

Multilevel Mixed-Effects Generalized Linear Models in STATA

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Prof. Dr. Matheus Albergaria



SUMMARY

- **Theoretical Fundamentals of Multilevel Models.**
- **Estimation of Multilevel Mixed-Effects Generalized Linear Models in Stata.**



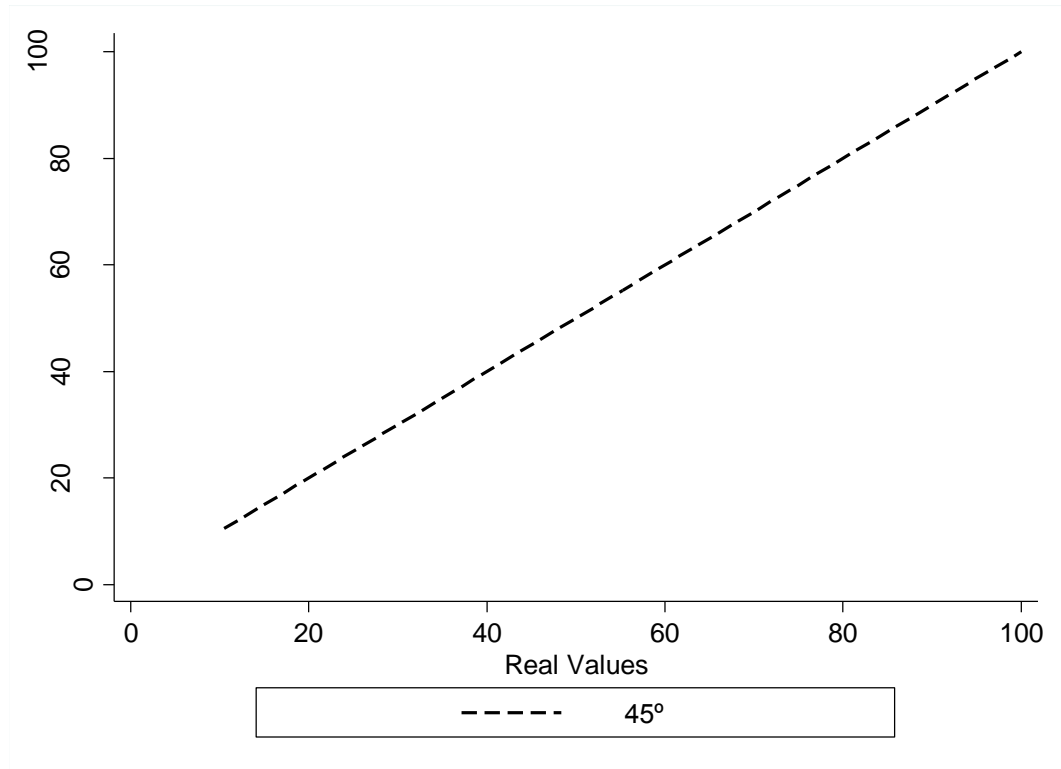
The background of the slide is a photograph of a grand, ornate spiral staircase. The staircase is viewed from above, looking down into the center. The walls of the staircase are decorated with intricate, golden-brown relief carvings of floral and geometric patterns. The metal railings are dark and feature elegant, swirling scrollwork. The overall lighting is soft, highlighting the textures and colors of the architecture. A semi-transparent purple rectangular box is centered over the image, containing the title text in white.

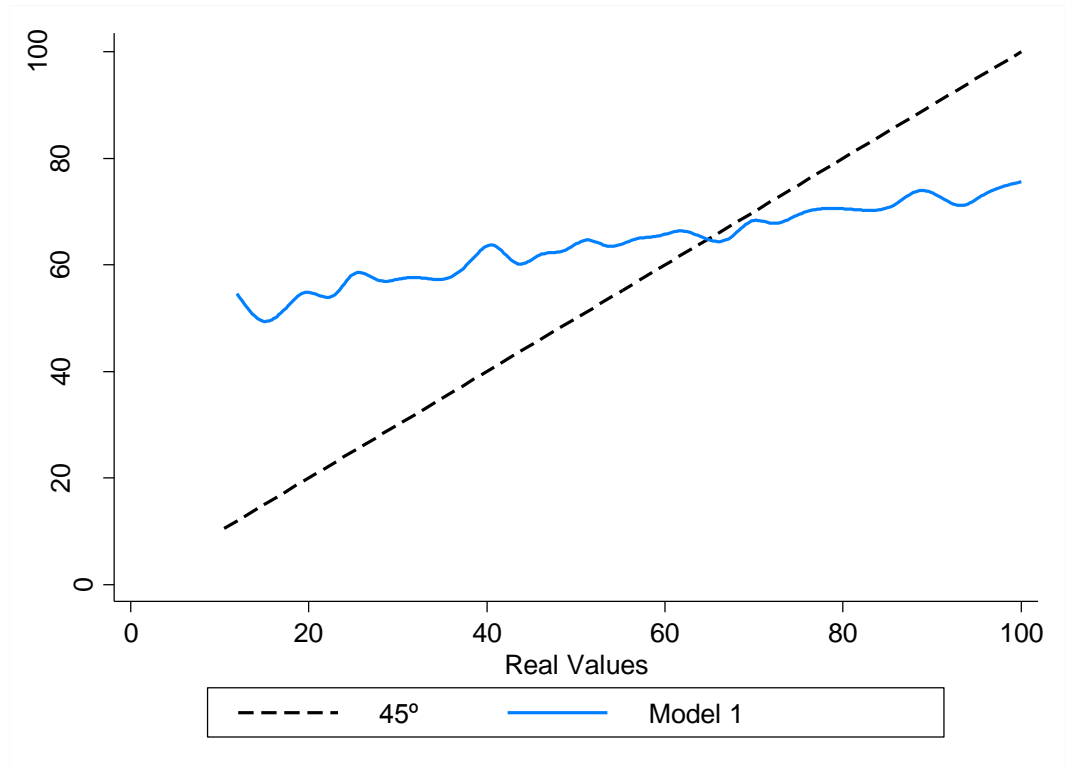
Theoretical Fundamentals of Multilevel Models

