

Bacon decomposition for understanding differences-in-differences with variation in treatment timing

July 11, 2019

Stata Conference

Andrew Goodman-Bacon (Vanderbilt University)

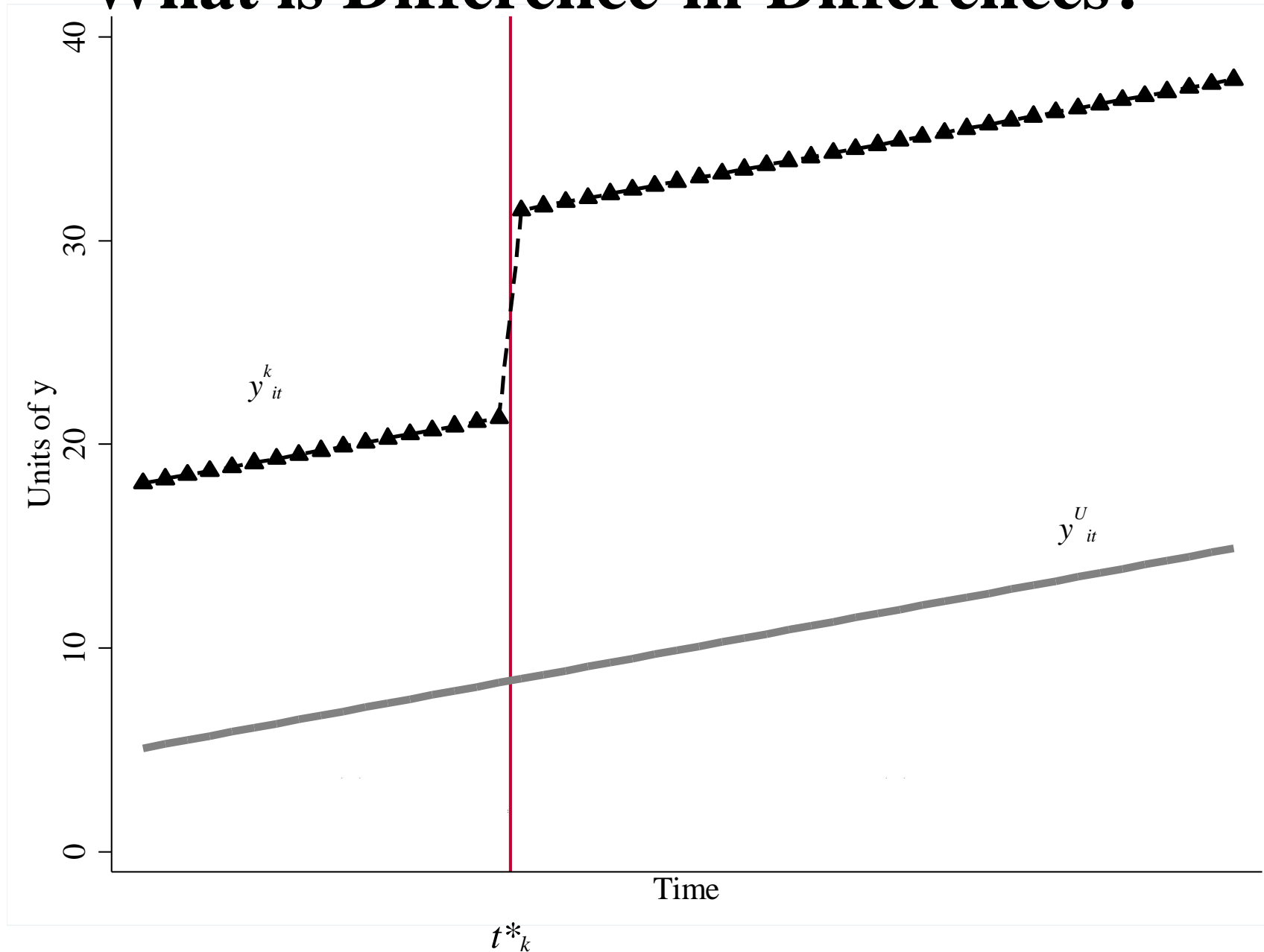
Austin Nichols (Abt Associates)

Thomas Goldring (University of Michigan)

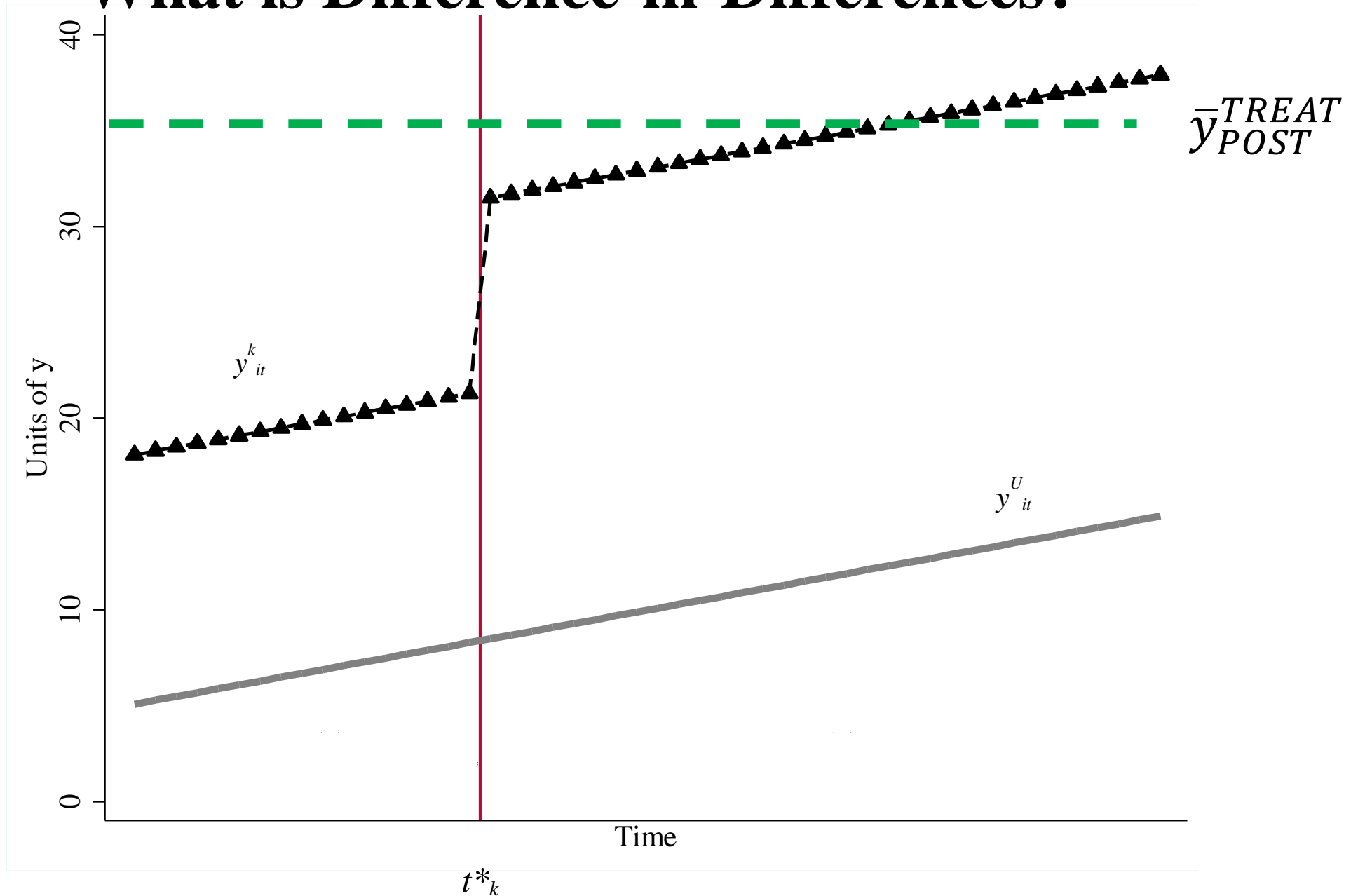
Overview

- In canonical difference-in-differences (DD), the regression version = function of pre/post and treat/control means.
- When treatment turns on at different times, the regression DD coefficient is a weighted average of canonical “2x2” DDs (Goodman-Bacon 2018)
 - Shows where such DDs “come from”
- This command calculates the component DDs and their weights, plots them (ie. shows variation), compares specifications
 - Future: conducts balance tests, analyzes estimand,

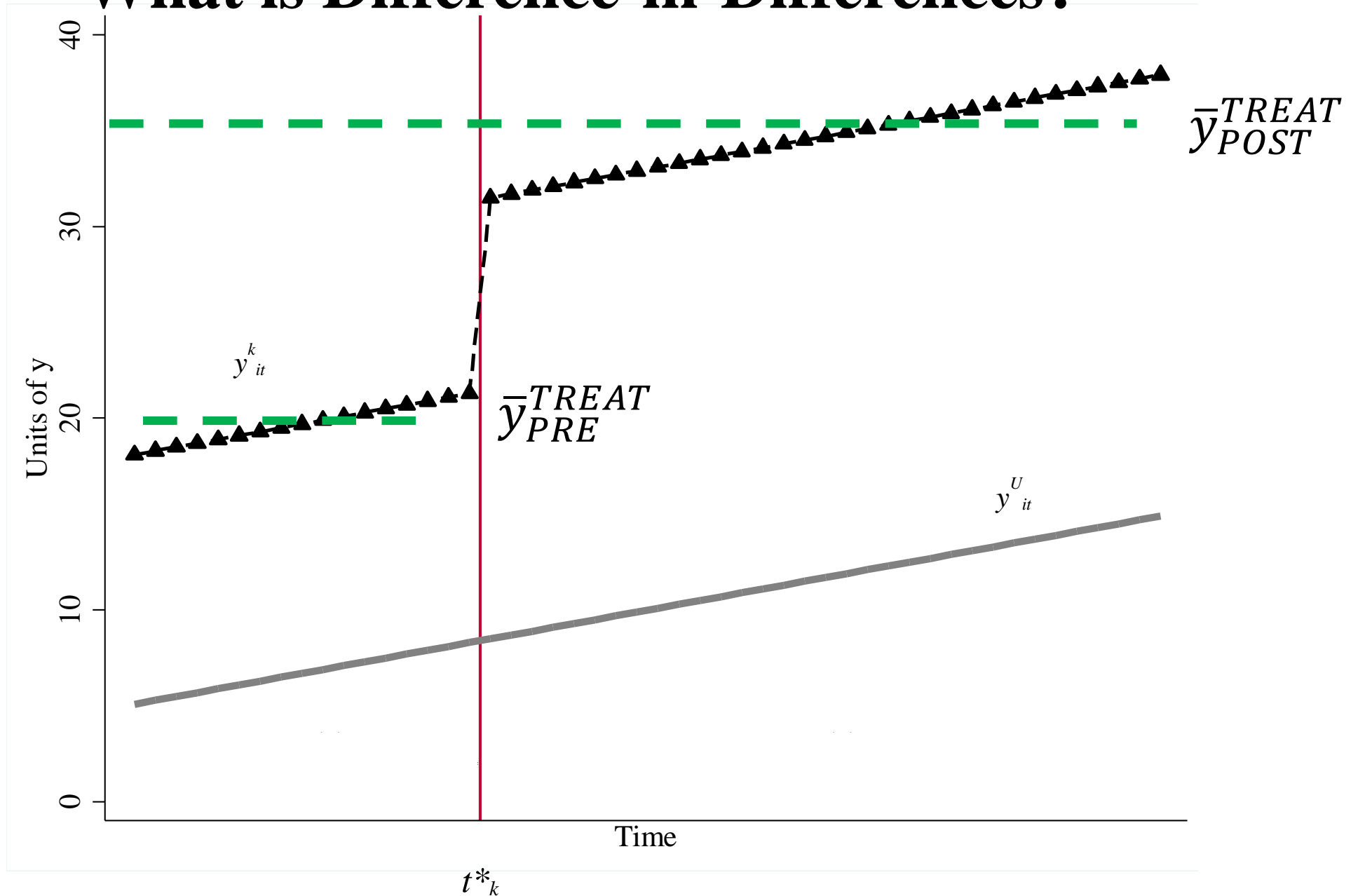
What is Difference-in-Differences?



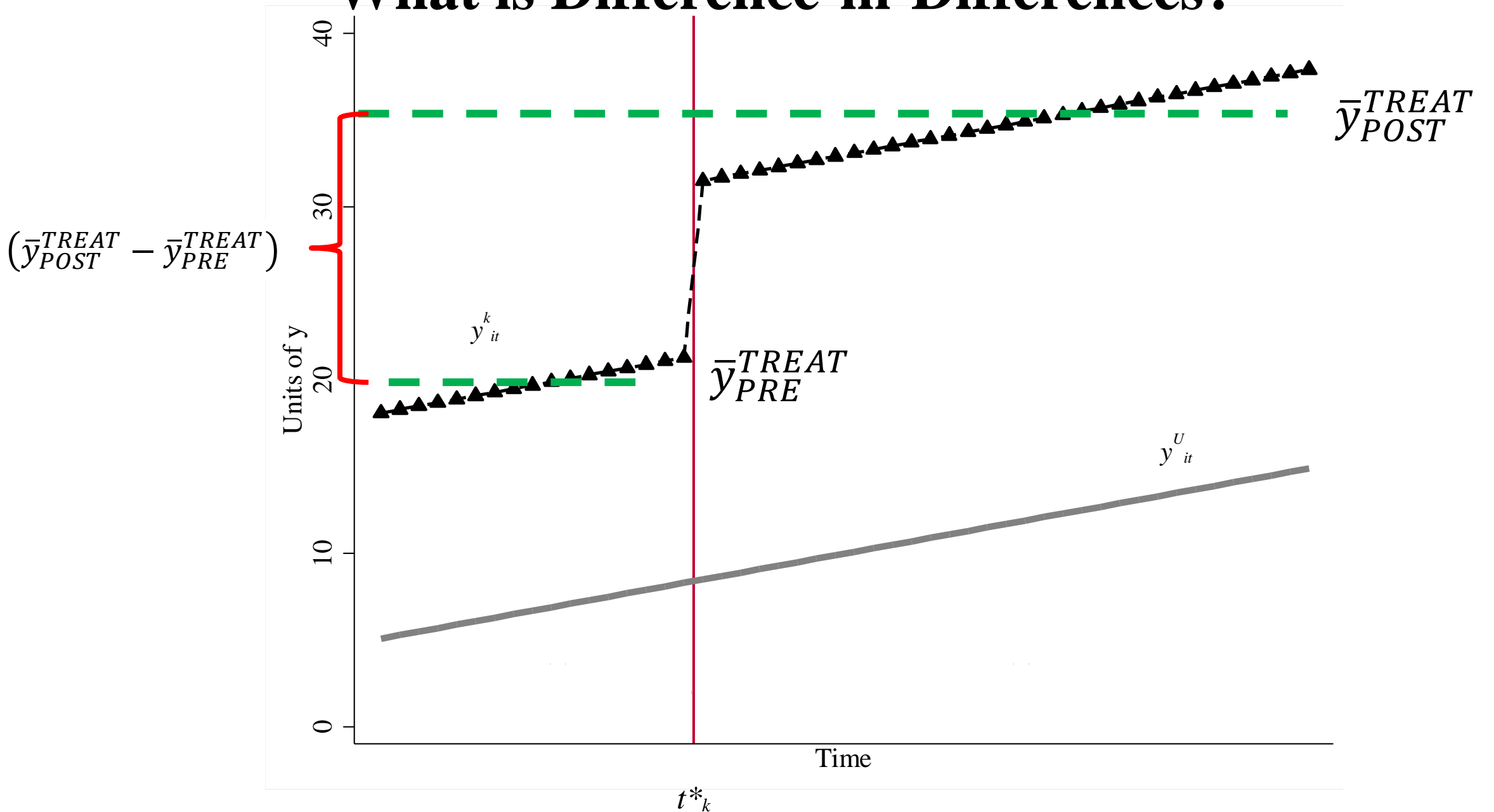
What is Difference-in-Differences?



