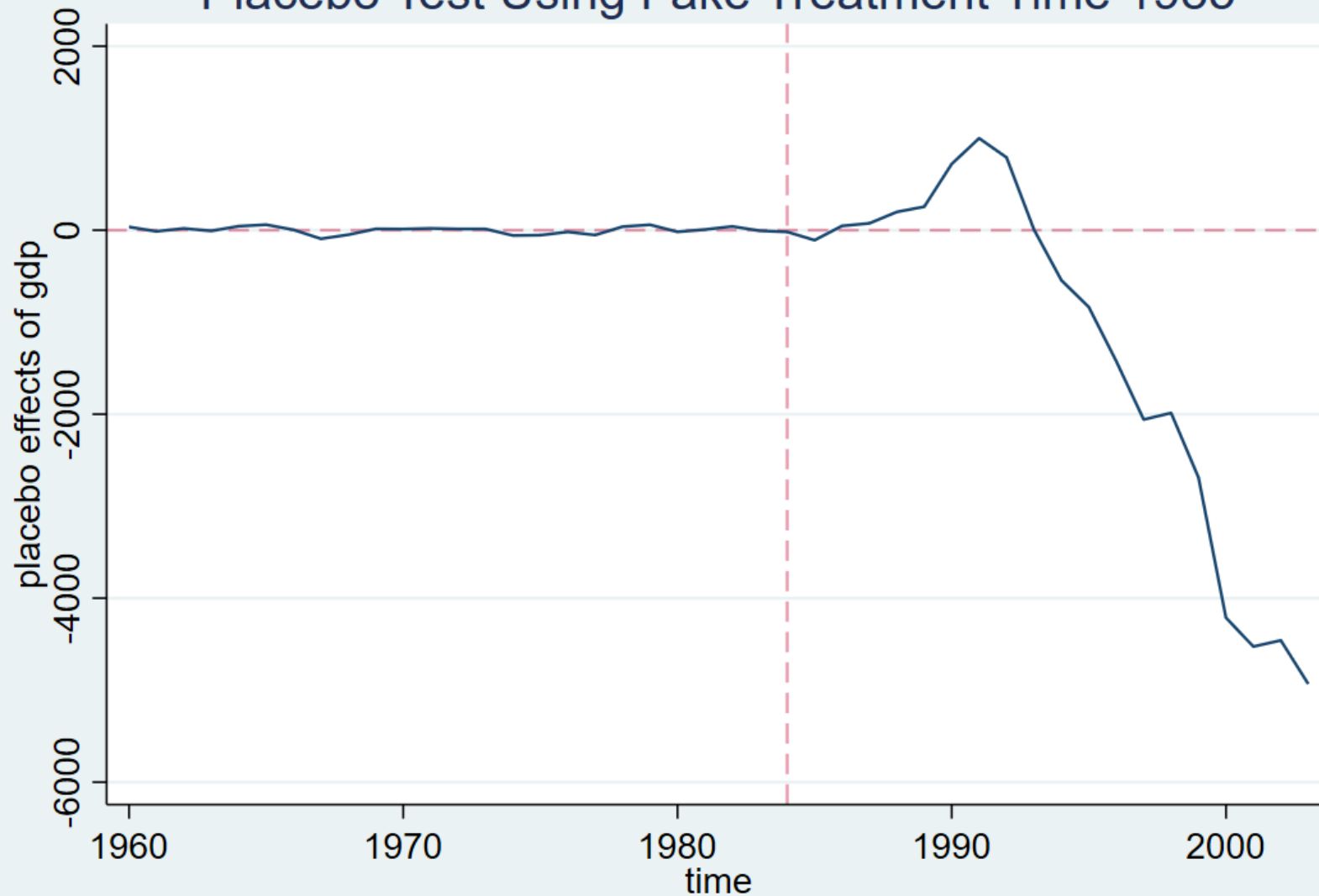




Placebo Test Using Fake Treatment Time 1985



Placebo Test Using Fake Treatment Time 1985



Reference

- Abadie, Alberto, Alexis Diamond, and Jens Hainmueller. 2015. Comparative Politics and the Synthetic Control Method. *American Journal of Political Science* 59(2): 495-510.
- Hsiao, Cheng, H. Steve Ching, and Shui Ki Wan. 2012. A Panel Data Approach for Program Evaluation: Measuring the Benefits of Political and Economic Integration of Hong Kong with Mainland China. *Journal of Applied Econometrics* 27(5): 705-740.
- Hsiao, Cheng, and Qiankun Zhou. 2019. Panel Parametric, Semiparametric, and Nonparametric Construction of Counterfactuals. *Journal of Applied Econometrics* 34(4): 463-481.

Reference

- Furnival, George M., and Robert W. Wilson, Jr. 1974. Regressions by Leaps and Bounds. *Technometrics* 16(4): 499-511.
- Guanpeng Yan, and Qiang Chen. 2021. Regression Control Method with Stata. Shandong University working paper.

Thank for Listening

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颜冠鹏

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