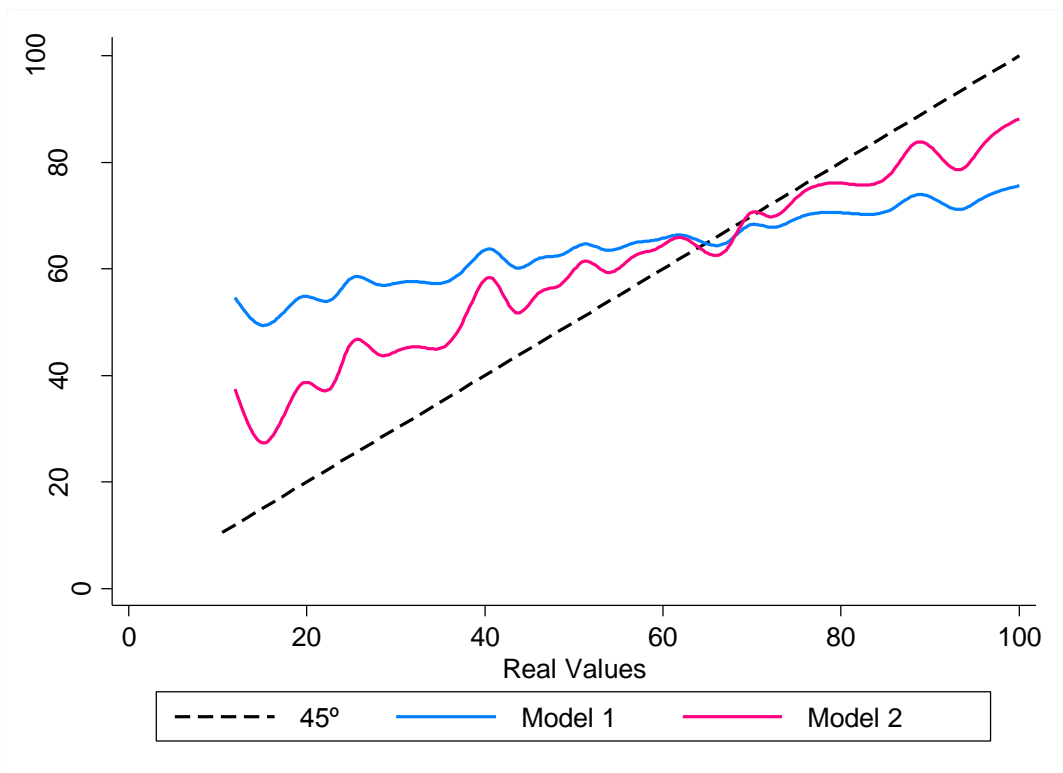


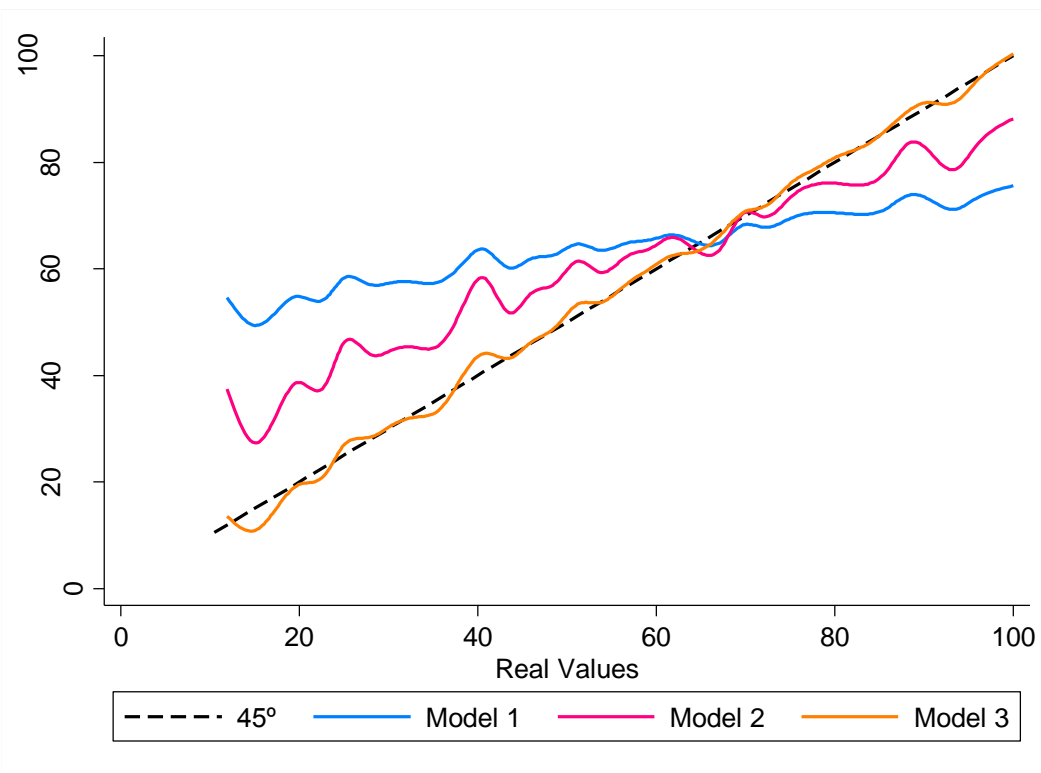
Different researchers, from the same database, can estimate different models and, consequently, obtain different predicted values of the phenomenon under study. The objective is to estimate models that, although simplifications of reality, present the best possible adherence between real and fitted values.