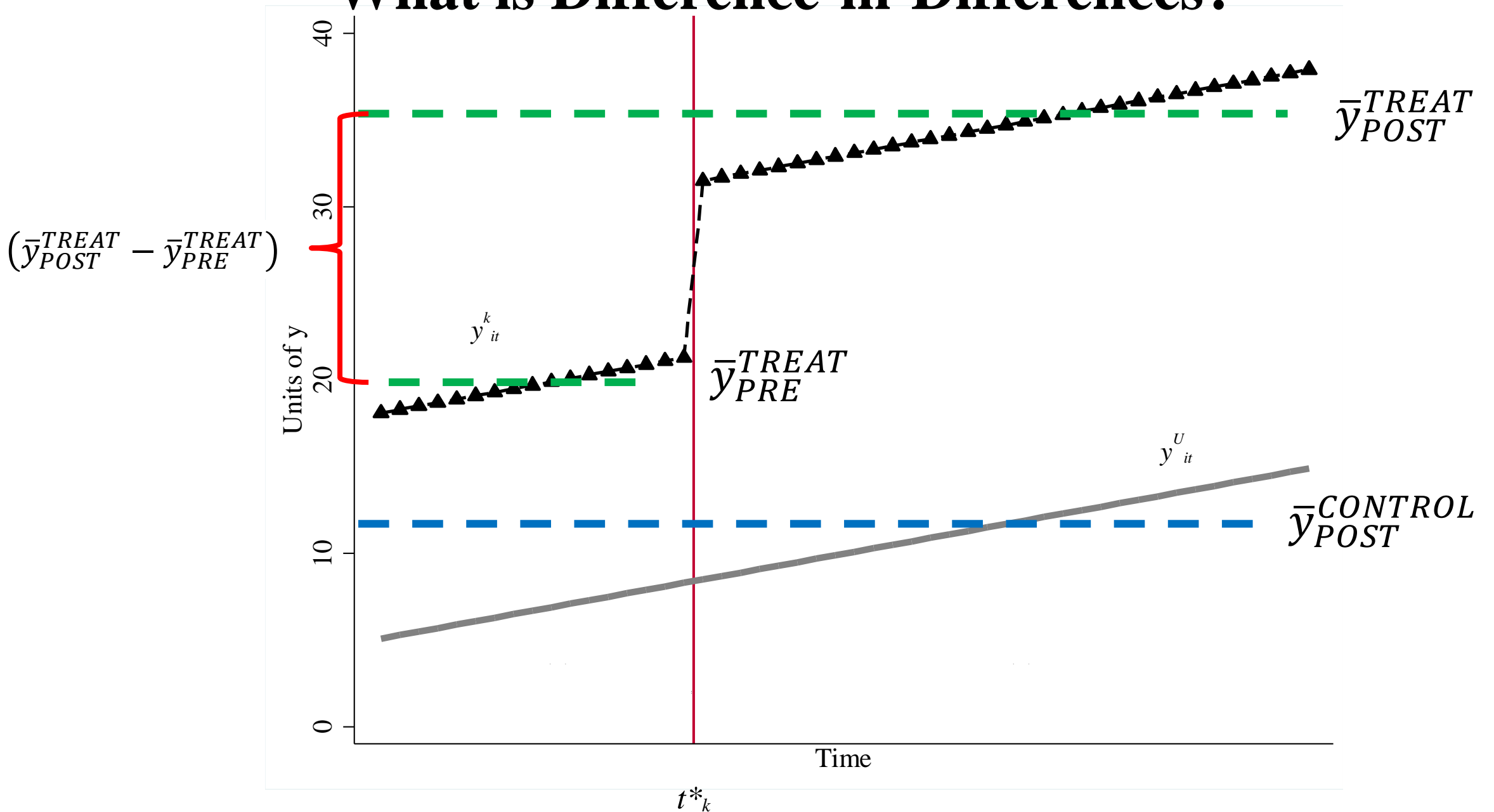
What is Difference-in-Differences?



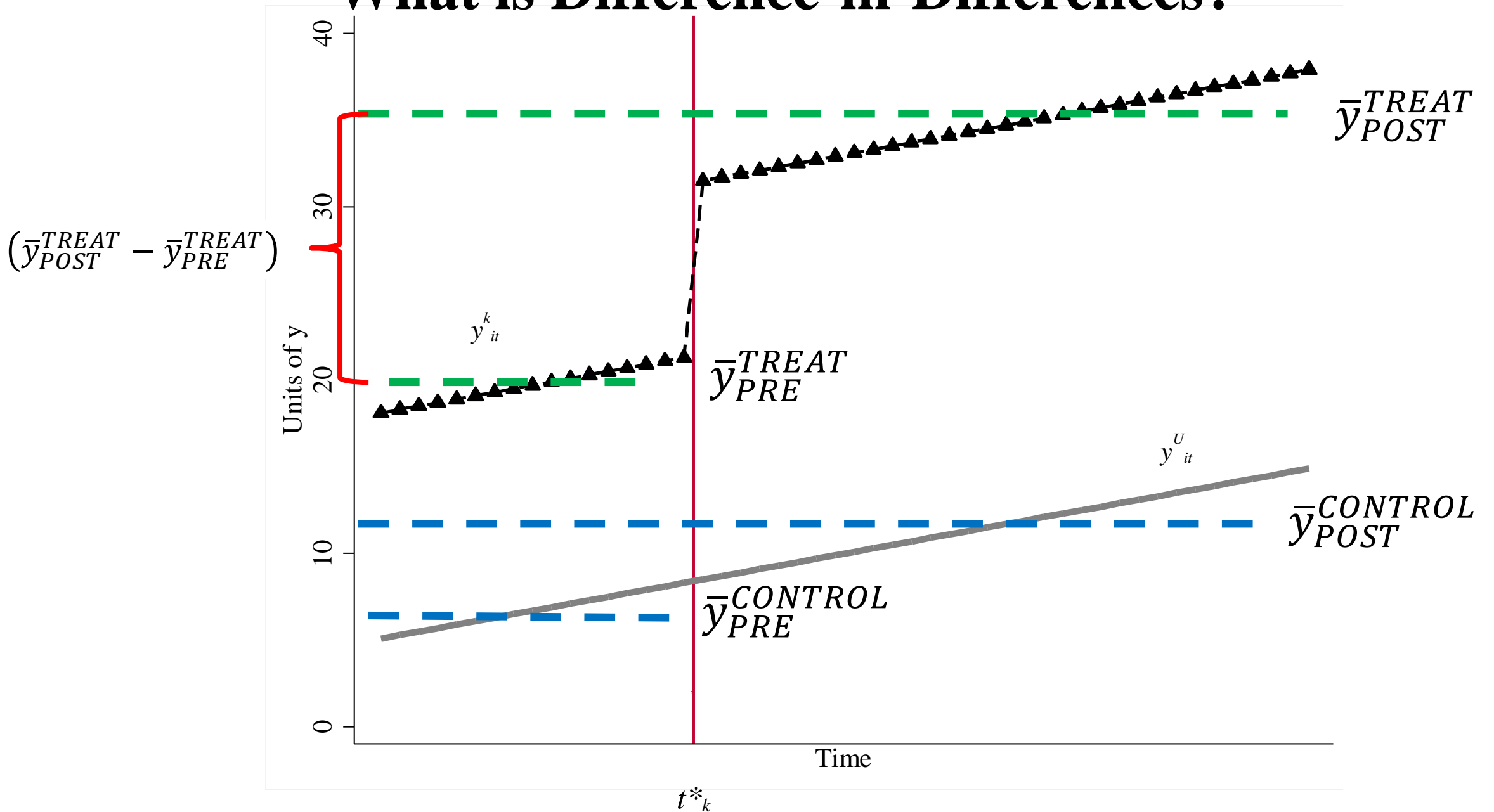
What is Difference-in-Differences?



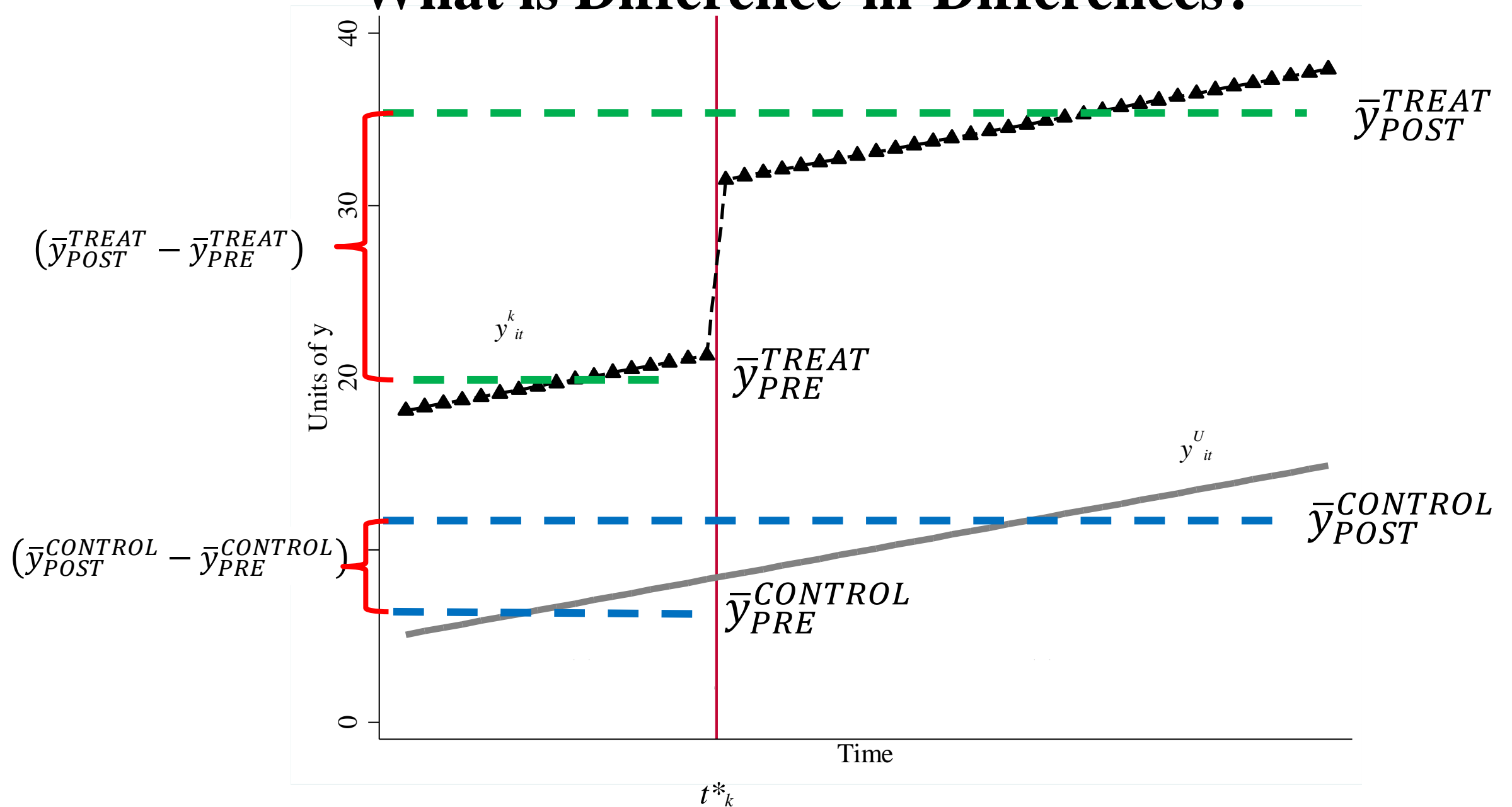
What is Difference-in-Differences?



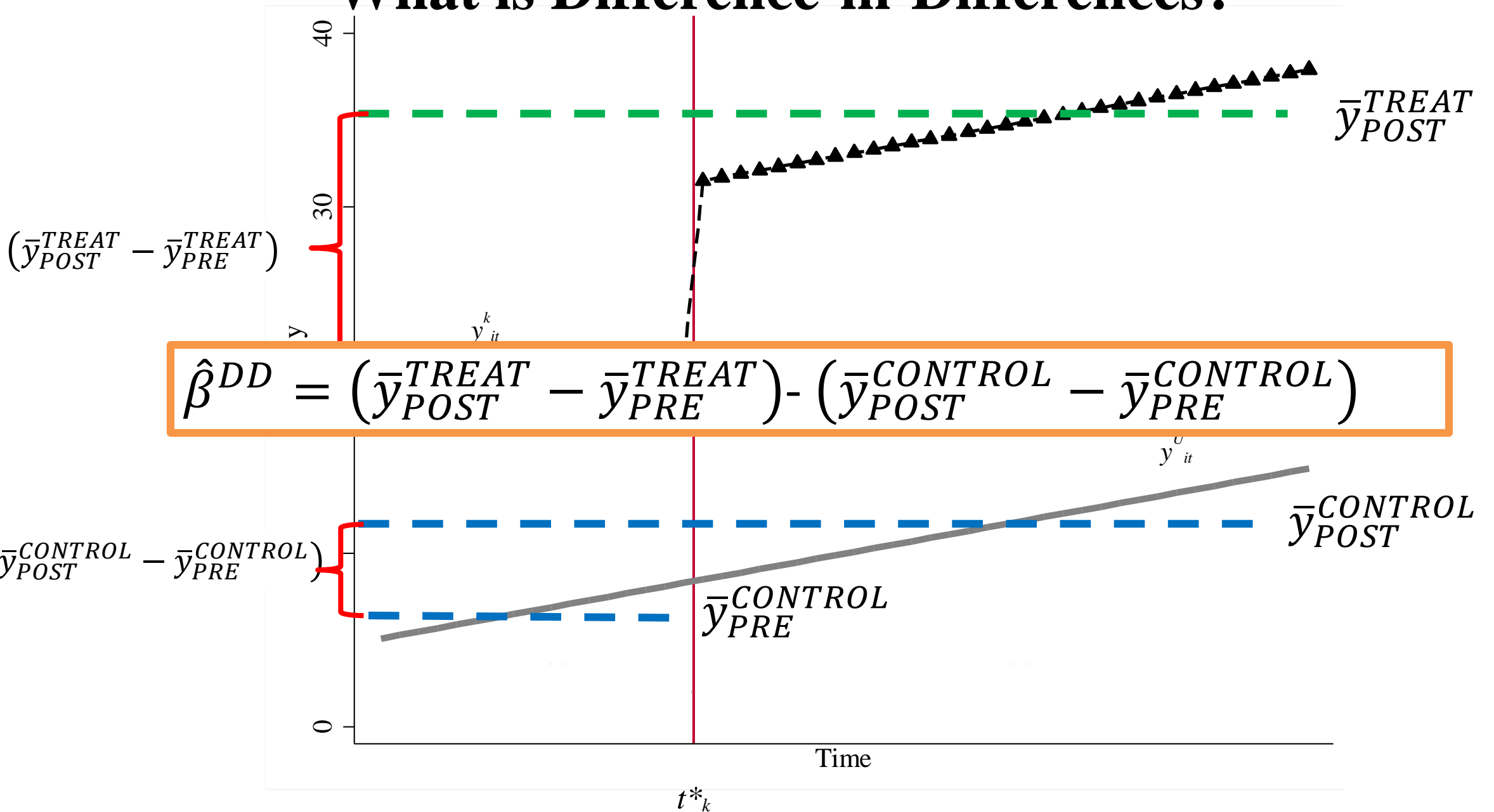
What is Difference-in-Differences?



What is Difference-in-Differences?



What is Difference-in-Differences?



What is Difference-in-Differences?

Wooldridge (2002):

$$\hat{\delta}_1 = (\bar{y}_{B,2} - \bar{y}_{B,1}) - (\bar{y}_{A,2} - \bar{y}_{A,1}) \quad (6.32)$$

This estimator has been labeled the **difference-in-differences (DID)** estimator in the recent program evaluation literature, although it has a long history in analysis of variance.

What is Difference-in-Differences?

Cameron and Trivedi (2007):

Then the OLS estimator reduces to

$$\hat{\phi} = \Delta \bar{y}^{\text{tr}} - \Delta \bar{y}^{\text{nt}}. \quad (22.43)$$

This estimator is called the **differences-in-differences (DID) estimator**, since one estimates the time difference for the treated and untreated groups and then takes the difference in the time differences.

What is Difference-in-Differences?

Angrist and Pischke (2009):

The population difference-in-differences,

$$\{E[Y_{ist}|s = NJ, t = Nov] - E[Y_{ist}|s = NJ, t = Feb]\} \\ - \{E[Y_{ist}|s = PA, t = Nov] - E[Y_{ist}|s = PA, t = Feb]\} = \delta,$$

What is Difference-in-Differences?

Imbens and Wooldridge (2007):

for those observations in the treatment group in the second period. The difference-in-differences estimate is

$$\hat{\delta}_1 = (\bar{y}_{B,2} - \bar{y}_{B,1}) - (\bar{y}_{A,2} - \bar{y}_{A,1}). \quad (1.2)$$

What is Difference-in-Differences?