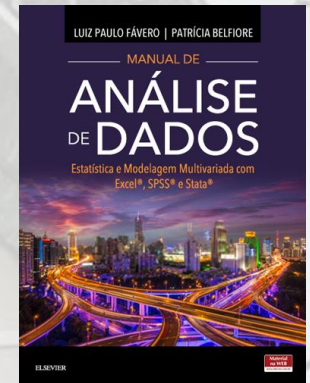
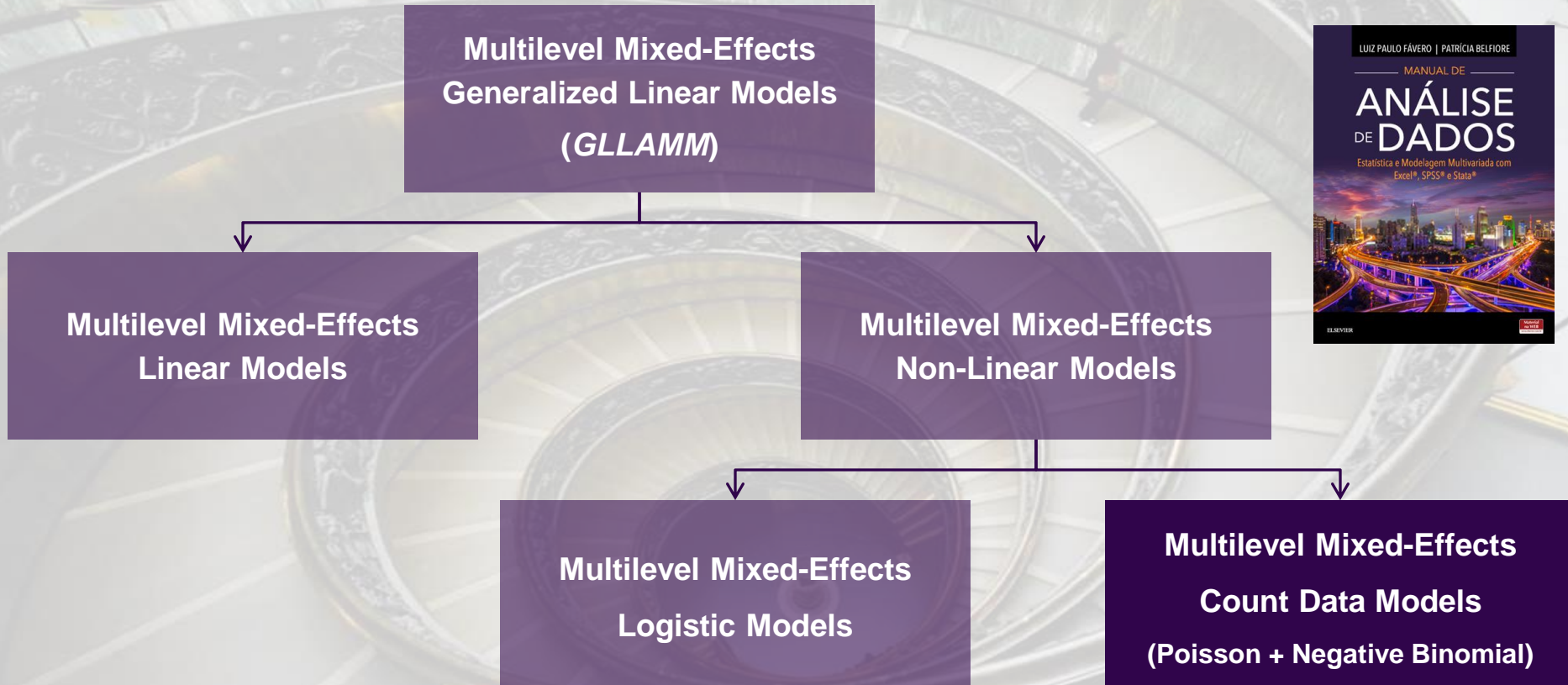
Silberzahn, R.; Uhlmann, E. L. Many hands make tight work. **Nature**, v. 526, p. 189-191, Out 2015.



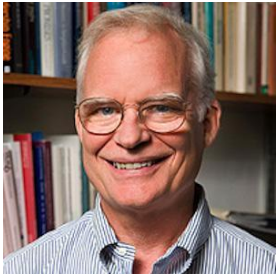






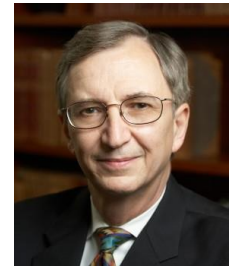


Multilevel models are models that recognize nested structure in the data.

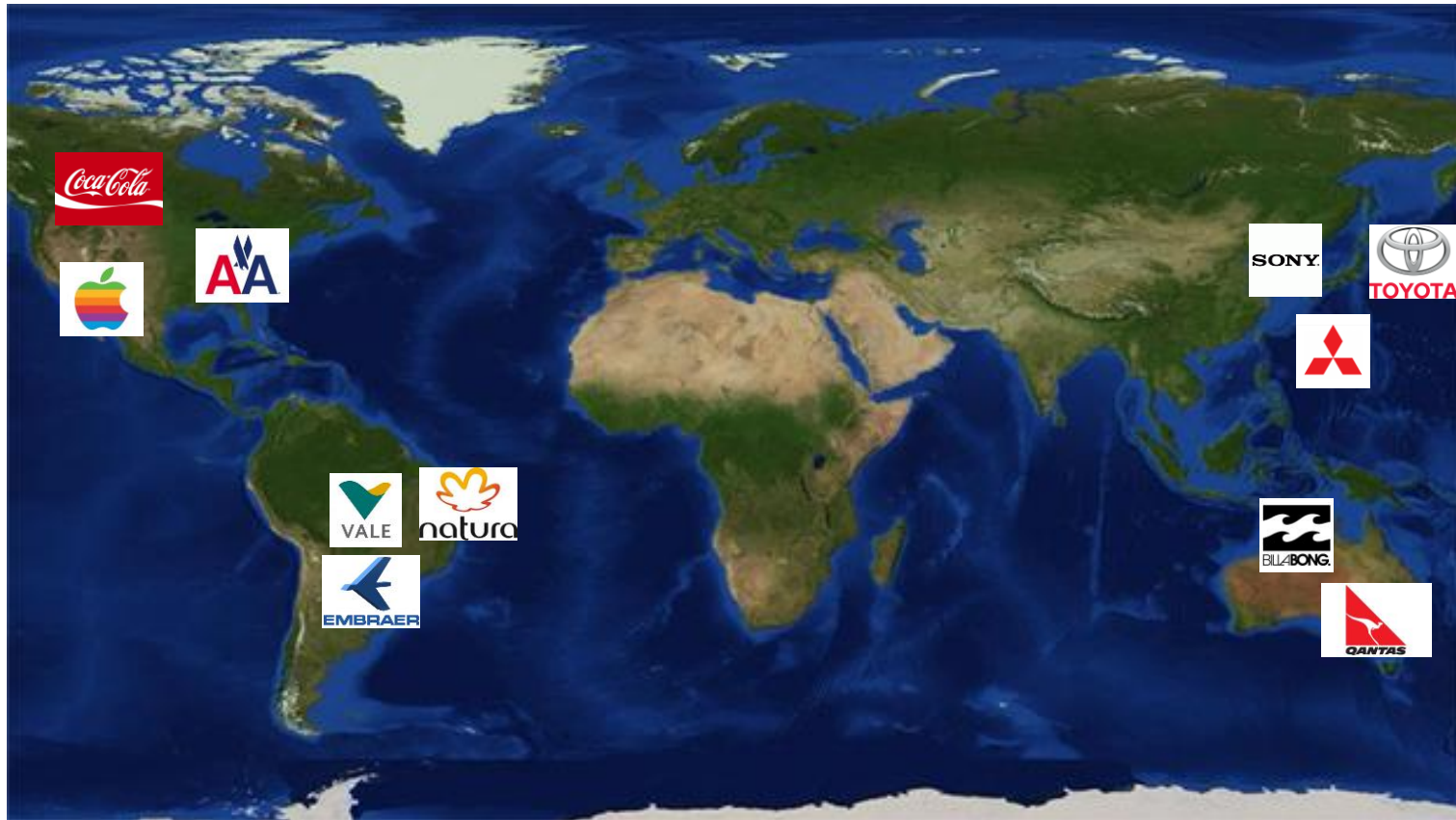


Stephen W. Raudenbush
University of Chicago

Hierarchical linear models: applications and data analysis methods. 2. ed. Thousand Oaks: Sage Publications, 2002.