Angrist and Krueger (1999):

$$\{E[Y_i | c = \text{Miami}, t = 1981] - E[Y_i | c = \text{Comparison}, t = 1981]\} \\ - \{E[Y_i | c = \text{Miami}, t = 1979] - E[Y_i | c = \text{Comparison}, t = 1979]\} = \delta. \quad (21)$$

What is Difference-in-Differences?

Heckman, LaLonde and Smith (1999):

then the *difference-in-differences* estimator given by

$$(\bar{Y}_{1t} - \bar{Y}_{0t'})_1 - (\bar{Y}_{0t} - \bar{Y}_{0t'})_0, \quad t > k > t'$$

What is Difference-in-Differences?

Meyer (1995):

In this case, an unbiased estimate of β can be obtained by difference in differences as

$$\begin{aligned}\widehat{\beta}_{\text{dd}} &= \Delta \bar{y}_1 - \Delta \bar{y}_0 \\ &= \bar{y}_1^1 - \bar{y}_0^1 - (\bar{y}_1^0 - \bar{y}_0^0),\end{aligned}\tag{4}$$

What is Difference-in-Differences?

Abadie (2005):

$D(i, 1) = 1$, and the individual-specific component, $\eta(i)$. This model is called “difference-in-differences” because under the identifying condition in equation (2) we have

$$\begin{aligned} \alpha = & \{E[Y(i, 1) \mid D(i, 1) = 1] - E[Y(i, 1) \mid D(i, 1) = 0]\} \\ & - \{E[Y(i, 0) \mid D(i, 1) = 1] - E[Y(i, 0) \mid D(i, 1) = 0]\}, \end{aligned} \tag{5}$$

What is Difference-in-Differences?

Athey and Imbens (2006):

i.e., $\varepsilon_i \perp (G_i, T_i)$, and is normalized to have mean zero. The standard DID estimand is

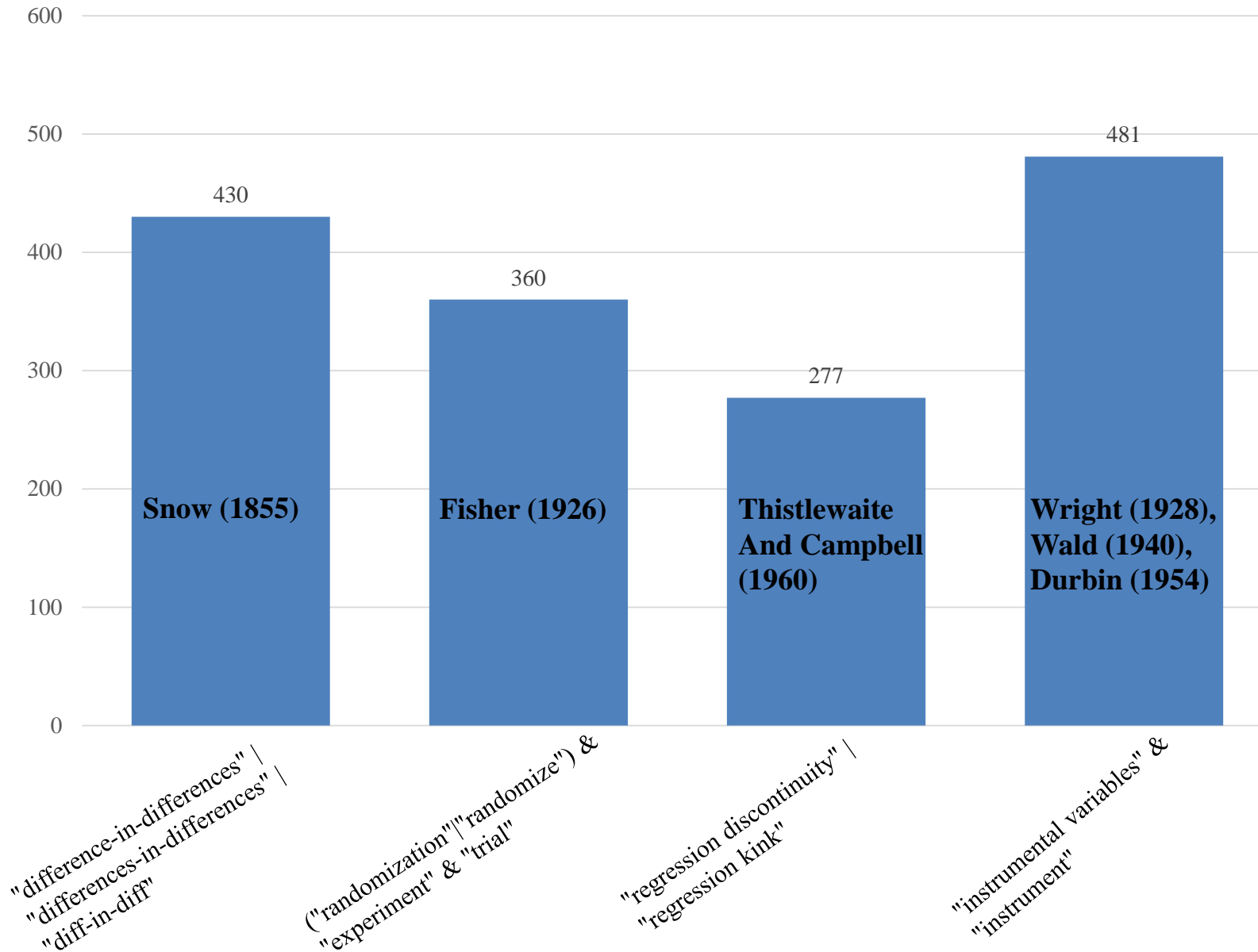
$$(2) \quad \tau^{\text{DID}} = [\mathbb{E}[Y_i | G_i = 1, T_i = 1] - \mathbb{E}[Y_i | G_i = 1, T_i = 0]] \\ - [\mathbb{E}[Y_i | G_i = 0, T_i = 1] - \mathbb{E}[Y_i | G_i = 0, T_i = 0]].$$

What is Difference-in-Differences?

DiNardo and Lee (2011):

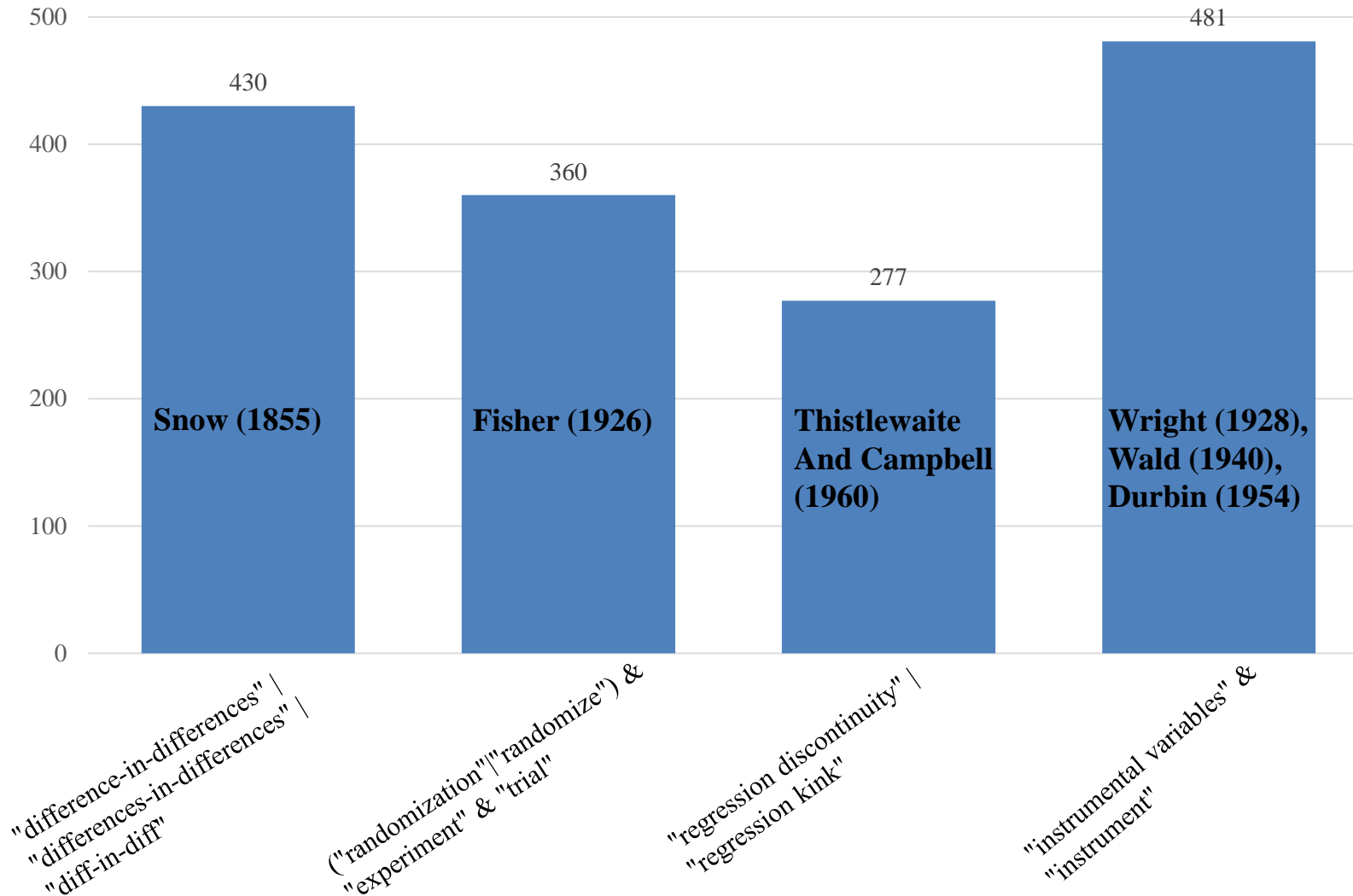
First, let us simplify the problem by considering the situation where the program was made available at only one point in time τ . This allows us to define $D = 1$ as those who were treated at time τ , and $D = 0$ as those who did not take up the program at that time.

Keywords in NBER Papers Since 2012



Keywords in NBER Papers Since 2012

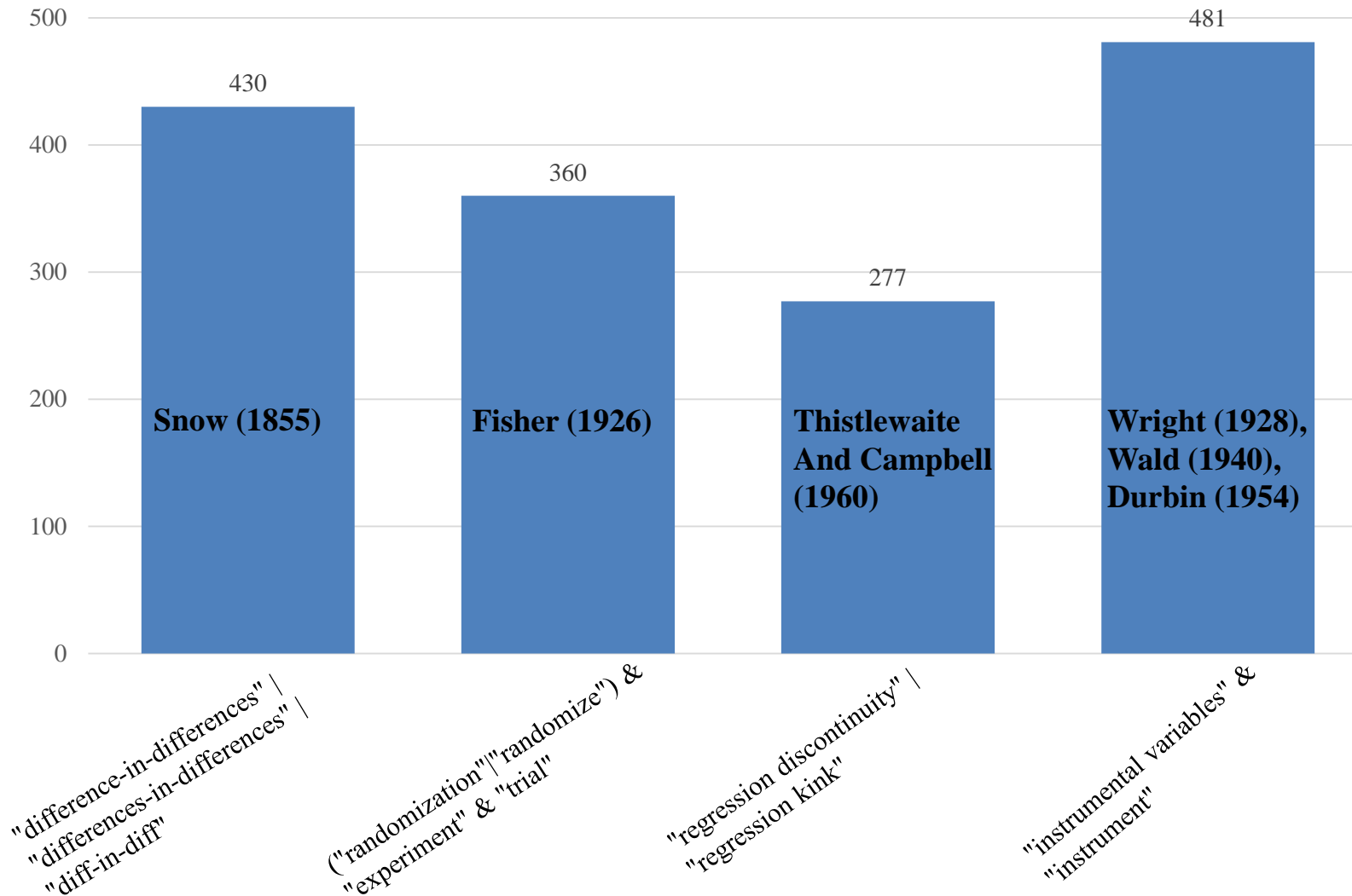
The last arrow in the quasi-experimental quiver is differences-in-differences, probably the most widely applicable design-based estimator.



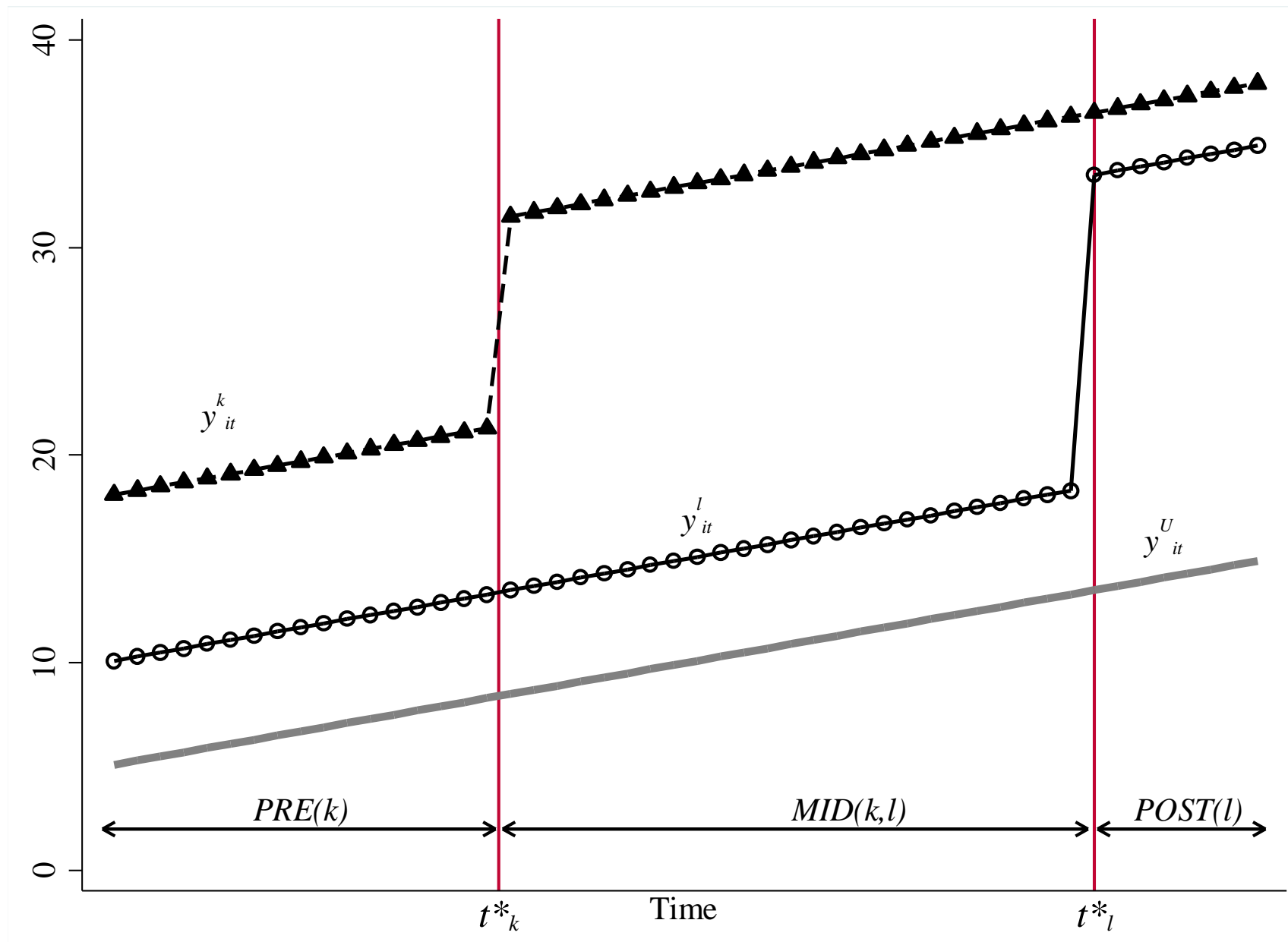
Keywords in NBER Papers Since 2012

2014/2015 AER/QJE/JPE/ReStud/JHE/JDE published 93 DD papers:

49% had timing variation



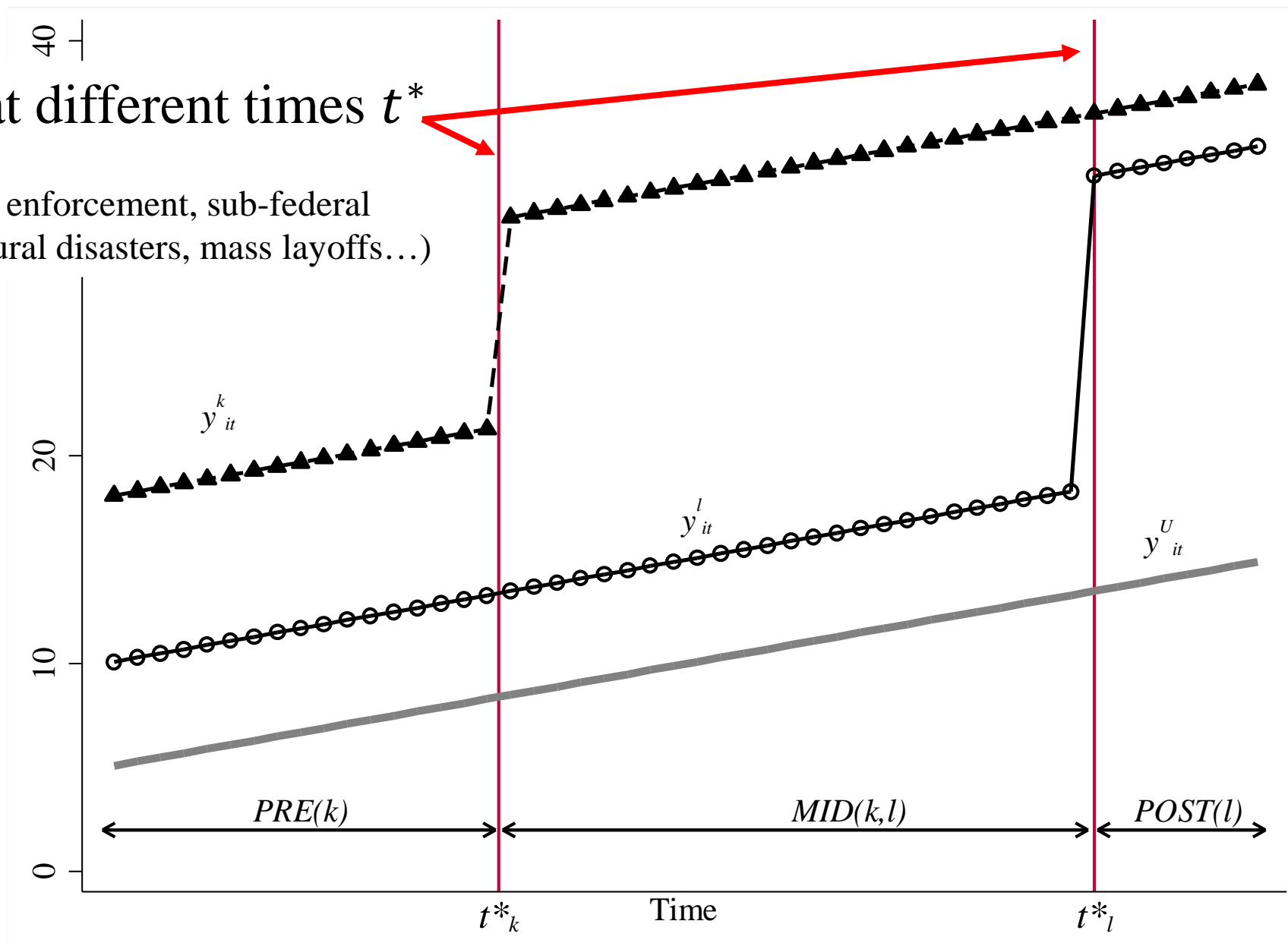
Variation in Timing



Variation in Timing

D_{it} turns on at different times t^*

(Federalism, judicial enforcement, sub-federal funding process, natural disasters, mass layoffs...)



Two-Way Fixed Effects Estimator

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

Unit fixed effects

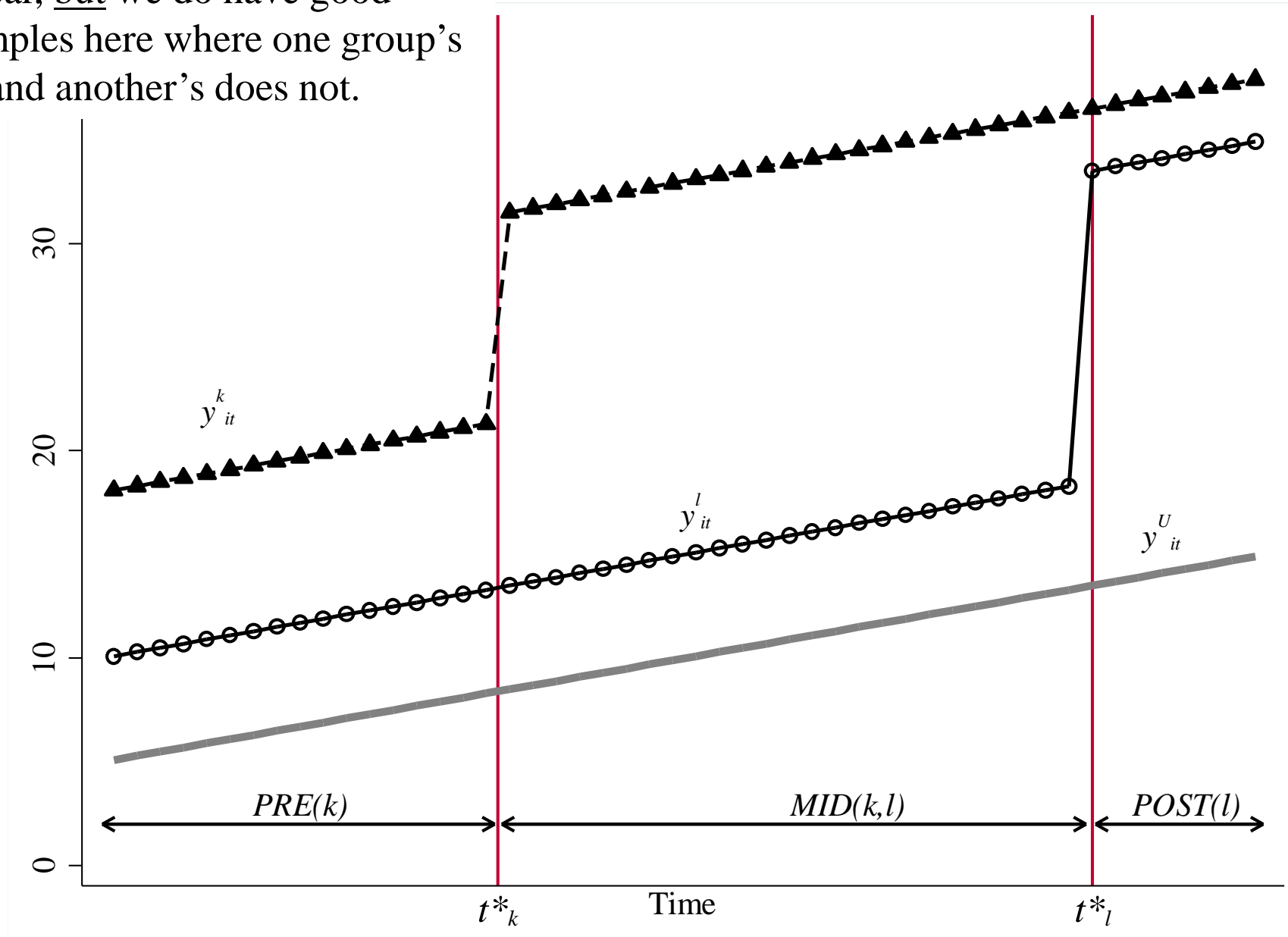
Time fixed effects

Treatment dummy

What is $\hat{\beta}^{DD}$?

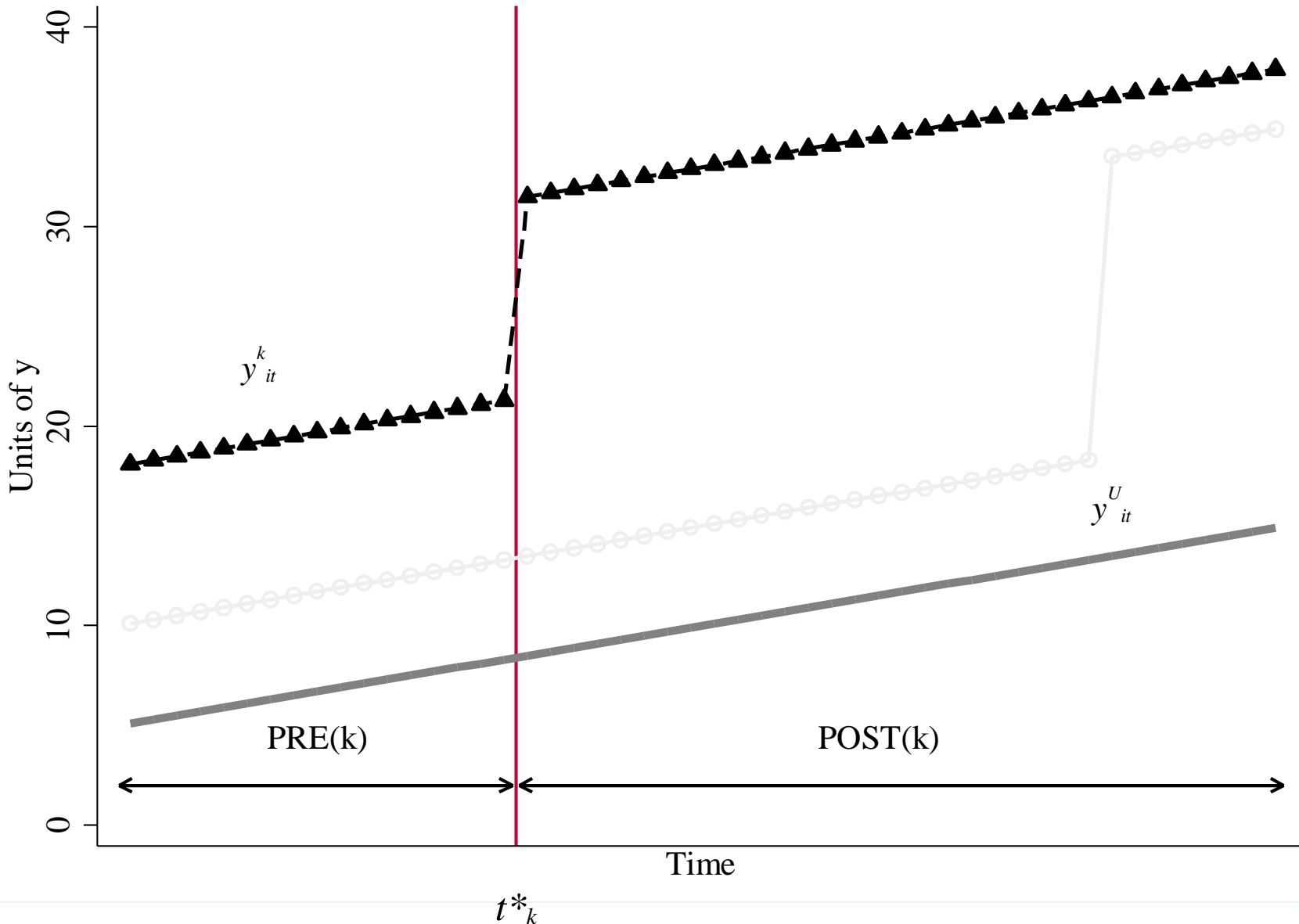
$$\hat{\beta}^{DD}?$$

This has been unclear, but we do have good intuition for subsamples here where one group's treatment changes and another's does not.