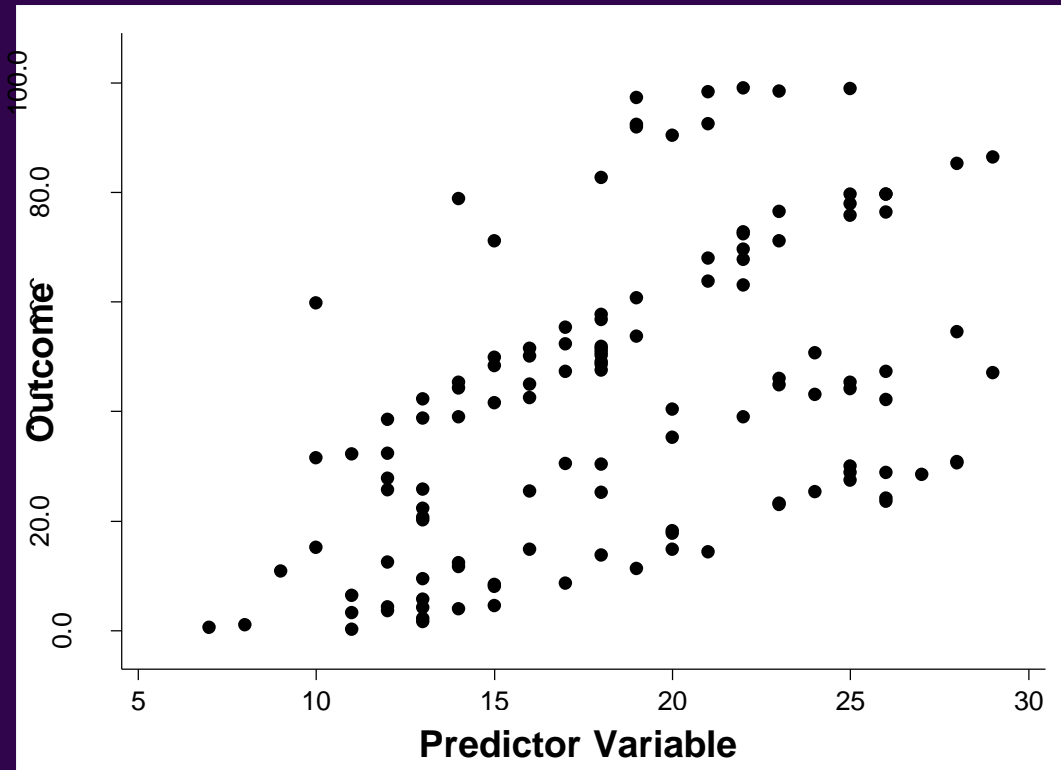
Anthony S. Bryk
Stanford University

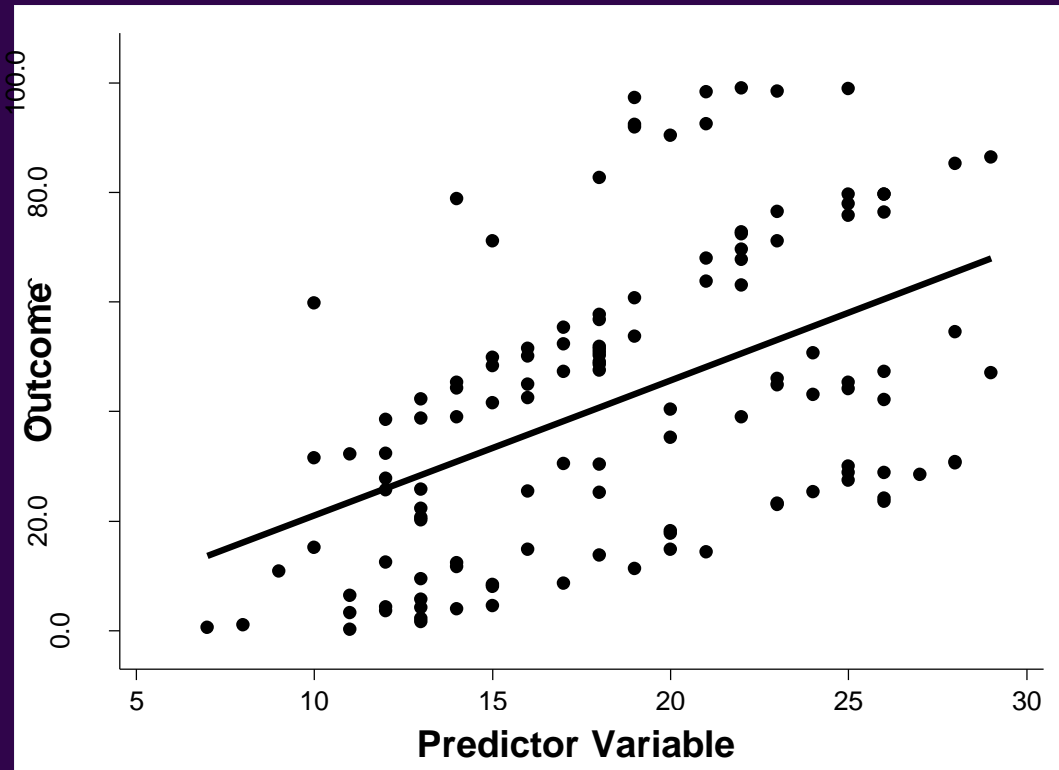


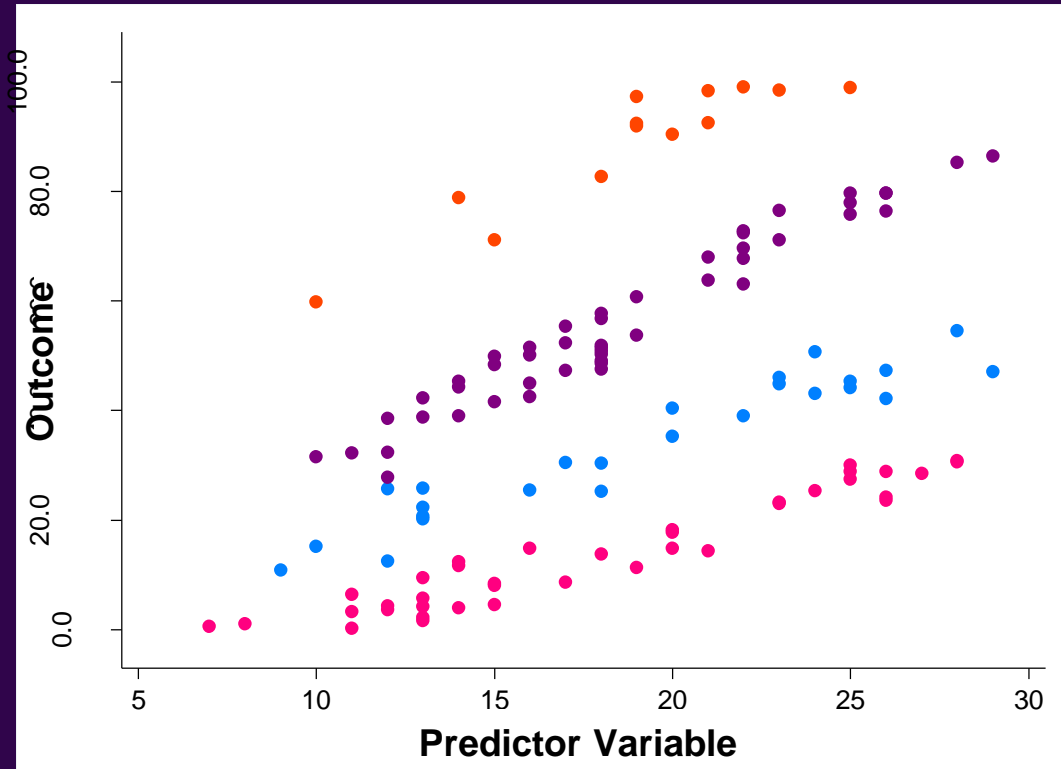
Level 1
Firm

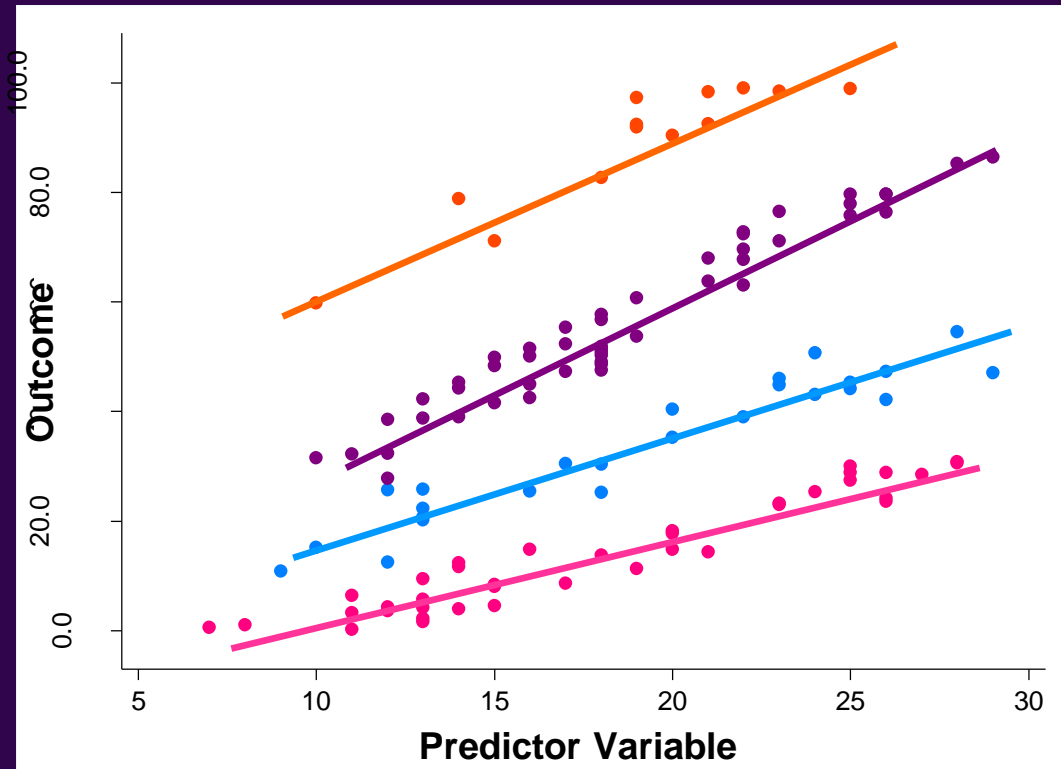


Level 2
Country









Context 1:

$$\underline{Y_{i1} = \beta_{01} + \beta_{11} \cdot X_{i1} + r_{i1}}$$

Context 2:

$$\underline{Y_{i2} = \beta_{02} + \beta_{12} \cdot X_{i2} + r_{i2}}$$

Context 3:

$$\underline{Y_{i3} = \beta_{03} + \beta_{13} \cdot X_{i3} + r_{i3}}$$

Context 4:

$$\underline{Y_{i4} = \beta_{04} + \beta_{14} \cdot X_{i4} + r_{i4}}$$

Level 1

$$\eta_{ij} = \beta_{0j} + \beta_{1j} \cdot X_{ij} + r_{ij}$$

Level 2

$$\beta_{0j} = \gamma_{00} + \gamma_{01} \cdot W_j + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11} \cdot W_j + u_{1j}$$

$$\eta_{ij} = \underbrace{(\gamma_{00} + \gamma_{01} \cdot W_j + u_{0j})}_{\text{intercept with random effects}}$$

intercept with
random effects

$$+ \underbrace{(\gamma_{10} + \gamma_{11} \cdot W_j + u_{1j})}_{\text{slope with random effects}} \cdot X_{ij} + r_{ij}$$

slope with
random effects

$$\eta_{ij} = \underbrace{\gamma_{00} + \gamma_{10} \cdot X_{ij} + \gamma_{01} \cdot W_j + \gamma_{11} \cdot W_j \cdot X_{ij}}_{\text{Fixed Effects}} + \underbrace{u_{0j} + u_{1j} \cdot X_{ij}}_{\text{Random Effects}} + r_{ij}$$

- **Traditional GLM models ignore interactions** between variables in the fixed effects component and between error terms and variables in the random effects component.

Multilevel statistical models. 4. ed. Chichester: John Wiley & Sons, 2011.

Harvey Goldstein
Centre for Multilevel Modelling
University of Bristol

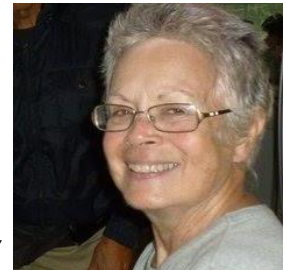


- If variances of error terms u_{0j} e u_{1j} are statistically different from zero, traditional GLM estimations will not be adequate.

Using multivariate statistics. 6. ed. Boston:
Pearson, 2013.



Barbara G. Tabachnick
California State University



Linda S. Fidell
California State University

- The inclusion of ***dummies*** representing groups do not capture **contextual effects**, because this procedure do not allow the split between observable and unobservable effects over the outcome variable.



Sophia Rabe-Hesketh
U. C. Berkeley

Multilevel and longitudinal modeling using Stata. 3.
ed. College Station: Stata Press, 2012.

Anders Skrondal
Norwegian Institute of Public Health
University of Oslo
U. C. Berkeley



Multilevel models allow the development of new and more complex research constructs.

Within a model structure with a single equation, there seems to be no connection between individuals and the society in which they live. In this sense, the use of level equations allows the researcher to 'jump' from one science to another: students and schools, families and neighborhoods, firms and countries. Ignoring this relationship means elaborate incorrect analyzes about the behavior of the individuals and, equally, about the behavior of the groups. Only the recognition of these reciprocal influences allows the correct analysis of the phenomena.

Methodology and epistemology of multilevel analysis.

London: Kluwer Academic Publishers, 2003.

Daniel Courgeau
*Institut National D'Études
Démographiques*





**Estimation of Multilevel
Mixed-Effects Generalized
Linear Models in STATA[®]**

Level 1

$$\ln(\lambda_{ijk}) = \pi_{0jk} + \pi_{1jk} \cdot Z_{1jk} + \pi_{2jk} \cdot Z_{2jk} + \dots + \pi_{Pjk} \cdot Z_{Pjk}$$

Level 2

$$\pi_{pjk} = b_{p0k} + \sum_{q=1}^{Q_p} b_{pqk} \cdot X_{qjk} + r_{pjk}$$

Level 3

$$b_{pqk} = \gamma_{pq0} + \sum_{s=1}^{S_{pq}} \gamma_{pqs} \cdot W_{sk} + u_{pqk}$$

Stata example

- Let's look at the relationship between traffic accidents and alcohol consumption (Fávero and Belfiore, 2017).
- We want to estimate the relationship between the number of traffic accidents and the consumption alcohol per person/day in the district, considering differences in cities and states.

Data description (file “TrafficAccidents.dta”)

```
. desc
```

```
obs:          1,062
vars:           5
size:         15,930
```

```
-----
      storage  display
variable name  type   format  variable label
-----
state          str2   %2s     state k (level 3)
city           int     %8.0g   city j (level 2)
district       int     %8.0g   municipal district i (level 1)
accidents      byte   %8.0g   number of traffic accidents in the district over the last year
alcohol        float  %9.2f   average consumption of alcohol per person/day in the district (in grams)
-----
```

```
Sorted by:
```