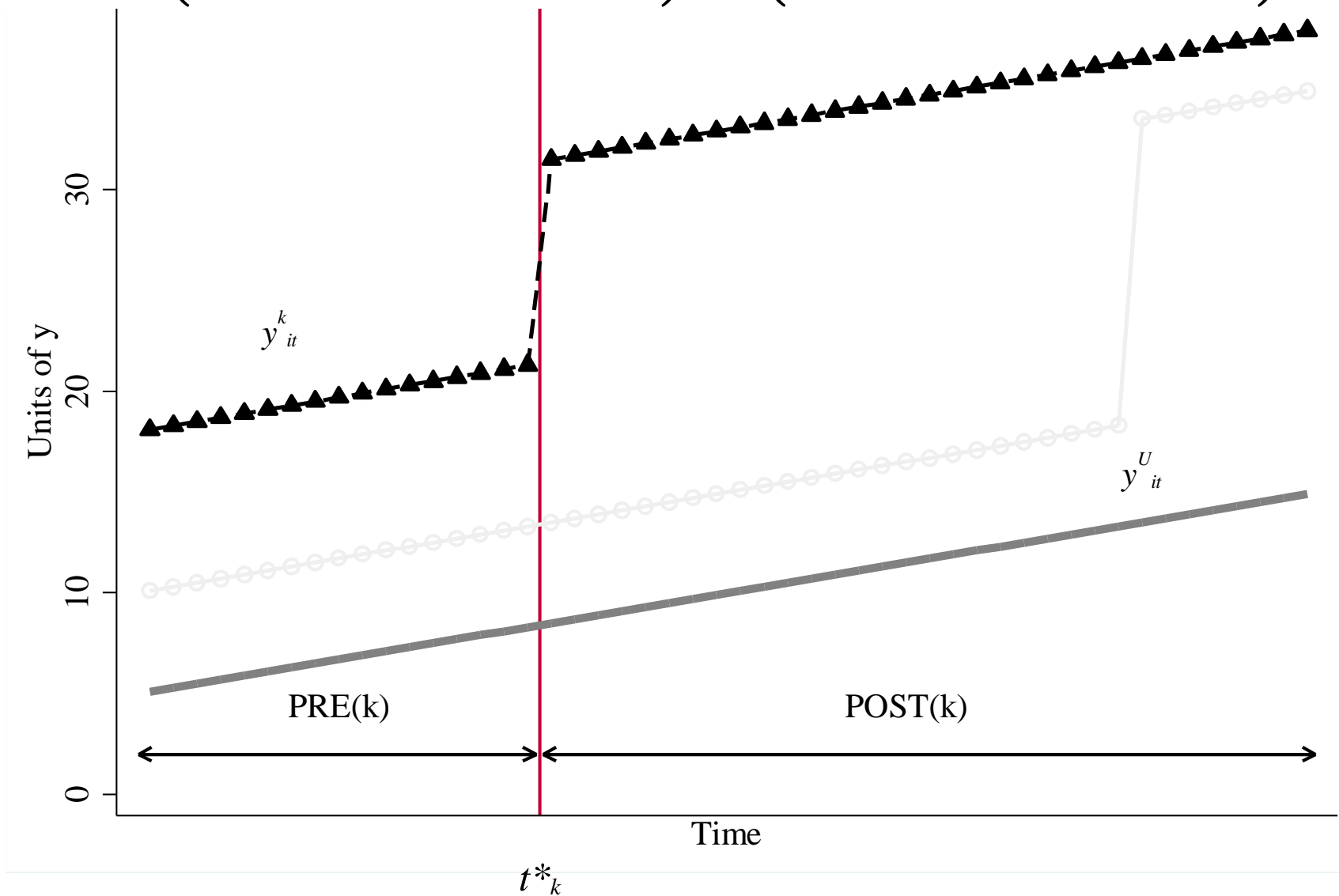
$$\hat{\beta}_{kU}^{2 \times 2}$$

A. Early Group vs. Untreated Group



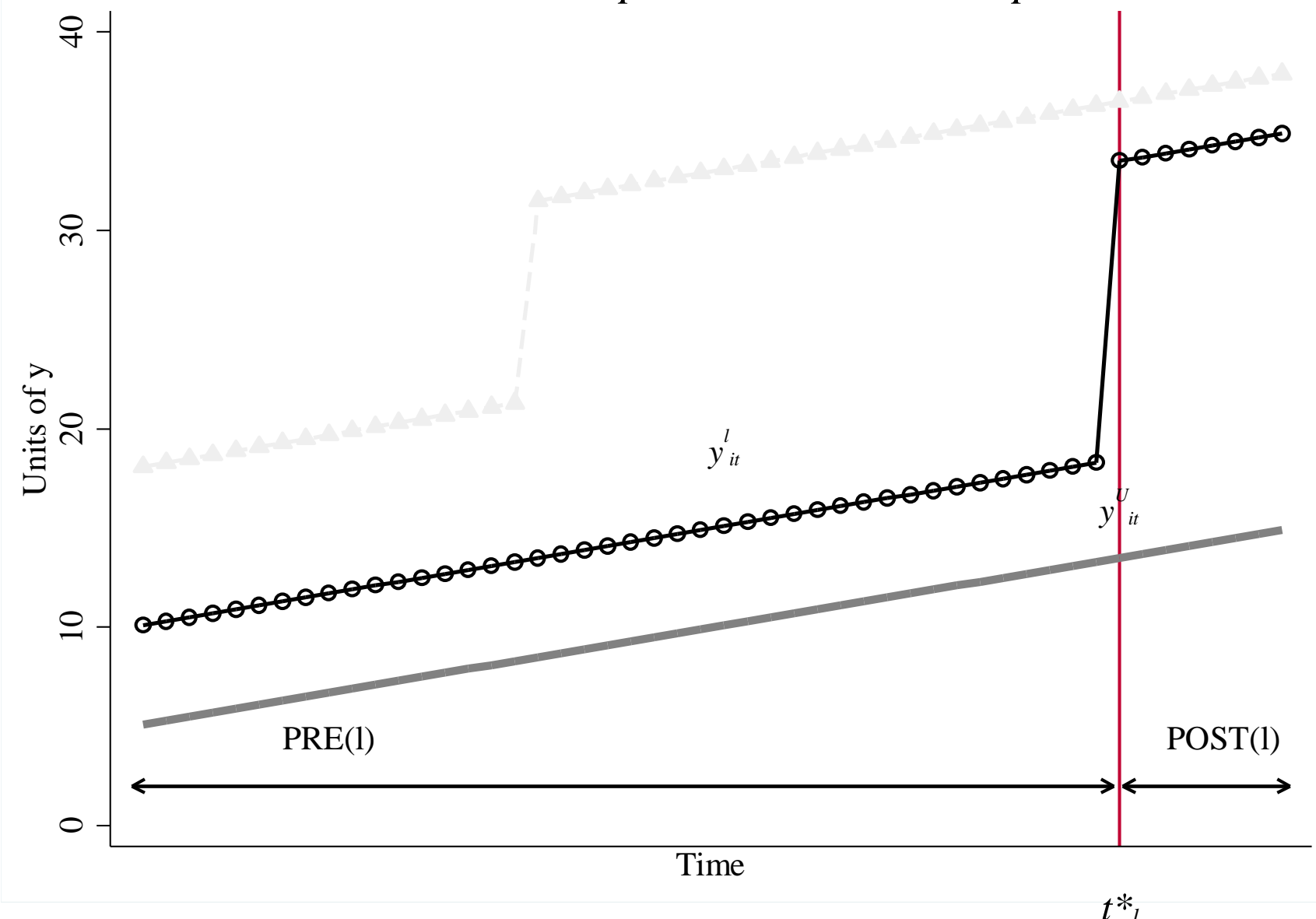
$$\hat{\beta}_{kU}^{2 \times 2}$$

$$\hat{\beta}_{kU}^{2 \times 2} = \left(\bar{y}_k^{POST(k)} - \bar{y}_k^{PRE(k)} \right) - \left(\bar{y}_U^{POST(k)} - \bar{y}_U^{PRE(k)} \right)$$



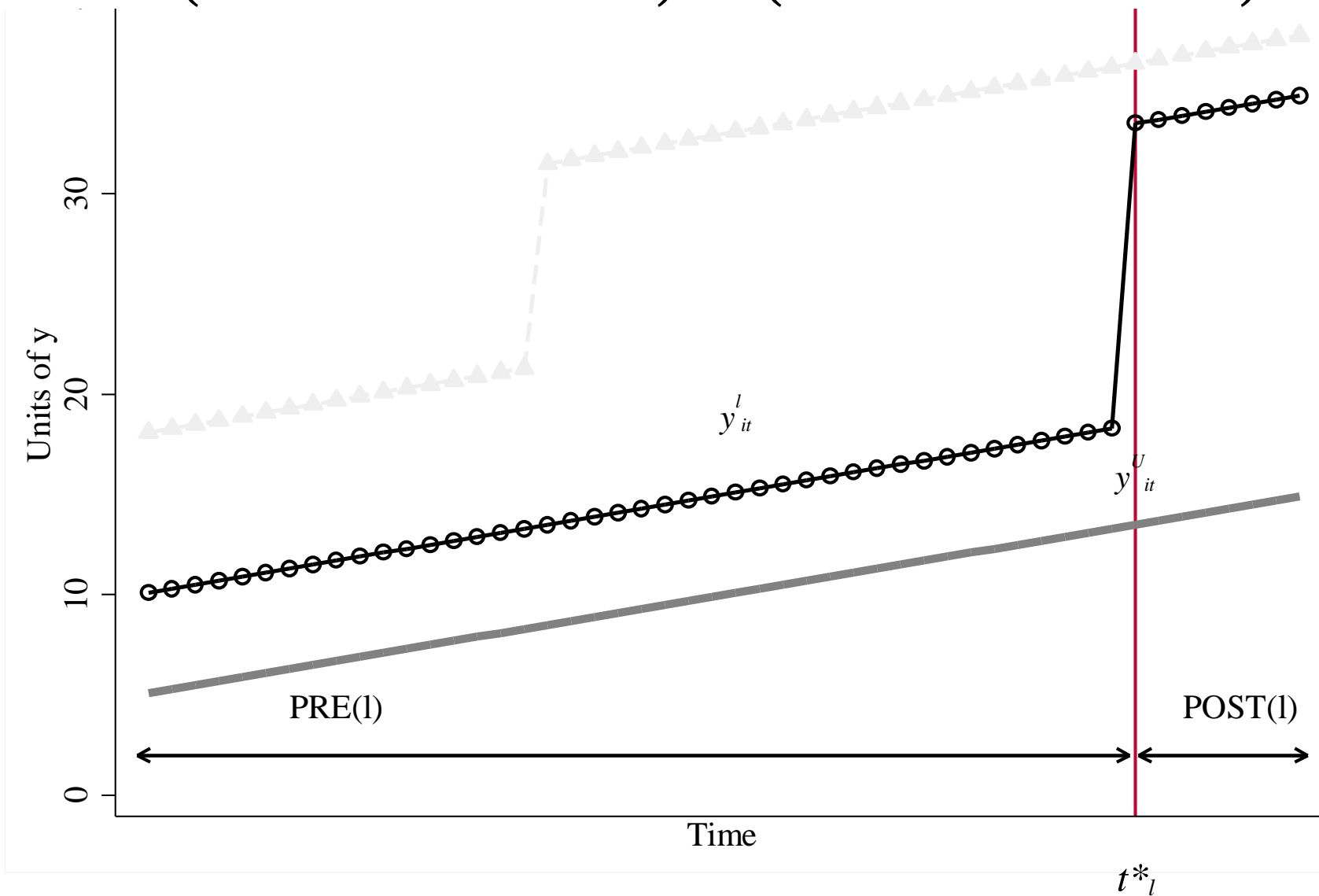
$$\hat{\beta}_{lU}^{2 \times 2}$$

B. Late Group vs. Untreated Group



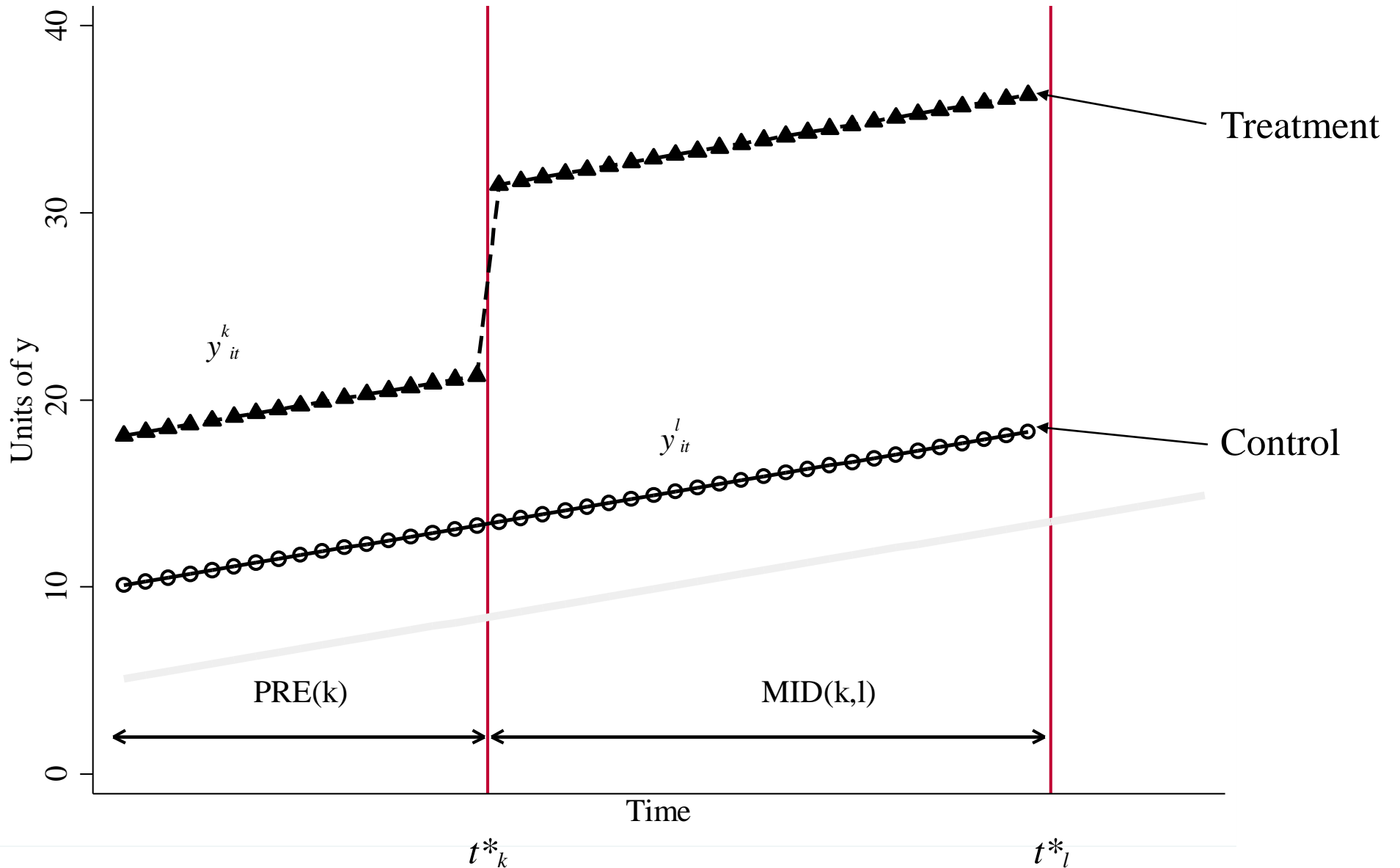
$$\hat{\beta}_{\ell U}^{2 \times 2}$$

$$\hat{\beta}_{\ell U}^{2 \times 2} = \left(\bar{y}_{\ell}^{POST(\ell)} - \bar{y}_{\ell}^{PRE(\ell)} \right) - \left(\bar{y}_U^{POST(\ell)} - \bar{y}_U^{PRE(\ell)} \right)$$



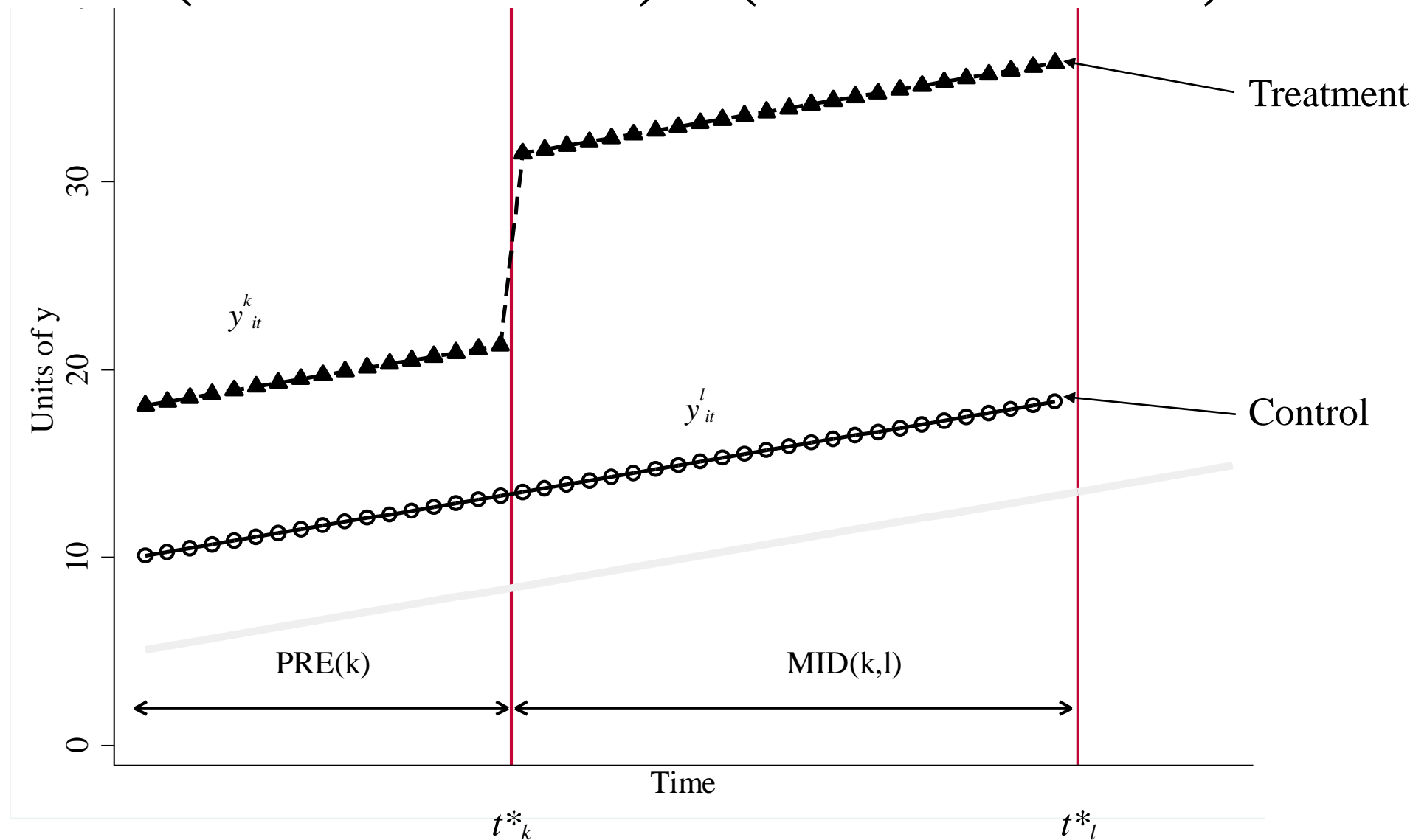
$$\hat{\beta}_{kl}^{2 \times 2, k}$$

C. Early Group vs. Late Group, before t^*_l



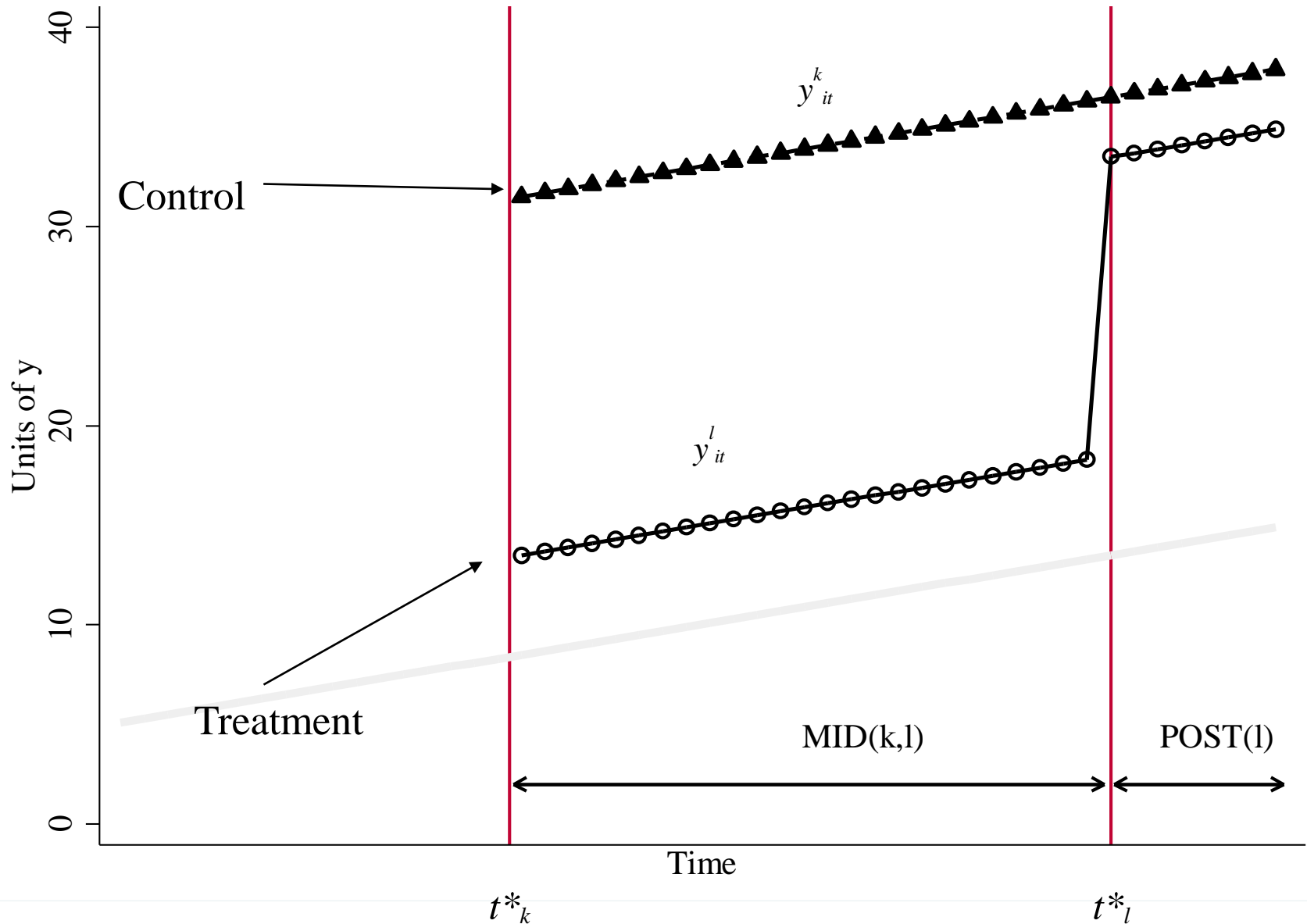
$$\hat{\beta}_{k\ell}^{2 \times 2, k}$$

$$\hat{\beta}_{k\ell}^{2 \times 2, k} = \left(\bar{y}_k^{MID(k,\ell)} - \bar{y}_k^{PRE(k)} \right) - \left(\bar{y}_\ell^{MID(k,\ell)} - \bar{y}_\ell^{PRE(k)} \right)$$

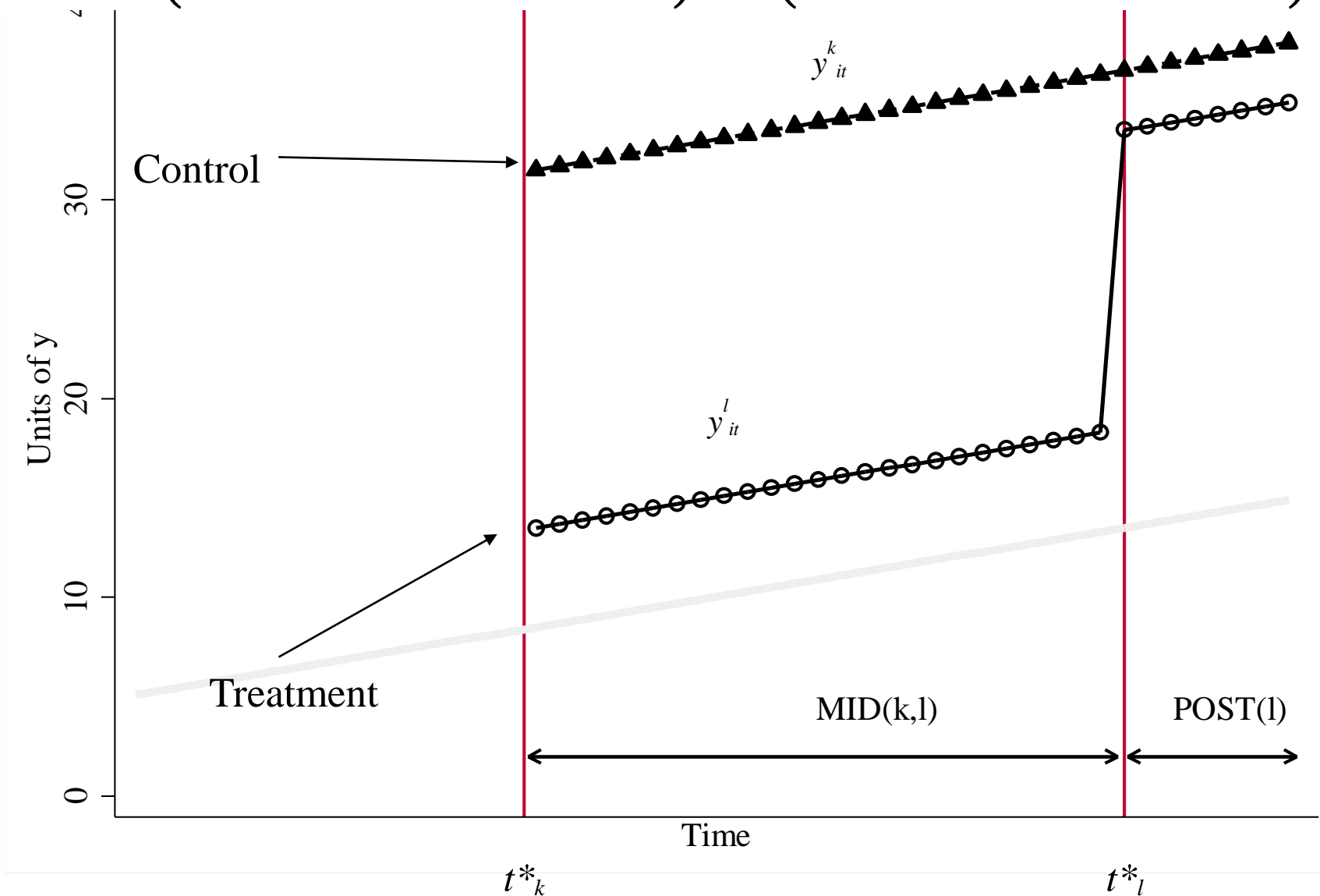


$$\widehat{\beta}_{k\ell}^{2 \times 2, \ell}$$

*D. Late Group vs. Early Group, after t^*_k*



$$\hat{\beta}_{k\ell}^{2 \times 2, \ell} = \left(\bar{y}_{\ell}^{POST(\ell)} - \bar{y}_{\ell}^{MID(k, \ell)} \right) - \left(\bar{y}_k^{POST(\ell)} - \bar{y}_k^{MID(k, \ell)} \right)$$



Difference-in-Differences Decomposition Theorem (3 Group Case)

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

For three groups:

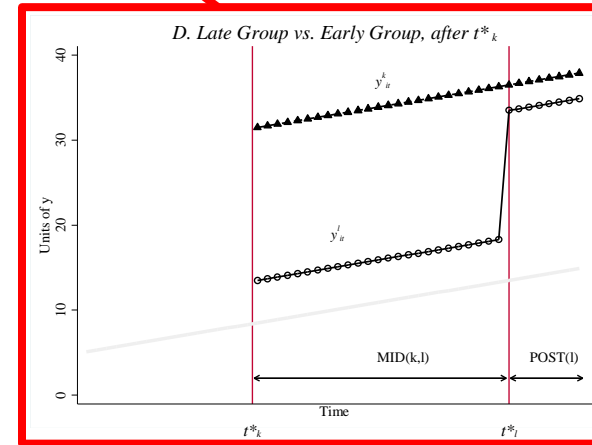
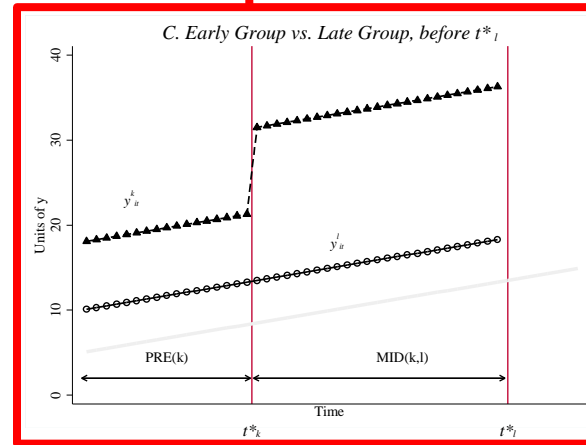
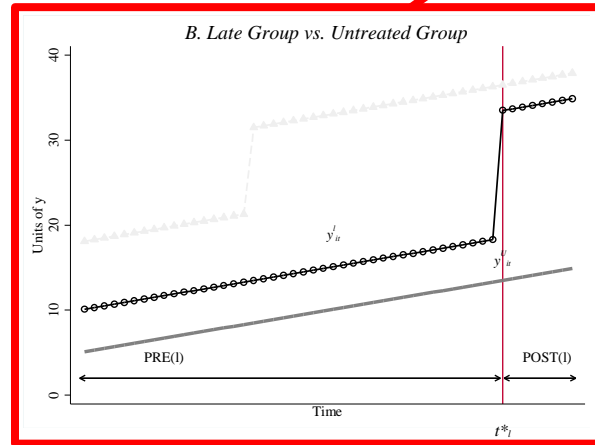
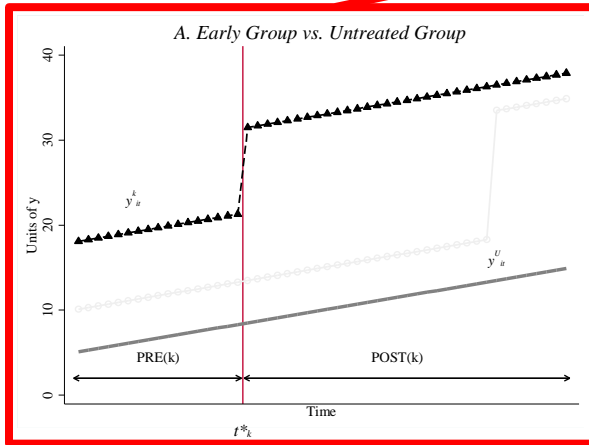
$$\hat{\beta}^{DD} = s_{kU} \hat{\beta}_{kU}^{2 \times 2} + s_{\ell U} \hat{\beta}_{\ell U}^{2 \times 2} + s_{k\ell}^k \hat{\beta}_{k\ell}^{2 \times 2, k} + s_{k\ell}^{\ell} \hat{\beta}_{k\ell}^{2 \times 2, \ell}$$

Difference-in-Differences Decomposition Theorem (3 Group Case)

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

For three groups:

$$\hat{\beta}^{DD} = s_{kU} \hat{\beta}_{kU}^{2 \times 2} + s_{lU} \hat{\beta}_{lU}^{2 \times 2} + s_{kl}^k \hat{\beta}_{kl}^{2 \times 2, k} + s_{kl}^l \hat{\beta}_{kl}^{2 \times 2, l}$$

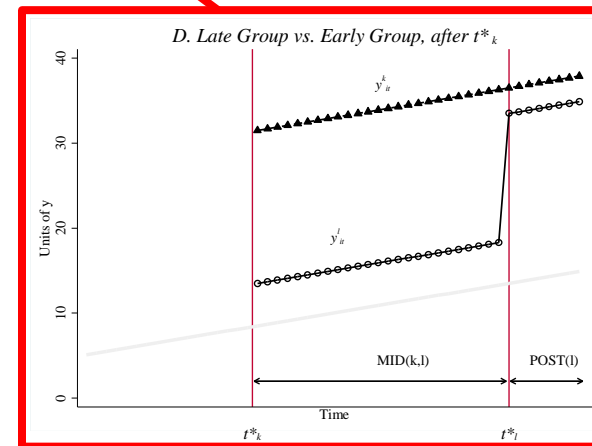
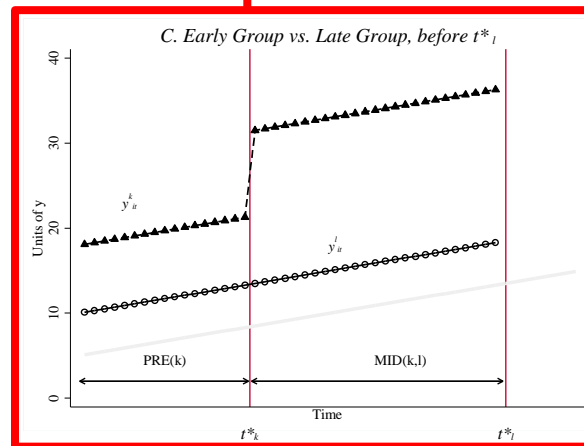
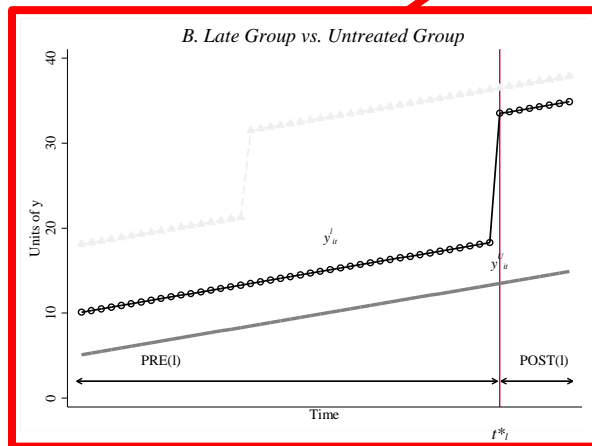
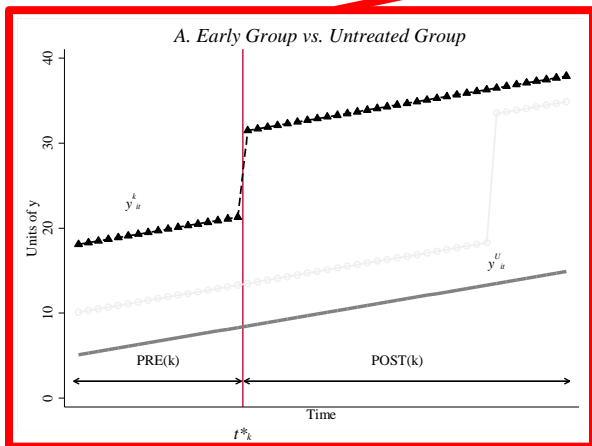


Difference-in-Differences Decomposition Theorem (3 Group Case)

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

For three groups:

$$\hat{\beta}^{DD} = s_{kU} \hat{\beta}_{kU}^{2 \times 2} + s_{\ell U} \hat{\beta}_{\ell U}^{2 \times 2} + s_{k\ell}^k \hat{\beta}_{k\ell}^{2 \times 2, k} + s_{k\ell}^{\ell} \hat{\beta}_{k\ell}^{2 \times 2, \ell}$$



2x2 DDs: subsamples with two groups (treat/control) and two periods (pre/post)

What do we learn from the 2x2 DDs?