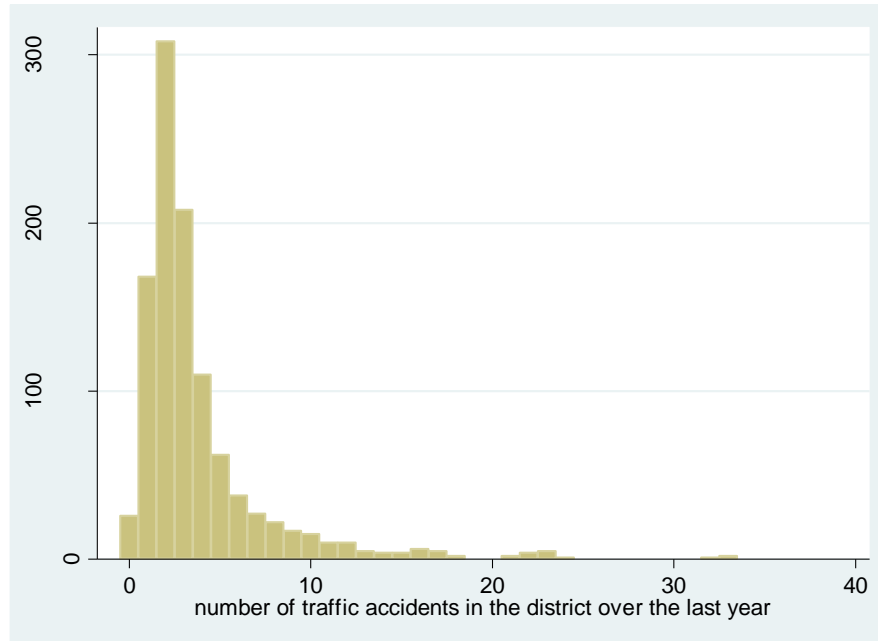

Data tabulation

```
. tab accidents
```

number of traffic accidents	Freq.	Percent	Cum.
0	26	2.45	2.45
1	168	15.82	18.27
2	308	29.00	47.27
3	208	19.59	66.85
4	110	10.36	77.21
5	62	5.84	83.05
6	38	3.58	86.63
7	27	2.54	89.17
8	22	2.07	91.24
9	17	1.60	92.84
10	15	1.41	94.26
11	10	0.94	95.20
12	10	0.94	96.14
13	5	0.47	96.61
14	4	0.38	96.99
15	4	0.38	97.36
16	6	0.56	97.93
17	5	0.47	98.40
18	2	0.19	98.59
21	2	0.19	98.78
22	4	0.38	99.15
23	5	0.47	99.62
24	1	0.09	99.72
32	1	0.09	99.81
33	2	0.19	100.00
Total	1,062	100.00	

Histogram

```
. hist accidents, discrete freq  
(start=0, width=1)
```



Mean and variance (possible existence of overdispersion)

```
. tabstat accidents, stats(mean var)

  variable |      mean  variance
-----+-----
 accidents |   3.812618  15.24007
```

Proposed Model

$$\ln(\text{accidents}_{ijk}) = \pi_{0jk} + \pi_{1jk} \cdot \text{alcohol}_{jk}$$

$$\pi_{0jk} = b_{00k} + r_{0jk}$$

$$\pi_{1jk} = b_{10k}$$

$$b_{00k} = \gamma_{000} + u_{00k}$$

$$b_{10k} = \gamma_{100}$$

$$\ln(\text{accidents}_{ijk}) = \gamma_{000} + \gamma_{100} \cdot \text{alcohol}_{jk} + u_{00k} + r_{0jk}$$

Multilevel Mixed-Effects Poisson Model

```
. meglm accidents alcohol || state: || city: , family(poisson) link(log) nolog
```

Mixed-effects GLM Number of obs = 1062
 Family: Poisson
 Link: log

Group Variable	No. of Groups	Observations per Group		
		Minimum	Average	Maximum
state	27	1	39.3	95
city	235	1	4.5	13

Integration method: mvaghermite Integration points = 7
 Wald chi2(1) = 5.60
 Log likelihood = -2295.9047 Prob > chi2 = 0.0180

accidents	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
alcohol	.0478279	.020216	2.37	0.018	.0082053	.0874506
_cons	.7293659	.2638594	2.76	0.006	.2122111	1.246521

state						
var(_cons)	.3857761	.12319			.2063103	.7213563

state>city						
var(_cons)	.0829691	.0142976			.059188	.1163053

LR test vs. Poisson regression: chi2(2) = 1279.65 Prob > chi2 = 0.0000

```
. estimates store mepoisson
```

Multilevel Mixed-Effects Negative Binomial Model

```
. meglm accidents alcohol || state: || city: , family(nbinomial) link(log) nolog
```

Mixed-effects GLM Number of obs = 1062
Family: negative binomial
Link: log
Overdispersion: mean

Group Variable	No. of Groups	Observations per Group		
		Minimum	Average	Maximum
state	27	1	39.3	95
city	235	1	4.5	13

Integration method: mvaghermite Integration points = 7
Wald chi2(1) = 4.38
Log likelihood = -2234.3721 Prob > chi2 = 0.0363

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
accidents						
alcohol	.0466768	.0222975	2.09	0.036	.0029746	.0903791
_cons	.7538477	.2843403	2.65	0.008	.196551	1.311144
/lnalpha	-2.258241	.1355339	-16.66	0.000	-2.523883	-1.9926
state var(_cons)	.3775391	.1205934			.2018698	.7060775
state>city var(_cons)	.0613878	.0138809			.0394104	.0956212

LR test vs. nbinomial regression: chi2(2) = 508.99 Prob > chi2 = 0.0000

```
. estimates store menegbin
```


Likelihood-ratio test

```
. lrtest mepoisson menegbin
```

```
Likelihood-ratio test                LR chi2(1)  =    123.07  
(Assumption: mepoisson nested in menegbin)  Prob > chi2 =    0.0000
```

```

predict lambda

predict u00 r0, remeans

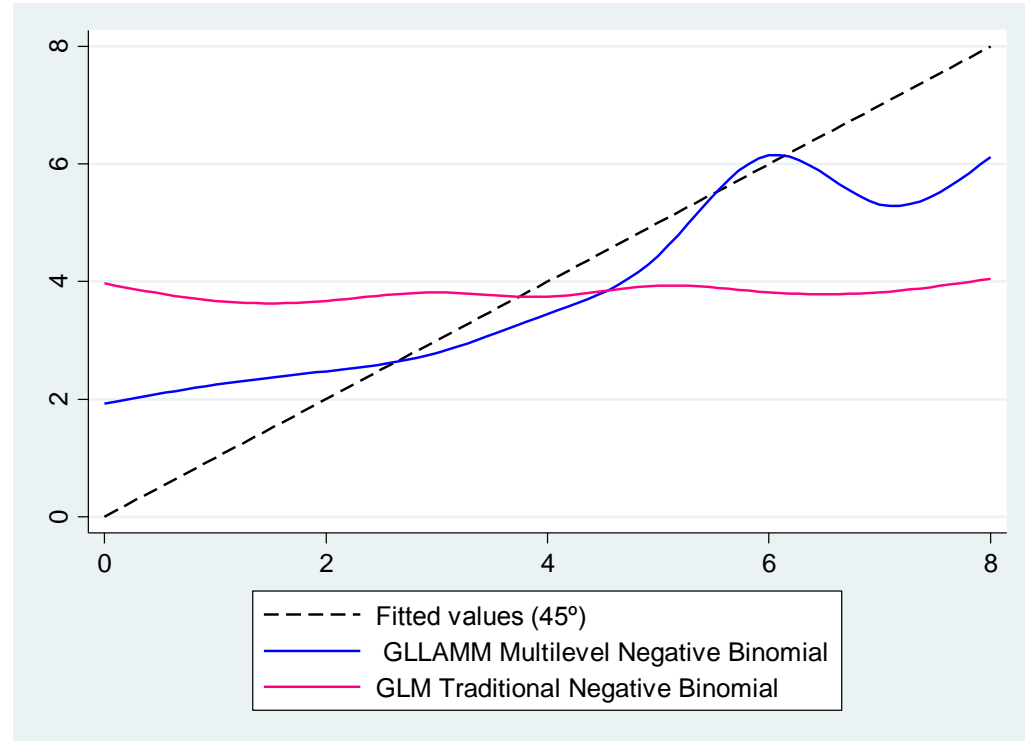
list state city accidents lambda u00 r0 if
state=="MT", sepby(city)