1. We didn't know what comparisons were being made:

“switchers vs untreated”?

“early vs late”?

“late vs early” (this is less obvious)?

It's all of those.

2. “What is the control group?”

Every group acts as a control (sometimes).

3. Clarifies theory

We understand the estimand (ATET) and ID assumption (common trends) for each 2x2; late vs. early comparisons are biased if effects vary over time.

Difference-in-Differences Decomposition Theorem (3 Group Case)

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

For three groups:

$$\hat{\beta}^{DD} = s_{kU} \hat{\beta}_{kU}^{2 \times 2} + s_{\ell U} \hat{\beta}_{\ell U}^{2 \times 2} + s_{k\ell}^k \hat{\beta}_{k\ell}^{2 \times 2, k} + s_{k\ell}^\ell \hat{\beta}_{k\ell}^{2 \times 2, \ell}$$

Difference-in-Differences Decomposition Theorem (3 Group Case)

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

For three groups:

$$\hat{\beta}^{DD} = s_{kU} \hat{\beta}_{kU}^{2 \times 2} + s_{\ell U} \hat{\beta}_{\ell U}^{2 \times 2} + s_{k\ell}^k \hat{\beta}_{k\ell}^{2 \times 2, k} + s_{k\ell}^\ell \hat{\beta}_{k\ell}^{2 \times 2, \ell}$$

A. Early Group vs. Untreated Group

Size:

$$(n_k + n_U)^2$$

×

Variance:

$$n_{kU} (1 - n_{kU}) \bar{D}_k (1 - \bar{D}_k)$$

t_k^*

B. Late Group vs. Untreated Group

Size:

$$(n_\ell + n_U)^2$$

×

Variance:

$$n_{\ell U} (1 - n_{\ell U}) \bar{D}_\ell (1 - \bar{D}_\ell)$$

t_ℓ^*

C. Early Group vs. Late Group, before t_k^*

Size:

$$((n_k + n_\ell)(1 - \bar{D}_\ell))^2$$

×

Variance:

$$n_{k\ell} (1 - n_{k\ell}) \frac{\bar{D}_k - \bar{D}_\ell}{1 - \bar{D}_\ell} \frac{1 - \bar{D}_k}{1 - \bar{D}_\ell}$$

t_k^* t_ℓ^*

D. Late Group vs. Early Group, after t_k^*

Size:

$$((n_k + n_\ell) \bar{D}_k)^2$$

×

Variance:

$$n_{k\ell} (1 - n_{k\ell}) \frac{\bar{D}_\ell}{\bar{D}_k} \frac{\bar{D}_k - \bar{D}_\ell}{\bar{D}_k}$$

t_k^* t_ℓ^*

Difference-in-Differences Decomposition Theorem (3 Group Case)

$$y_{it} = \alpha_i + \alpha_t + \hat{\beta}^{DD} D_{it} + u_{it}$$

For three groups:

$$\hat{\beta}^{DD} = s_{kU} \hat{\beta}_{kU}^{2 \times 2} + s_{\ell U} \hat{\beta}_{\ell U}^{2 \times 2} + s_{k\ell}^k \hat{\beta}_{k\ell}^{2 \times 2, k} + s_{k\ell}^\ell \hat{\beta}_{k\ell}^{2 \times 2, \ell}$$

A. Early Group vs. Untreated Group

Size:

$$(n_k + n_U)^2$$

×

Variance:

$$n_{kU} (1 - n_{kU}) \bar{D}_k (1 - \bar{D}_k)$$

B. Late Group vs. Untreated Group

Size:

$$(n_\ell + n_U)^2$$

×

Variance:

$$n_{\ell U} (1 - n_{\ell U}) \bar{D}_\ell (1 - \bar{D}_\ell)$$

C. Early Group vs. Late Group, before t_k^*

Size:

$$((n_k + n_\ell)(1 - \bar{D}_\ell))^2$$

×

Variance:

$$n_{k\ell} (1 - n_{k\ell}) \frac{\bar{D}_k - \bar{D}_\ell}{1 - \bar{D}_\ell} \frac{1 - \bar{D}_k}{1 - \bar{D}_\ell}$$

D. Late Group vs. Early Group, after t_k^*

Size:

$$((n_k + n_\ell) \bar{D}_k)^2$$

×

Variance:

$$n_{k\ell} (1 - n_{k\ell}) \frac{\bar{D}_\ell}{\bar{D}_k} \frac{\bar{D}_k - \bar{D}_\ell}{\bar{D}_k}$$

Weights: $\frac{(\text{subsample share})^2 (\text{subsample variance of FE-adjusted } D)}{\text{total variance of FE-adjusted } D}$

What do we learn from the weights?

1. Relative importance of each kind of comparison.

“switchers vs untreated”?

“early vs late”?

“late vs early” (this is less obvious)?

More important if big group (bigger sample size) or treated closer to middle of the panel (bigger variance).

“How much” comes from timing vs comparisons to untreated.

2. Importance of specific 2x2 DDs.

Sometimes a few terms dominate.

3. Clarifies theory

The estimand and ID assumption are “variance weighted”; can compare estimand to “parameters of interest” and conduct a proper balance test

What will the command do?

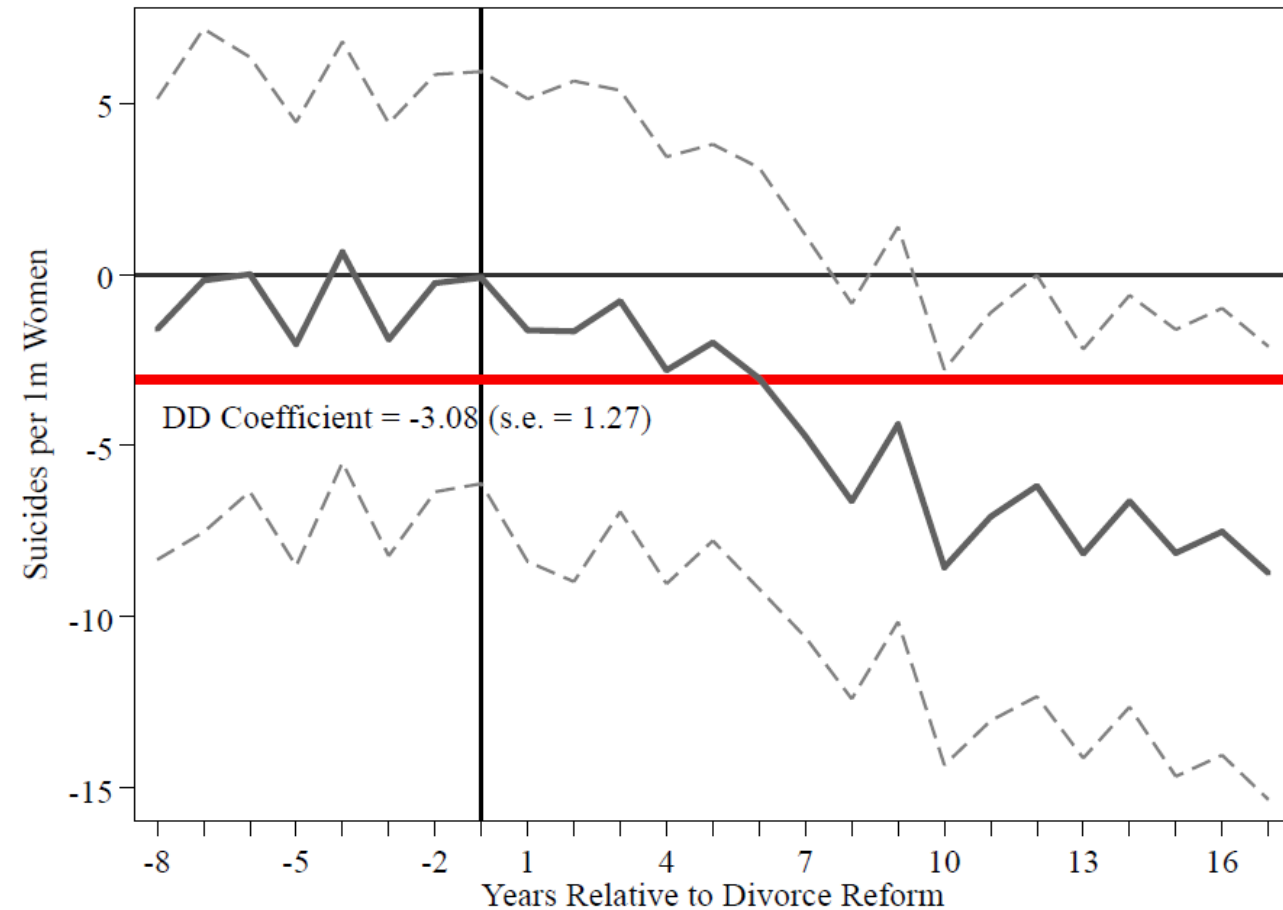
- Describe where the DD “comes from”
 - Which 2x2s matter most? (sources of variation)
 - How different are the 2x2 DDs? (heterogeneity)

Replication: The Effect of Unilateral Divorce on Suicide (Stevenson and Wolfers 2006)

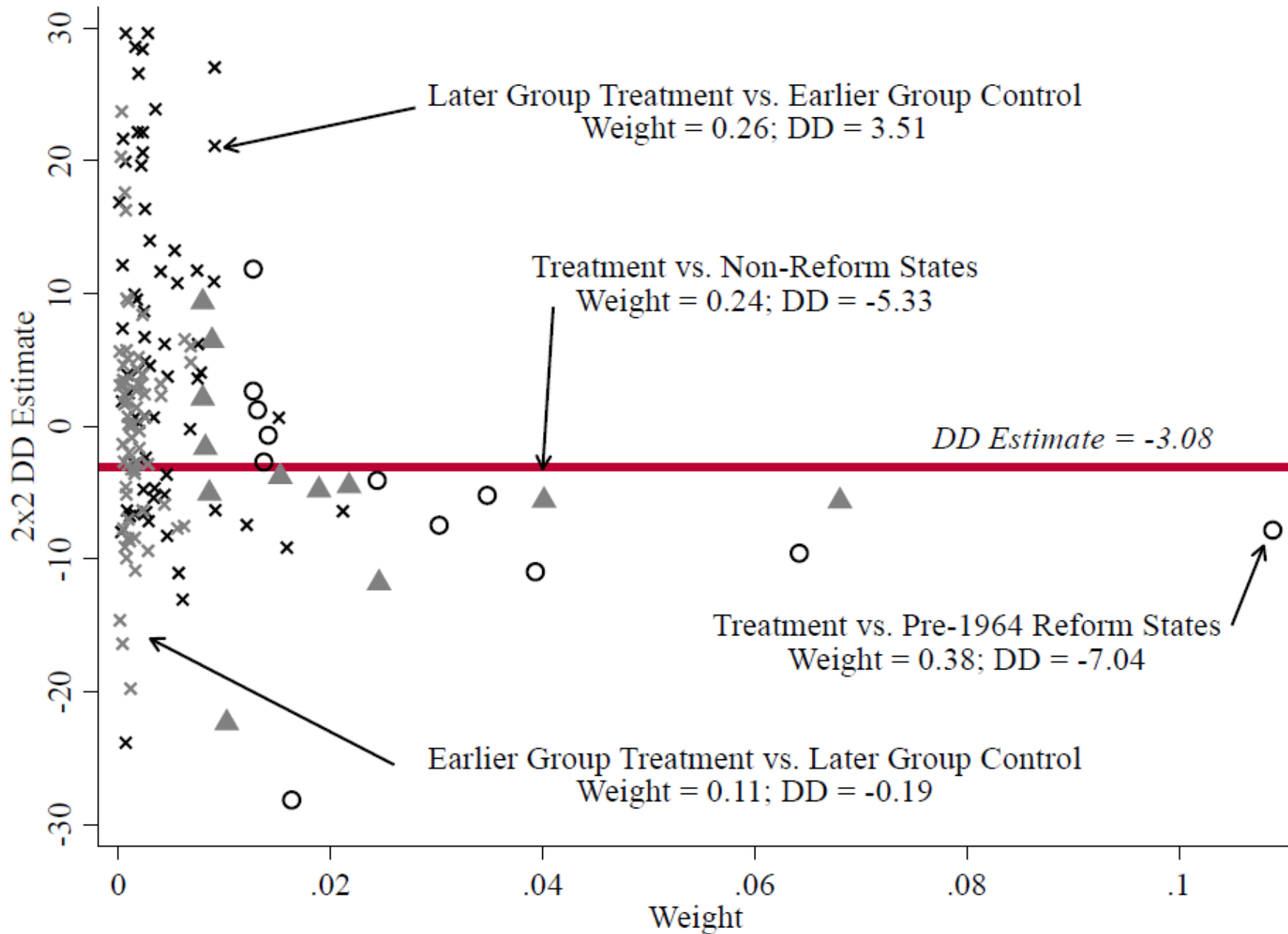
State-year panel of female suicide rates 1964-1996

12 timing groups vs Non-reform: 12 2x2 DDs
 12 timing groups vs Pre-64 reform: 12 2x2 DDs
 12 timing groups vs 12 timing groups: 12x11 = 132 2x2 DDs

| No-Fault Divorce Year (t_k^*) | Number of States | Share of States (n_k) | Treatment Share (\bar{D}_k) |
|-----------------------------------|------------------|---------------------------|---------------------------------|
| Non-Reform States | 5 | 0.10 | . |
| Pre-64 Reform States | 8 | 0.16 | . |
| 1969 | 2 | 0.04 | 0.85 |
| 1970 | 2 | 0.04 | 0.82 |
| 1971 | 7 | 0.14 | 0.79 |
| 1972 | 3 | 0.06 | 0.76 |
| 1973 | 10 | 0.20 | 0.73 |
| 1974 | 3 | 0.06 | 0.70 |
| 1975 | 2 | 0.04 | 0.67 |
| 1976 | 1 | 0.02 | 0.64 |
| 1977 | 3 | 0.06 | 0.61 |
| 1980 | 1 | 0.02 | 0.52 |
| 1984 | 1 | 0.02 | 0.39 |
| 1985 | 1 | 0.02 | 0.36 |



Graphing the Decomposition: Divorce Example



What will the command do?

- Describe where the DD “comes from”
 - Which 2x2s matter most? (sources of variation)
 - How different are the 2x2 DDs? (heterogeneity)
- Calculate why estimates differ across specifications:
 - Is it the weights, the 2x2 DDs, or both?

Comparing two weighted averages

$$\hat{\beta}^{DD} = s' \hat{\beta}^{2 \times 2}$$

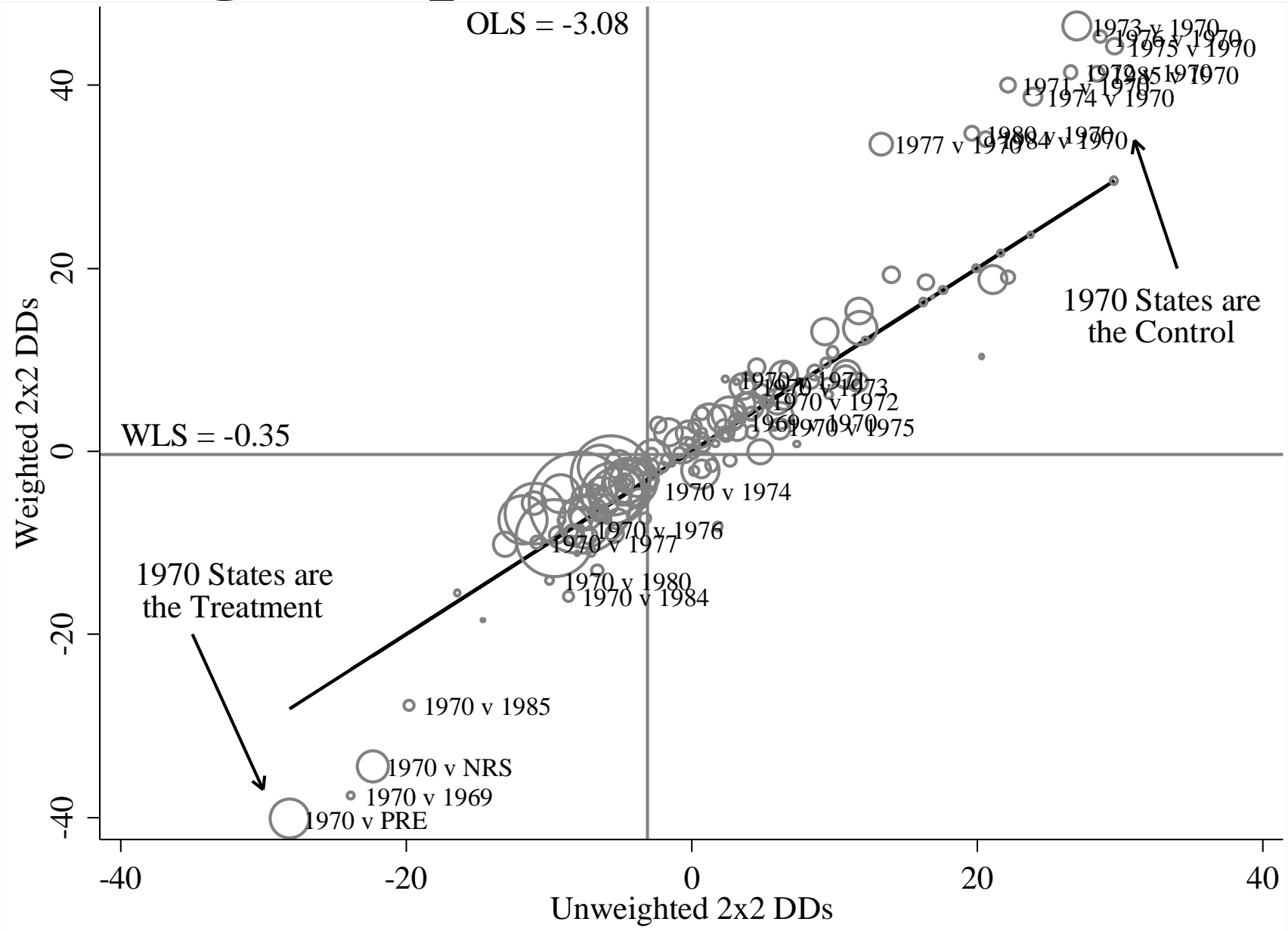
Now imagine an alternative specification that also has this form:

$$\hat{\beta}_{alt}^{DD} = s'_{alt} \hat{\beta}_{alt}^{2 \times 2}$$

If $\hat{\beta}_{alt}^{DD} \neq \hat{\beta}^{DD}$, why? (Oaxaca/Blinder/Kitagawa decomposition)

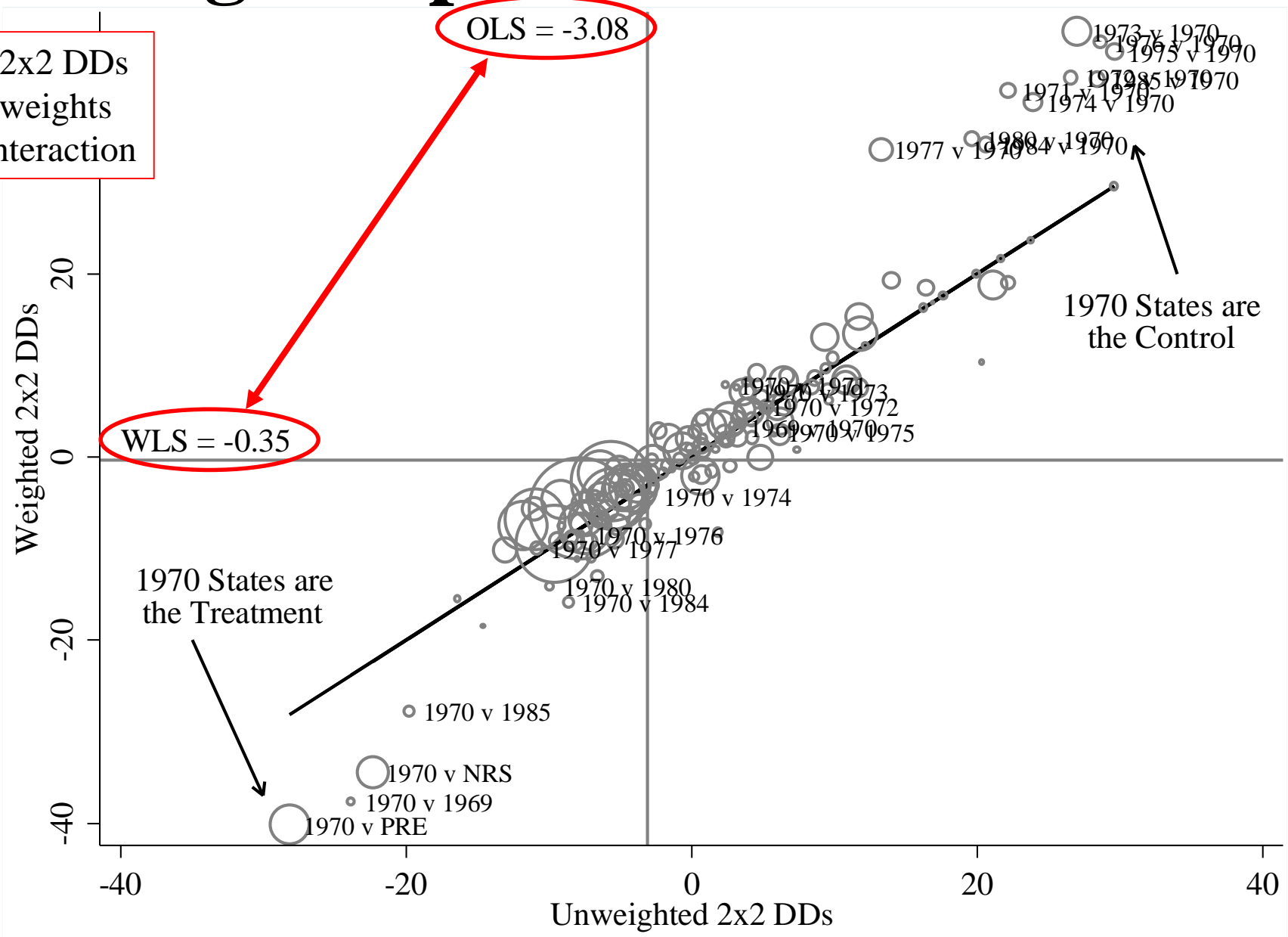
$$\underbrace{s' (\hat{\beta}_{alt}^{2 \times 2} - \hat{\beta}^{2 \times 2})}_{\text{Due to 2x2 DDs}} + \underbrace{(s'_{alt} - s') \hat{\beta}^{2 \times 2}}_{\text{Due to weights}} + \underbrace{(s'_{alt} - s') (\hat{\beta}_{alt}^{2 \times 2} - \hat{\beta}^{2 \times 2})}_{\text{Due to interaction}}$$

Plotting components: WLS vs. OLS



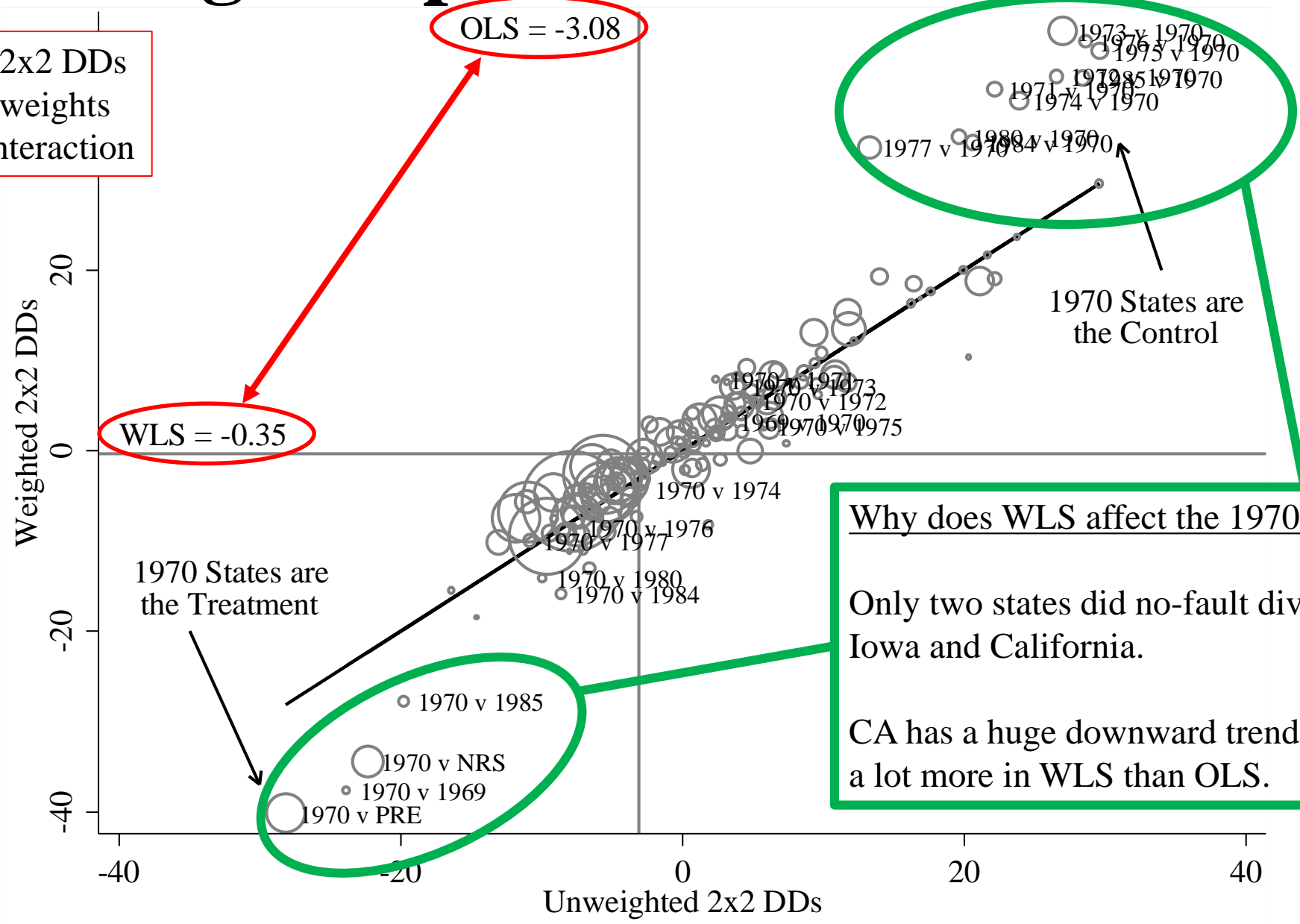
Plotting components: WLS vs. OLS

53% from 2x2 DDs
38% from weights
9% from interaction



Plotting components: WLS vs. OLS

53% from 2x2 DDs
38% from weights
9% from interaction



Why does WLS affect the 1970 states so much?

Only two states did no-fault divorce in 1970: Iowa and California.

CA has a huge downward trend and it matters a lot more in WLS than OLS.

Conclusion

- When treatment timing varies, the (two-way fixed effects) regression DD coefficient is a weighted average of simple 2x2 DDs (Goodman-Bacon 2018)
- This command will plot the 2x2 DDs against their weight to highlight where identification comes from and how heterogeneous are the 2x2 DDs.
 - “How much” variation comes from timing?
 - What is “the” control group?
 - Weights do NOT rely on outcome data (can apply to it to samples you don’t yet have)
- This command allows users to analyze why estimates change under different specifications (e.g. weights, controls, triple-diff)
- Future: test covariate balance (accounting for timing), compare estimand to other parameters of interest, adjust for bias from time-varying effects.

References

Goodman-Bacon, Andrew. 2018. "Difference-in-Differences with Variation in Treatment Timing." *National Bureau of Economic Research Working Paper Series* No. 25018. doi: 10.3386/w25018.

Stevenson, Betsey, and Justin Wolfers. 2006. "Bargaining in the Shadow of the Law: Divorce Laws and Family Distress." *The Quarterly Journal of Economics* 121 (1):267-288.

Textbooks and Survey Articles that describe 2x2 DD:

Angrist, Joshua D., and Alan B. Krueger. 1999. "Chapter 23 - Empirical Strategies in Labor Economics." In *Handbook of Labor Economics*, edited by Orley C. Ashenfelter and David Card, 1277-1366. Elsevier.

Angrist, Joshua David, and Jörn-Steffen Pischke. 2009. *Mostly harmless econometrics : an empiricist's companion*. Princeton: Princeton University Press.

Angrist, Joshua David, and Jörn-Steffen Pischke. 2015. *Mastering 'metrics : the path from cause to effect*. Princeton ; Oxford: Princeton University Press.

Cameron, Adrian Colin, and P. K. Trivedi. 2005. *Microeconometrics : methods and applications*. Cambridge ; New York: Cambridge University Press.

Heckman, James J., Robert J. Lalonde, and Jeffrey A. Smith. 1999. "Chapter 31 - The Economics and Econometrics of Active Labor Market Programs." In *Handbook of Labor Economics*, edited by Orley C. Ashenfelter and David Card, 1865-2097. Elsevier.

Meyer, Bruce D. 1995. "Natural and Quasi-Experiments in Economics." *Journal of Business & Economic Statistics* 13 (2):151-161. doi: 10.2307/1392369.

Wooldridge, Jeffrey M. 2010. *Econometric analysis of cross section and panel data*. 2nd ed. Cambridge, Mass.: MIT Press.

Recent Research on DD with Timing

- Abraham, Sarah, and Liyang Sun. 2018. "Estimating Dynamic Treatment Effects in Event Studies with Heterogeneous Treatment Effects." Working Paper.
- Athey, Susan, and Guido W. Imbens. 2018. "Design-based Analysis in Difference-in-Differences Settings with Staggered Adoption." Working Paper.
- Bitler, Marianne P., Jonah B. Gelbach, and Hilary W. Hoynes. 2003. "Some Evidence on Race, Welfare Reform, and Household Income." *The American Economic Review* 93 (2):293-298. doi: 10.2307/3132242.
- Borusyak, Kirill, and Xavier Jaravel. 2017. "Revisiting Event Study Designs." Harvard University Working Paper.
- Callaway, Brantly, and Pedro Sant'Anna. 2018. "Difference-in-Differences With Multiple Time Periods and an Application on the Minimum Wage and Employment." Working Paper.
- Chernozhukov, Victor, Iván Fernández-Val, Jinyong Hahn, and Whitney Newey. 2013. "Average and Quantile Effects in Nonseparable Panel Models." *Econometrica* 81 (2):535-580. doi: 10.3982/ECTA8405.
- de Chaisemartin, C., and X. D'Haultfœuille. 2018a. "Fuzzy Differences-in-Differences." *The Review of Economic Studies* 85 (2):999-1028. doi: 10.1093/restud/rdx049.
- de Chaisemartin, C., and X. D'Haultfœuille. 2018b. "Two-way fixed effects estimators with heterogeneous treatment effects." Working Paper.
- Freyaldenhoven, Simon, Christian Hansen, and Jesse M. Shapiro. 2018. "Pre-event Trends in the Panel Event-study Design." National Bureau of Economic Research Working Paper Series No. 24565. doi: 10.3386/w24565.
- Gibbons, Charles, E., Juan Carlos Suárez Serrato, and Michael Urbancic, B. 2018. Broken or Fixed Effects? In *Journal of Econometric Methods*.
- Imai, Kosuke, In Song Kim, and Erik Wang. 2018. "Matching Methods for Causal Inference with Time-Series Cross-Section Data." Working Paper.
- Krolikowski, Pawel. 2017. "Choosing a Control Group for Displaced Workers." *ILR Review*:0019793917743707. doi: 10.1177/0019793917743707.
- Sloczynski, Tymon. 2017. "A General Weighted Average Representation of the Ordinary and Two-Stage Least Squares Estimands." Working Paper.
- Strezhnev, Anton. 2018. "Semiparametric Weighting Estimators for Multi-Period Difference-in-Differences Designs." *Working Paper*.
- Wooldridge, Jeffrey M. 2005. "Fixed-Effects and Related Estimators for Correlated Random-Coefficient and Treatment-Effect Panel Data Models." *The Review of Economics and Statistics* 87 (2):385-390.