```

	state	city	accidents	lambda	u00	r0
669.	MT	148	2	1.600369	-.815816	-.0064477
670.	MT	148	2	1.63053	-.815816	-.0064477
671.	MT	148	1	1.63053	-.815816	-.0064477
672.	MT	148	1	1.585499	-.815816	-.0064477
673.	MT	148	2	1.499133	-.815816	-.0064477
674.	MT	149	0	1.415119	-.815816	-.1107979
675.	MT	149	3	1.441788	-.815816	-.1107979
676.	MT	149	1	1.428391	-.815816	-.1107979
677.	MT	149	1	1.441788	-.815816	-.1107979
678.	MT	149	1	1.338034	-.815816	-.1107979
679.	MT	149	1	1.388943	-.815816	-.1107979
680.	MT	149	2	1.415119	-.815816	-.1107979
681.	MT	149	1	1.350584	-.815816	-.1107979
682.	MT	149	1	1.350584	-.815816	-.1107979
683.	MT	149	2	1.40197	-.815816	-.1107979
684.	MT	149	1	1.376037	-.815816	-.1107979
685.	MT	149	1	1.441788	-.815816	-.1107979
686.	MT	150	2	1.667662	-.815816	.01607
687.	MT	150	2	1.576821	-.815816	.01607
688.	MT	150	1	1.621606	-.815816	.01607
689.	MT	150	2	1.547654	-.815816	.01607
690.	MT	150	1	1.547654	-.815816	.01607
691.	MT	150	2	1.533273	-.815816	.01607
692.	MT	151	1	1.462078	-.815816	-.031476
693.	MT	151	2	1.489632	-.815816	-.031476
694.	MT	151	1	1.517706	-.815816	-.031476

```
quietly nbreg accidents alcohol
predict lambda trad

graph twoway lfit accidents accidents ||
mspline lambda accidents || mspline
lambda trad accidents ||, legend(label(2 "
GLLAMM Multilevel Negative Binomial")
label(3 "GLM Traditional Negative Binomial
"))
```





Deep interactions

Methods of parameter estimations

Sample clustering

Estimation of models with better adjustment between real and fitted values!



Andrew Gelman

Multilevel Conference, 31 Oct 2015, Columbia University, NYC.

- Multilevel Mixed-Effects Generalized Linear Models: still employed with parsimony today.
- Stata 15 has a full command suite for the estimation of these models.
- Several research opportunities, both in theoretical and applied terms, in areas such as microeconomics, finance, transportation, real estate, leisure, ecology, education, and health.

COURGEAU, D. **Methodology and epistemology of multilevel analysis**. London: Kluwer Academic Publishers, 2003.

FÁVERO, L. P.; BELFIORE, P. **Manual de análise de dados: estatística e modelagem multivariada com Excel[®], SPSS[®] e Stata[®]**. Rio de Janeiro: Elsevier, 2017.

RABE-HESKETH, S.; SKRONDAL, A. **Multilevel and longitudinal modeling using Stata: continuous responses (Vol. I)**. 3. ed. College Station: Stata Press, 2012.





We must widen the circle of our love till it embraces the whole village; the village in its turn must take into its fold the district; the district the province; and so on, until the scope of our love becomes co-terminous with the world.



The Stata logo is displayed in a white, bold, sans-serif font against a dark purple rectangular background.

Thank you!

Multilevel Mixed-Effects Generalized Linear Models in Stata

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Prof. Dr. Matheus Albergaria - matheus.albergaria@usp.br




```
*****
*Stata do-file for the Presentation "Multilevel Mixed-Effects Generalized
*Linear Models in Stata", by Luiz Paulo Fávero and Matheus Albergaria
*2017 Brazilian Stata Users Group Meeting
*Universidade de São Paulo (USP), São Paulo, Brazil
*December 8th, 2017

*This do-file was written by Luiz Paulo Fávero and Matheus Albergaria
*The data file is "TrafficAccidents.dta". For more details, see:
*Fávero, L.P.; Belfiore, P. (2017) "Manual de Análise de Dados:
*estatística e modelagem multivariada com Excel, SPSS e Stata (Chapter 16)
*****

*Open Dataset
use C:\TrafficAccidents.dta

*Data Description
desc

*Data Tabulation
tab accidents
hist accidents, discrete freq
```

```
*Descriptive Statistics
tabstat accidents, stats(mean var)

*Multilevel Mixed-Effects Count Models Estimation
meglm accidents alcohol || state: || city: , family(poisson) link(log) nolog
estimates store mepoisson
meglm accidents alcohol || state: || city: , family(nbinomial) link(log) nolog
estimates store menegbin

*Fitting Distinct Models
lrtest mepoisson menegbin

predict lambda

predict u00 r0, remeans
list state city accidents lambda u00 r0 if state == "MT", seby(city)
quietly nbreg accidents alcohol
predict lambdatrad
graph twoway lfit accidents accidents || mspline lambda accidents || ///
mspline lambdatrad accidents ||, ///
legend(label(2 " GLLMM Multilevel Negative Binomial") ///
label(3 "GLM Traditional Negative Binomial "))
```